

CPSC 540 Tutorial

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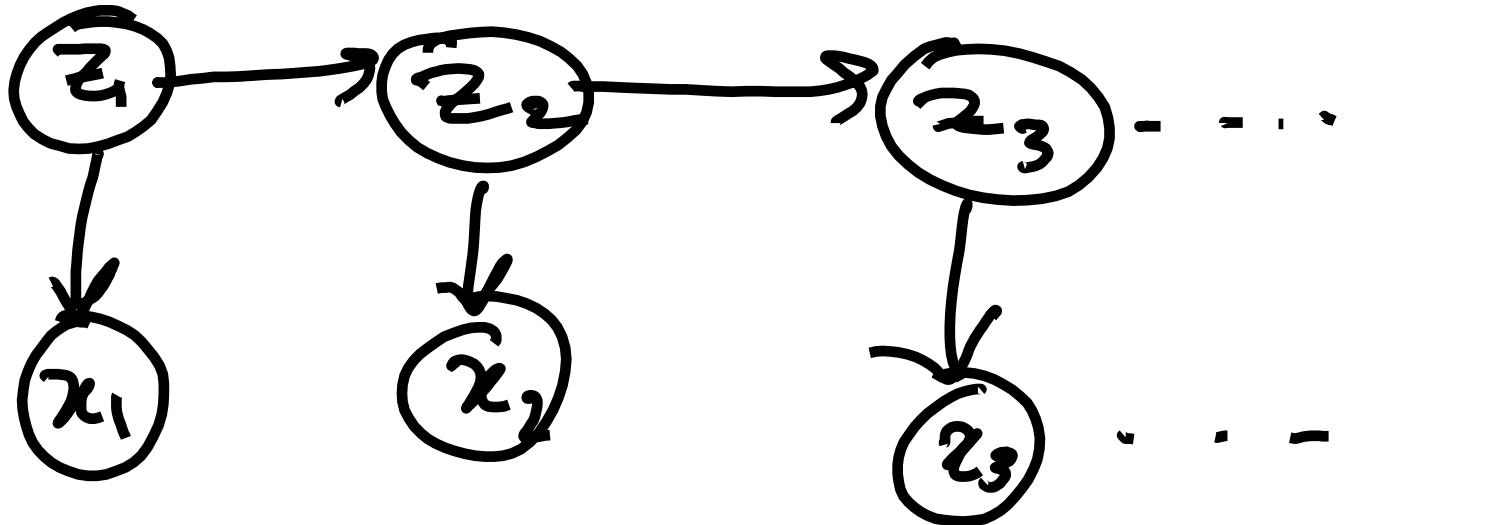
Motivation

- Action detection in video
 - You are given a video and asked to extract actions happening in it. What would you do?
- One way to solve it: using HMM
 - Each frame is known variable and action happening in there is latent!

HMM

- HMM can be used to model sequential data or stochastic processes.
 - Including two models
 - A discrete time, discrete state Markov chain with hidden states
$$z_t \in \{1, 2, \dots, K\}$$
 - An observation model $p(x_t | z_t)$
 - Observation model could be discrete or continuous
 - In discrete case, it could be observation matrix (we consider this case here)
$$p(x_t = l | z_t = k) = B_{kl}$$
 - In continuous case, we can use conditional Gaussian.
$$p(x_t | z_t = k; \theta) = N(x_t | \mu_k, \Sigma_k)$$
- Applications: Speech recognition, Activity recognition, Part of speech tagging, Gene finding, ...

Model Parameters



$P(z_1 = k) = \pi_k \rightarrow$ initial distribution

$P(z_t = k | z_{t-1} = j, \theta) = A_{jk} \rightarrow$ transition Prob.

$P(x_t = l | z_t = k) = B_{kl} \rightarrow$ observation Prob.

Inference with HMM

- Given parameters, we can do the following inference in HMM
 - Filtering: computing the belief state (online) $P(z_t | x_{1:t})$
 - Smoothing: computing the state offline $P(z_t | x_{1:T})$
 - Fixed lag smoothing: $P(z_{t-\ell} | x_{1:t})$
 - Prediction: $P(z_{t+h} | x_{1:t})$.
 - MAP estimation: $\overrightarrow{\underset{z_{1:T}}{\arg\max}} P(z_{1:T} | x_{1:T})$
 - Posterior sampling: $\tilde{z}_{1:T} \sim P(z_{1:T} | x_{1:T})$
 - Probability of evidence: $\tilde{P}(x_{1:T}) = \sum_{z_{1:T}} P(x_{1:T}, z_{1:T})$

Filtering using Forward Algorithm : $P(z_t=j|x_{1:t}, \theta)$

$$P(z_t=j|x_{1:t-1}) = \sum_i P(z_t=j|z_{t-1}=i) P(z_{t-1}=i|x_{1:t-1})$$

$$d_t(j) = P(z_t=j|x_{1:t}) = P(z_t=j|x_t, x_{1:t-1})$$

$$= \frac{1}{Z_t} P(x_t|z_t=j, x_{1:t-1}) P(z_t=j|x_{1:t-1})$$

$$\bar{z}_t = P(x_t|x_{1:t-1}) = \sum_j P(z_t=j|x_{1:t-1}) P(x_t|z_t=j)$$

$$P(z_t|x_t) = \frac{P(x_t|z_t) P(z_t)}{\sum_i P(x_t|z_i) P(z_i)}$$

Forward Backward algorithm for Smoothing

$$P(z_t=j | x_{1:T}) = P(z_t=j | x_{1:t}, x_{t+1:T}) = \frac{P(z_t, x_{t+1:T} | x_{1:t})}{P(x_{1:T})}$$

$$\propto P(z_t=j, x_{t+1:T} | x_{1:t}) = P(x_{t+1:T} | z_t, x_{1:t}) P(z_t | x_{1:t})$$

We know how to compute $\alpha_t(j) = P(z_t=j | x_{1:t})$

let $\beta_t(j) = P(x_{t+1:T} | z_t=j)$

$$\begin{aligned} &= \sum_i P(z_{t+1}=i, x_{t+1:T} | z_t=j) \\ &= P(z_{t+1}=i, x_{t+1}, x_{t+2:T} | z_t=j) \end{aligned}$$

$$\begin{aligned}\beta_+(j) &= \sum_i P(x_{t+2:T} | z_{t+1}=i, z_t \neq j, x_{t+1}) P(x_{t+1}, z_{t+1}=i | z_t=j) \\ &= \sum_i P(x_{t+2:T} | z_{t+1}=i) P(x_{t+1} | z_{t+1}=i) P(z_{t+1}=i | z_t=j)\end{aligned}$$

$\beta_{t+1}(i)$
 $B_{i\cdot}$
 A_{ji}

Base case : $\beta_T(j) = P(x_T | z_T=j) = B_{j\cdot}$

so $P(z_t | x_{1:T}) \propto \alpha_+(j) \beta_+(j)$

Two-slice smoothed marginals

Two slice marginal : $\xi_{t,t+1}^{(i,j)} \triangleq P(z_t=i, z_{t+1}=j | x_{1:T})$

$$\xi_{t,t+1}^{(i,j)} = P(z_t, z_{t+1} | x_{t+1:T}, x_{1:t})$$

$$\propto P(x_{t+1:T}, z_t, z_{t+1} | x_{1:t})$$

$$\propto P(x_{t+1:T} | z_t, z_{t+1}, x_{1:t}) P(z_t, z_{t+1} | x_{1:t})$$

$$\propto P(x_{t+1:T} | z_{\text{end}}) \underbrace{P(z_{t+1} | z_t)}_{\alpha_{t+1}^{ij}} \underbrace{P(z_t | x_{1:t})}_{\alpha_t(i)}.$$

$$\begin{aligned}
 P(x_{t+1:T} | z_{t+1}) &= P(x_{t+1}, x_{t+2:T} | z_{t+1}) \\
 &= P(x_{t+1} | z_{t+1}) P(x_{t+2:T} | z_{t+1})
 \end{aligned}$$

$B_j.$ $\beta_{t+1}(j)$

$$\text{So } \sum_{t,t+1} (i,j) \alpha_t(j) B_j. \beta_{t+1}(j) A_{ij}$$

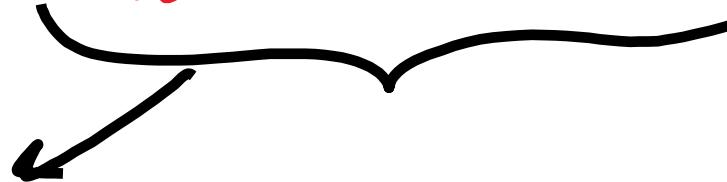
The Viterbi Algorithm

$$\text{goal : } z^* = \arg \max_{z_{1:T}} P(z_{1:T} | x_{1:T})$$

$$S_t(j) \triangleq \max_{z_{1:t-1}} P(z_{1:t-1}, z_t=j | x_{1:t})$$

$$P(z_{1:t-1}, z_t=j | x_{1:t-1}, x_t) \propto P(x_t | z_t) P(z_{1:t-1}, z_t=j | x_{1:t-1})$$
$$P(z_t=j, z_{1:t-1} | x_{1:t-1}) = P(z_t | z_{t-1}) P(z_{1:t-1} | x_{1:t-1})$$

$$(*) \rightarrow \delta_t(j) = \max_{z_{1:t-1}} P(z_{1:t-1} | x_{1:t-1}) A_{ij} B_j.$$



$$\max_{z_{1:t-2}} \max_i P(z_{1:t-2}, z_{t-1}=i | x_{1:t-1}) = \max_i \delta_{t-1}(i)$$

$$\text{So: } \delta_t(j) = \max_i \delta_{t-1}(i) A_{ij} B_j.$$

To keep track of most likely previous state

$$a_t(j) = \arg \max_i \delta_{t-1}(i) A_{ij} B_j.$$

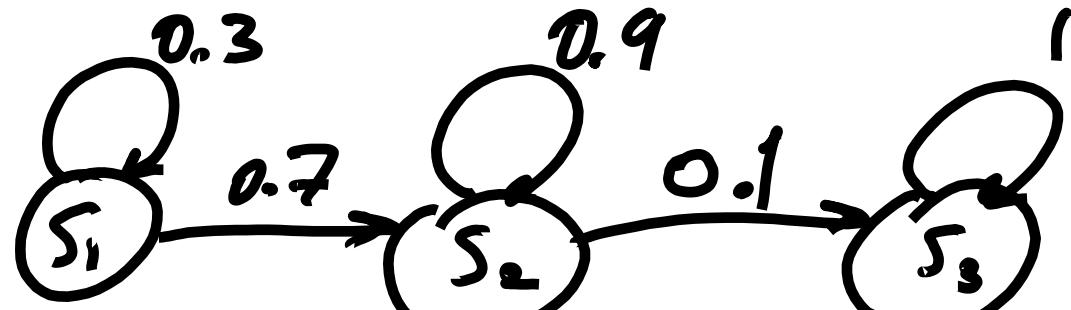
Base case : $\delta_i(z) = I_1 B_j$.

The algorithm terminates when we compute $Z_T^* = \arg \max_i \delta_T(i)$

To compute most probable sequence using tree back

$$Z_t^* = a_{t+1}(Z_{t+1}^*)$$

Example



$$A_{11} = 0.3, A_{12} = 0.7, A_{21} = 0.1, \dots$$

obs.	S_1	S_2	S_3
c_1	0.5	0	0
c_2	0.3	0	0
c_3	0.2	0.2	0
c_4	0	0.7	0.1
c_5	0	0.1	0
c_6	0	0	0.5
c_7	0	0	0.4

$$A = \begin{bmatrix} 1 & 2 & 3 \\ 0.3 & 0.7 & 0 \\ 0 & 0.9 & 0.1 \\ 0 & 0 & 1 \end{bmatrix}$$

Data: $C_1 C_3 C_4 C_6 = x_{1:4}$

$$\pi = (1, 0, 0)$$

$$\delta_1(1) = 1 \times 0.5 = 0.5, \delta_1(2) = 0, \delta_1(3) = 0 \quad t=1$$

$$\delta_2(1) = \max_i \delta_1(i) A(i, 1) B_{1C_3} \quad t=2$$

for $i=1$ $\delta_1(i) = 0.5$ but for $i=2, 3$ $\delta_i = 0$

$$so \delta_2(1) = \delta_1(1) * A(1, 1) * B_{1C_3} = 0.5 * 0.2 * 0.3 = 0.03$$

$\delta_2(2)$ like $\delta_2(1)$ just $\delta_1(1) \neq 0$

$$so \delta_2(2) = \delta_1(1) A(1, 2) B_{2C_3} = 0.5 * 0.7 * 0.2 = 0.07$$

like before $\delta_2(3) = \delta_1(0) A(1,3) \underbrace{B_3}_{\downarrow 0} / C_3 = 0$

So $a_2(1)=1$, $a_2(2)=1$ $a_2(3)=1 \text{ or } 2 \text{ or } 3$

Now

$t=3$

Since $B_{1C4}=0 \Rightarrow \delta_3(1) = \max_i \delta_2(i) A(i,1) B_{1C4} = 0$

$\delta_3(2) = \max \left\{ \begin{array}{l} \delta_2(1) A(1,2) B_{2C4} = 0.03 \times 0.7 \times 0.7 = 0.0147 \\ \delta_2(2) A(2,2) B_{2C4} = 0.07 \times 0.9 \times 0.7 = 0.0441 \\ \delta_2(3) A(3,2) B_{2C4} = 0 \end{array} \right.$

$\delta_3(2) = 0.0441$

$$S_3^{(3)=\max} \left\{ \begin{array}{l} \delta_2(1) A(1,3) B_{3c4} = 0 \\ \delta_2(2) A(2,3) B_{3c4} = 0.07 * 0.1 \times 0.1 = 0.0007 \\ \cancel{\delta_2(3) A(3,3) B_{3c4} = 0} \end{array} \right.$$

$$a_3(1) = 1, 2, 3, a_3(2) = 2, a_3(3) = 2$$

$$t=4:$$

$$\text{Since } B_{1c6} = 0 \Rightarrow \delta_4(1) = 0$$

$$\text{Since } B_{2c6} = 0 \Rightarrow \delta_4(2) = 0$$

$$\delta_4(3) = \max \left\{ \begin{array}{l} \delta_3(1) \ A(1,3) \ B_{3C6} = 0 \\ \delta_3(2) \ A(2,3) \ B_{3C6} = 0.0441 \times 0.1 \times 0.5 = 0.002205 \\ \delta_3(3) \ A(3,3) \ B_{3C6} = 0.007 \times 1 \times 0.5 = 0.0035 \end{array} \right.$$

$$\delta_4(3) = 0.0035$$

$$a_4(1) = 1, 2, 3 , \quad a_4(2) = 1, 2, 3 , \quad a_4(3) = 2$$

$$Z^* = 3 , \quad a_4(3) = 2 , \quad a_3(2) = 2 , \quad a_2(2) = 1 , \quad a_1(1) = 1$$

$$Z_{1234} = (1, 2, 2, 3)$$

Learning for HMM

- To do the inference we need to estimate the parameters of model
- When hidden variables are observed: MLE
- When hidden variables are not observed: EM

Fully observed case

A data sample from HMM: $x_{1:T}^i, z_{1:T}^i, \theta = (\pi, A, B)$

$$P(x, z | \theta) = \prod_{i=1}^N P(x^i, z^i | \theta) = \prod_{i=1}^N P(x^i | z^i, \theta) P(z^i | \theta)$$
$$= \prod_{i=1}^N \left[\prod_{t=1}^{T_i} P(x_t^i | z_t^i, \theta) \prod_{t=2}^{T_i} P(z_t^i | z_{t-1}^i, \theta) P(z_1^i | \theta) \right]$$

$$\sum_l B_{jl} = 1, \quad \sum_j \pi_j = 1 \quad \sum_j A_{ij} = 1$$

$$\log P(x, z | \theta) = \sum_{i=1}^N \sum_{t=1}^{T_i} \log P(x_t^i | z_t^i, \theta) + \sum_{i=1}^N \sum_{t=1}^{T_i} \log P(z_t^i | z_{t-1}^i, \theta) \\ + \sum_{i=1}^N \log P(z_i^i | \theta)$$

$$B_{jl} = P(z_t = l | z_t = j) \quad \text{let } j \in \{1..K\}, l \in \{1..S\}$$

$$P(x_t^i | z_t^i, \theta) = \prod_{j=1}^K \prod_{l=1}^S (P(x_t^i = l | z_t^i = j, \theta))^{I(x_t^i = l, z_t^i = j)}$$

$$\log P(x_t^i | z_t^i, \theta) = \sum_j \sum_l I(x_t^i = l, z_t^i = j) \log B_{jl}$$

$$\log P(z_t^i | z_{t-1}^i, \theta) = \sum_i \sum_{t=2} \sum_j \sum_k I(z_t^i = k, z_{t-1}^i = j) \log A_{jk}$$

$$\log P(z_1^i | \theta) = \sum_i \sum_{j=1}^K I(z_1^i = j) \log \pi_j$$

$$\frac{\partial \log P(X, Z | \theta)}{\partial B_{jl}} = \sum_{j=1}^J \sum_{i=1}^N \sum_{t=1}^{T_i} I(x_t^i = l, z_t^i = j) \log B_{jl} + \lambda_j (1 - \sum_{k=1}^K B_{jk})$$

$$= \sum_{i=1}^N \sum_{t=1}^{T_i} \frac{I(x_t^i = l, z_t^i = j)}{B_{jl}} - \lambda_j = 0$$

$$\Rightarrow B_{jl} = \sum_i \sum_t \frac{I(x_t^i = l, z_t^i = j)}{\lambda_j}$$

$$\sum_l \beta_{jl} = 1 \Rightarrow \lambda_j = \sum_l \sum_i \sum_t I(z_t^i = l, z_{t-1}^i = j)$$

In similar way, we can derive for π_j and a_{ij}

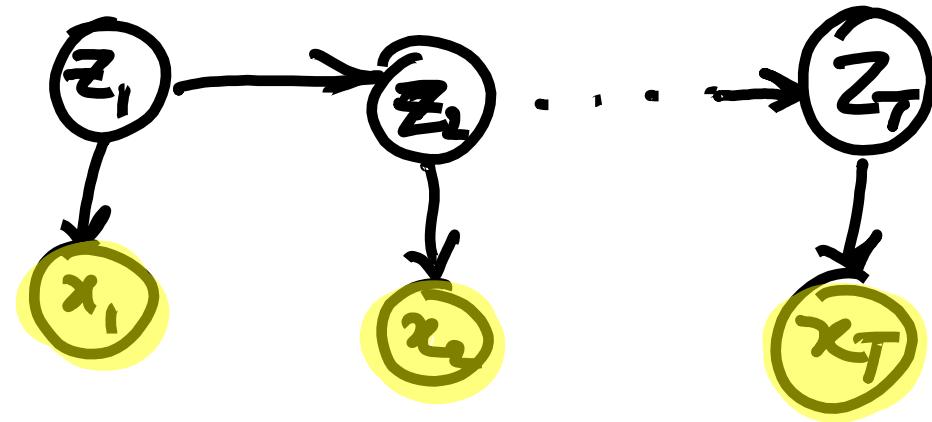
$$a_{jk} = \frac{\sum_i \sum_{t=2}^{T_i} I(z_t^i = k, z_{t-1}^i = j)}{\sum_k \sum_i \sum_{t=2}^{T_i} I(z_t^i = k, z_{t-1}^i = j)}$$

$$\pi_j = \frac{\sum_i I(z_i^i = j)}{\sum_{j=1}^k \sum_{i=1}^n I(z_i^i = j) = N}$$

Partially observed case

Observed data : $X_{1:T}$

Hidden data : $Z_{1:T}$



likelihood: $P(X | \theta) = ?$

X_i 's are not independent!

Use EM to estimate parameter and make model simpler
or likelihood

E-step:

$$\begin{aligned} \log P(x|\theta) &= \log \int p(x,z|\theta) dz = \log \int \frac{p(x,z|\theta)}{q(z|\theta^t)} q(z|\theta^t) dz \\ &\geq E_z \log p(x,z|\theta) - \text{Const (w.r.t. } \theta) \end{aligned}$$

$$Q(\theta, \theta^t) = E_z \log p(x,z|\theta) = E_z \log p(x|z, \theta) + E_z \log p(z|\theta)$$

$$P(x|z, \theta) = \prod_{i=1}^N \prod_{t=1}^T \prod_{j=1}^K \prod_{l=1}^S P(x_t^i = l | z_t^i = j)^{\mathbb{I}(x_t^i = l, z_t^i = j)}$$

$$P(z|\theta) = \prod_i^N \prod_{t=2}^{T_i} \prod_{j=1}^K \prod_{k=1}^K P(z_t^i = k | z_{t-1}^i = j; \theta)^{I(x_{t-k}, x_{t-1}^i = j)}$$

$$\times \prod_{i=1}^N \prod_{j=1}^K P(z_i^i = j | \theta)^{I(z_i^i = j)}$$

$$Q(\theta|\theta^t) = E_{\mathcal{Z}} \sum_i \sum_t \sum_{j \neq l} \mathbb{I}(z_t^i = j) I(x_t^i = 1) \log \beta_{jl}$$

$$+ E_{\mathcal{Z}} \sum_{i=1}^N \sum_{t=2}^{T_c} \sum_{j=1}^K \sum_{k=1}^K \mathbb{I}(z_t^i = k, z_{t-1}^i = j) \log A_{jk}$$

$$+ E_{\mathcal{Z}} \sum_{i=1}^N \sum_{j=1}^K \mathbb{I}(z_i^i = j) \log \pi_j$$

$$Q(\theta | \theta^t) = \sum_{i=1}^N \sum_{t=1}^{T_i} \sum_{j=1}^K \sum_{k=1}^S p(z_t^i | x^i; \theta^t) \mathbb{I}(x_t^i = k) \log \beta_{jk}$$

$$+ \sum_{i=1}^N \sum_{t=2}^{T_i} \sum_{j=1}^K \sum_{k=1}^K p(z_t^i = k, z_{t-1}^i = j | x^i; \theta^t) \log A_{jk}$$

$$+ \sum_{i=1}^N \sum_{j=1}^K p(z_1^i = j | x^i; \theta^t) \log \pi_j$$

$$P(z_t^i | x^i; \theta) = \gamma_{i,t}(j) \quad (\text{smoothing})!$$

$$P(z_t^i = k, z_{t-1}^i = j | x^i, \theta) = \gamma_{i,t}(j, k) \cdot (\text{Two step smooth marginals})$$

$$P(z_i^i = j | x^i, \theta) = \frac{P(x^i | z_i^i = j) P(z_i^i = j | \theta)}{\sum_j P(x^i | z_i^i = j) P(z_i^i = j | \theta)} = \gamma_i(j)$$

so:

$$\begin{aligned} Q(\theta | \theta^*) &= \sum_i \sum_t \sum_j \left[\gamma_{i,t}(j) I(x_t^i = l) \log B_{jl} + \sum_k \sum_{j,k} \left\{ \gamma_{i,t}(j, k) \log A_{jk} \right\} \right] \\ &+ \sum_i \sum_j \gamma_i(j) \log \pi_j \end{aligned}$$

M-step:

We compute for B_{jl} . Other parameters can be calculated similarly.

Extra constraint: $\forall j \sum_l B_{jl} = 1, \sum_k A_{jk} = 1, \sum_j \pi_j = 1$

$$\begin{aligned} S_0: \frac{\partial Q}{\partial B_{jl}} &= \frac{\partial}{\partial B_{jl}} \left\{ \sum_{i=1}^N \sum_{t=1}^{T_i} \gamma_{i,t}(j) I(x_t^i = l) \log B_{jl} + \lambda_j (1 - \sum_l B_{jl}) + \text{const} \right\} \\ &= \sum_{i=1}^N \sum_{t=1}^{T_i} \frac{\gamma_{i,t}(j) I(x_t^i = l)}{B_{jl}} - \lambda_j = 0 \Rightarrow B_{jl} = \frac{\sum_{i=1}^N \sum_{t=1}^{T_i} \gamma_{i,t}(j) I(x_t^i = l)}{\lambda_j} \end{aligned}$$

$$\sum_k \beta_{jk} = 1 \Rightarrow \gamma_j = \sum_l \sum_i \sum_t \gamma_{i,t}(j) I(x_t^i = l) = N_j$$

Exercise: drive the updates for π_j, A_{jk} ?

Exercise: drive MLE and E-M when $p(x|z)$ is Gaussian?