CPSC 540: Machine Learning Kernel Density Estimation

Mark Schmidt

University of British Columbia

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Last Time: Expectation Maximization

- EM considers learning with observed data O and hidden data H.
- In this case the "observed" log-likelihood has a nasty form,

$$\log p(O \mid \Theta) = \log \left(\sum_{H} p(O, H \mid \Theta) \right).$$

- EM applies when "complete" likelihood, $p(O, H \mid \Theta)$, has a nice form.
- EM iterations take the form of a weighted "complete" NLL,

$$\Theta^{t+1} = \operatorname*{argmax}_{\Theta} \left\{ \sum_{H} \alpha_{H} \log p(O, H \mid \Theta) \right\},$$

where $\alpha_H = p(H \mid O, \Theta^t)$.

• For mixture models, has a closed-form solution for common distributions.

Monotonicity of EM

• Classic result is that EM iterations are monotonic:

$$\log p(O \mid \Theta^{t+1}) \ge \log p(O \mid \Theta^t),$$

- We don't need a step-size and this is useful for debugging.
- We can show this by proving that the below picture is "correct":



- The Q function leads to a global bound on the original function.
- At Θ^t the bound matches original function.
 - \bullet So if you improve on the Q function, you improve on the original function.

Monotonicity of EM

• Let's show that the Q function gives a global upper bound on NLL:

$$\begin{split} -\log p(O\mid\Theta) &= -\log\left(\sum_{H}p(O,H\mid\Theta)\right) & \text{(marginalization rule)} \\ &= -\log\left(\sum_{H}\alpha_{H}\frac{p(O,H\mid\Theta)}{\alpha_{H}}\right) & \text{(for }\alpha_{H}\neq0) \\ &\leq -\sum_{H}\alpha_{H}\log\left(\frac{p(O,H\mid\Theta)}{\alpha_{H}}\right), \end{split}$$

because $-\log(z)$ is convex and the α_H are a convex combination.

Monotonicity of EM

• Using that log turns multiplication into addition we get

$$\begin{split} -\log p(O\mid\Theta) &\leq -\sum_{H} \alpha_{H} \log \left(\frac{p(O,H\mid\Theta)}{\alpha_{H}}\right) \\ &= -\sum_{H} \alpha_{H} \log p(O,H\mid\Theta) + \sum_{H} \alpha_{H} \log \alpha_{H} \\ &\underbrace{Q(\Theta\mid\Theta^{t})}_{\text{negative entropy}} \\ &= -Q(\Theta\mid\Theta^{t}) - \text{entropy}(\alpha), \end{split}$$

so we have the first part of the picture, $-\log p(O \mid \Theta^{t+1}) \leq -Q(\Theta \mid \Theta^t) + \text{const.}$

- ullet Entropy is a measure of how "random" the $lpha_H$ values are.
- ullet Bound gets tighter for hidden data H that is more "predictable".
- Now we need to show that this holds with equality at Θ^t .

Bound on Progress of Expectation Maximization

• To show equality at Θ^t we use definition of conditional probability,

$$p(H \mid O, \Theta^t) = \frac{p(O, H \mid \Theta^t)}{p(O \mid \Theta^t)} \quad \text{or} \quad \log p(O \mid \Theta^t) = \log p(O, H \mid \Theta^t) - \log p(H \mid O, \Theta^t)$$

• Multiply by α_H and summing over H values,

$$\sum_{H} \alpha_{H} \log p(O \mid \Theta^{t}) = \underbrace{\sum_{H} \alpha_{H} \log p(O, H \mid \Theta^{t} - \sum_{H} \alpha_{H} \log \underbrace{p(H \mid O, \Theta^{t})}_{\alpha_{H}}}_{Q(\Theta^{t} \mid \Theta^{t})}.$$

• Which gives the result we want:

$$\log p(O \mid \Theta^t) \underbrace{\sum_{H} \alpha_H} = Q(\Theta^t \mid \Theta^t) + \mathrm{entropy}(\alpha),$$

Bound on Progress of Expectation Maximization

Thus we have the two bounds

$$\log p(O \mid \Theta) \ge Q(\Theta \mid \Theta^t) + \mathsf{entropy}(\alpha)$$
$$\log p(O \mid \Theta^t) = Q(\Theta^t \mid \Theta^t) + \mathsf{entropy}(\alpha).$$

• Subtracting these and using $\Theta = \Theta^{t+1}$ gives a stronger result,

$$\log p(O \mid \Theta^{t+1}) - \log p(O \mid \Theta^t) \ge Q(\Theta^{t+1} \mid \Theta^t) - Q(\Theta^t \mid \Theta^t),$$

that we improve objective by at least the decrease in Q.

- Inequality holds for any choice of Θ^{t+1} .
 - ullet Approximate M-steps are ok: we just need to decrease Q to improve likelihood.
- For imputation, we instead improve "complete" log-likelihood, $\log p(O, H \mid \Theta^t)$.
 - Which isn't quite what we want, treats hidden data as a "parameter".

Convergence of Expectation Maximization

We've shown that

$$\log p(O \mid \Theta^{t+1}) - \log p(O \mid \Theta^t) \ge Q(\Theta^{t+1} \mid \Theta^t) - Q(\Theta^t \mid \Theta^t),$$

that guaranteed progress is at least as large as difference in Q.

- Does this imply convergence?
 - Yes, the algorithm can't keep improving if the likelihood is bounded above.
- Does this imply convergence to a local optimum or a stationary point?
 - No, although many papers wrongly say that it does.
 - Could have maximum of 3 and objective values of $1, 1.5, 1.75, 1.875, \dots$

Convergence Rate of Expectation Maximization

- Can we say EM converges to stationary point or analyze convergence rate?
- If $\log p(O \mid \Theta)$ is differentiable, then we can show that

$$\nabla \log p(O \mid \Theta^t) = \nabla Q(\Theta^t \mid \Theta^t),$$

that gradient of bound agrees with gradient of function at Θ^t .

• If the bound Q is L-Lipschitz continuous, then we have

$$-Q(\Theta^{t+1} \mid \Theta^t) \le -Q(\Theta^t \mid \Theta^t) - \frac{1}{2L} \|\nabla Q(\Theta^t)\|^2,$$

since optimizing Q does at least as well as one iteration of gradient descent.

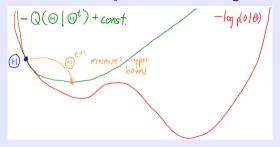
ullet Using our relationships between Q and objective f gives our usual progress bound

$$f(\Theta^{t+1}) \le f(\Theta^t) - \frac{1}{2L} \|\nabla f(\Theta^t)\|^2,$$

so EM has convergence rate at least as fast as gradient descent.

Convergence Rate of Expectation Maximization

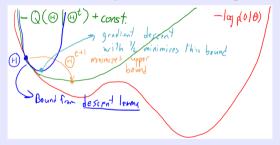
• Expectation maximization decreases f at least much as gradient descent:



ullet Slight subtle point: we measure L across Q values, rather than for f.

Convergence Rate of Expectation Maximization

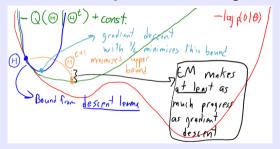
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Convergence Rate of Expectation Maximization

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Outline

Convergence of EM

2 Kernel Density Estimation

A Non-Parametric Mixture Model

• The classic parametric mixture model has the form

$$p(x^i) = \sum_{c=1}^k p(z^i = c)p(x^i \mid z^i = c).$$

• A natural way to define a non-parametric mixture model is

$$p(x^i) = \sum_{i=1}^n p(z^i = j)p(x^i \mid z^i = j),$$

where we have one mixture for every training example i.

ullet Common example: z^i is uniform and $x^i \mid z^i$ is Gaussian with mean x^j ,

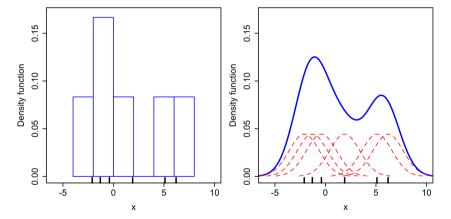
$$p(x^i) = \frac{1}{n} \sum_{i=1}^n \mathcal{N}(x^i \mid x^j, \sigma^2 I),$$

and we use a shared covariance $\sigma^2 I$ (σ can be estimated by cross-validation).

• This is a special case of kernel density estimation (or Parzen window).

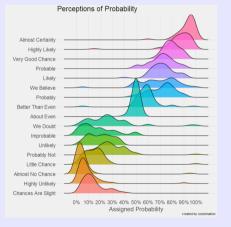
Histogram vs. Kernel Density Estimator

• Think of kernel density estimator as a generalization of histogram:



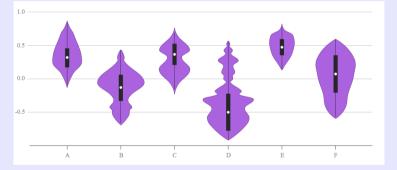
Kernel Density Estimator for Visualization

• Visualization of people's opinions about what "likely" and other words mean.



Violin Plot: Added KDE to a Boxplot

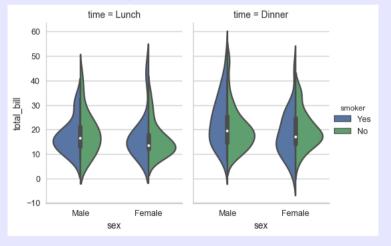
• Violin plot adds KDE to a boxplot:



https://datavizcatalogue.com/methods/violin_plot.html

Violin Plot: Added KDE to a Boxplot

• Violin plot adds KDE to a boxplot:



Kernel Density Estimation

• The 1D kernel density estimation (KDE) model uses

$$p(x^{i}) = \frac{1}{n} \sum_{j=1}^{n} k_{\sigma} \underbrace{(x^{i} - x^{j})}_{x},$$

where the PDF k is the "kernel" and the parameter σ is the "bandwidth".

• In the previous slide we used the (normalized) Gaussian kernel,

$$k_1(r) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{r^2}{2}\right), \quad k_{\sigma}(r) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{r^2}{2\sigma^2}\right).$$

• Note that we can add a "bandwith" (standard deviation) σ to any PDF k_1 , using

$$k_{\sigma}(r) = \frac{1}{\sigma} k_1 \left(\frac{r}{\sigma}\right),$$

from the change of variables formula for probabilities $(\left|\frac{d}{dr}\left[\frac{r}{\sigma}\right]\right| = \frac{1}{\sigma})$.

• Under common choices of kernels, KDEs can model any continuous density.

Efficient Kernel Density Estimation

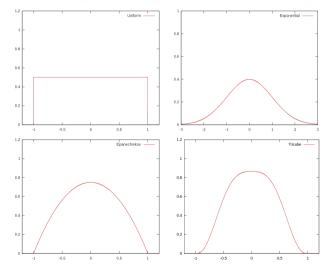
- KDE with the Gaussian kernel is slow at test time:
 - We need to compute distance of test point to every training point.
- A common alternative is the Epanechnikov kernel,

$$k_1(r) = \frac{3}{4} (1 - r^2) \mathcal{I}[|r| \le 1].$$

- This kernel has two nice properties:
 - Epanechnikov showed that it is asymptotically optimal in terms of squared error.
 - It can be much faster to use since it only depends on nearby points (use hashing).
 - You can use hashing to quickly find neighbours in training data.
- It is non-smooth at the boundaries but many smooth approximations exist.
 - Quartic, triweight, tricube, cosine, etc.

Visualization of Common Kernel Functions

Histogram vs. Gaussian vs. Epanechnikov vs. tricube:



Multivariate Kernel Density Estimation

• The multivariate kernel density estimation (KDE) model uses

$$p(x^{i}) = \frac{1}{n} \sum_{j=1}^{n} k_{A} (\underbrace{x^{i} - x^{j}}_{r}),$$

• The most common kernel is a product of independent Gaussians,

$$k_I(r) = \frac{1}{(2\pi)^{\frac{d}{2}}} \exp\left(-\frac{\|r\|^2}{2}\right).$$

 \bullet We can add a bandwith matrix A to any kernel using

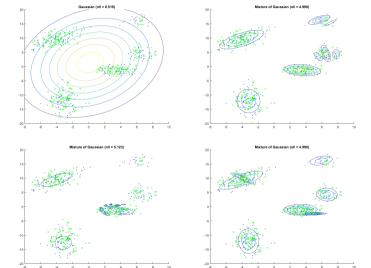
$$k_A(r) = rac{1}{|A|} k_1(A^{-1}r)$$
 (generalizes $k_\sigma(r) = rac{1}{\sigma} k_1\left(rac{r}{\sigma}
ight)$),

and in Gaussian case we get a multivariate Gaussian with $\Sigma = AA^T$.

- To reduce number of parameters, we typically:
 - Use a product of independent distributions and use $A = \sigma I$ for some σ .

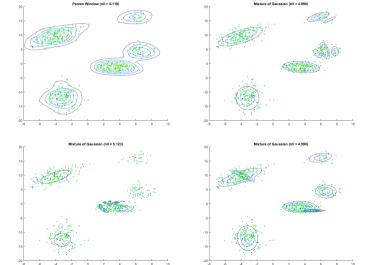
KDE vs. Mixture of Gaussian

• By fixing mean/covariance/k, we don't have to worry about local optima.



KDE vs. Mixture of Gaussian

ullet By fixing mean/covariance/k, we don't have to worry about local optima.



Mean-Shift Clustering

- Mean-shift clustering uses KDE for clustering:
 - Define a KDE on the training examples, and then for test example \hat{x} :
 - Run gradient descent to maximize p(x) starting from \hat{x} .
 - Clusters are points that reach same local minimum.
- https://spin.atomicobject.com/2015/05/26/mean-shift-clustering
- ullet Not sensitive to initialization, no need to choose k, can find non-convex clusters.
- Similar to density-based clustering from 340.
 - But doesn't require uniform density within cluster.
 - And can be used for vector quantization.
- "The 5 Clustering Algorithms Data Scientists Need to Know":
 - https://towardsdatascience.com/ the-5-clustering-algorithms-data-scientists-need-to-know-a36d136ef68

Continuous Mixture Models

• We've been discussing mixture models where z^i is discrete,

$$p(x^{i}) = \sum_{z^{i}=1}^{k} p(z^{i})p(x^{i} \mid z^{i} = c).$$

ullet We can also consider mixtures models where z^i is continuous,

$$p(x^i) = \int_{z^i} p(z^i) p(x^i \mid z^i = c) dz^i.$$

- Unfortunately, computing the integral might be hard.
 - But if both probabilities are Gaussian then it's straightforward.

Probabilistic PCA

• In 340 we discussed PCA, which approximates (centered) x^i by

$$x^i \approx W^T z^i$$
.

• In probabilistic PCA we assume that

$$x^i \sim \mathcal{N}(W^T z^i, \sigma^2 I), \quad z^i \sim \mathcal{N}(0, I).$$

• Continuous mixture integral will be marginal of a joint Gaussian, and gives

$$x^i \mid W \sim \mathcal{N}(0, W^T W + \sigma^2 I).$$

- Regular PCA is obtained as the limit of σ^2 going to 0.
 - Shows that PCA is just fitting a multivariate Gaussian with a restricted form for Σ .
 - Allows you to do things like mixture of PCAs.

Factor Analysis

- A related method for discovering latent factors is factor analysis (FA).
 - A standard tool and widely-used across science and engineering.

Trait	Description
O penness	Being curious, original, intellectual, creative, and open to new ideas.
Conscientiousness	Being organized, systematic, punctual, achievement- oriented, and dependable.
Extraversion	Being outgoing, talkative, sociable, and enjoying social situations.
A greeableness	Being affable, tolerant, sensitive, trusting, kind, and warm.
Neuroticism	Being anxious, irritable, temperamental, and moody.

https://new.edu/resources/big-5-personality-traits

- Historical applications are measures of intelligence and personality traits.
 - Some controversy, like trying to find factors of intelligence due to race.

(without normalizing for socioeconomic factors)

Factor Analysis

ullet FA approximates (centered) x^i by

$$x^i \approx W^T z^i$$
,

and assumes z^i and $x^i \mid z^i$ are Gaussian.

- Which should sound familiar...
- Are PCA and FA the same?
 - Both are more than 100 years old.
 - There are many online discussions about whether they are the same.
 - Some software packages run PCA when you call their FA method.
 - Some online discussions claiming they are completely different.

PCA vs. Factor Analysis

In probabilistic PCA we assume

$$x^i \mid z^i \sim \mathcal{N}(W^T z^i, \sigma^2 I), \quad z^i \sim \mathcal{N}(0, I),$$

and we obtain PCA as $\sigma \to 0$.

In FA we assume

$$x^i \mid z^i \sim \mathcal{N}(W^T z^i, \mathbf{D}), \quad z^i \sim \mathcal{N}(0, I),$$

where D is a diagonal matrix.

- The difference is that you can have a noise variance for each dimension.
 - So FA has extra degrees of freedom in variance of original variables.
 - In practice there often isn't a huge difference.

Summary

- Monotonicity of EM: EM is guaranteed not to decrease likelihood.
 - Very-recent results giving convergence rates.
- Kernel density estimation: Non-parametric density estimation method.
 - Allows smooth variations on histograms.
- Probabilistic PCA:
 - Continuous mixture models based on Gaussian assumptions.
 - Factor analysis extends probabilistic PCA with different noise in each dimension.
 - Very similar but not identical to PCA.
- Next time: the sad truth about rain in Vancouver.

Alternate View of EM as BCD

• We showed that given α the M-step minimizes in Θ the function

$$F(\Theta, \alpha) = -\mathbb{E}_{\alpha}[\log p(O, H \mid \Theta)] - \mathsf{entropy}(\alpha).$$

- The E-step minimizes this function in terms of α given Θ .
 - Setting $\alpha_H = p(H \mid O, \Theta)$ minimizes it.
- Note that F is not the NLL, but F and the NLL have same stationary points.
- From this perspective, we can view EM as a block coordinate descent method.
- This perspective is also useful if you want to do approximate E-steps.

Alternate View of EM as KL-Proximal

ullet Using definitions of expectation and entropy and lpha in the last slide gives

$$\begin{split} F(\Theta, \alpha) &= -\sum_{H} p(H \mid O, \theta^{t}) \log p(O, H \mid \Theta) + \sum_{H} p(H \mid O, \theta^{t}) \log p(H \mid O, \theta^{t}) \\ &= -\sum_{H} p(H \mid O, \theta^{t}) \log \frac{p(O, H \mid \theta)}{p(H \mid O, \theta^{t})} \\ &= -\sum_{H} p(H \mid O, \theta^{t}) \log \frac{p(H \mid O, \theta)p(O \mid \theta)}{p(H \mid O, \theta^{t})} \\ &= -\sum_{H} \log p(O \mid \Theta) - \sum_{H} p(H \mid O, \theta^{t}) \log \frac{p(H \mid O, \theta)}{p(H \mid O, \theta^{t})} \\ &= NLL(\Theta) + \mathsf{KL}(p(H \mid O, \theta^{t}) \mid\mid p(H \mid O, \theta)). \end{split}$$

- From this perspective, we can view EM as a "proximal point" method.
 - Classical proximal point method uses $\frac{1}{2} \|\theta^t \theta\|^2$, EM uses KL divergence.
- From this view we can see that EM doesn't depend on parameterization of Θ .
- If we linearize NLL and we multiply KL term by $1/\alpha_k$ (step-size), we get the natural gradient method.