

# CS340 Machine learning Midterm review

# Topics

- Bayesian statistics
  - Information theory
  - Decision theory

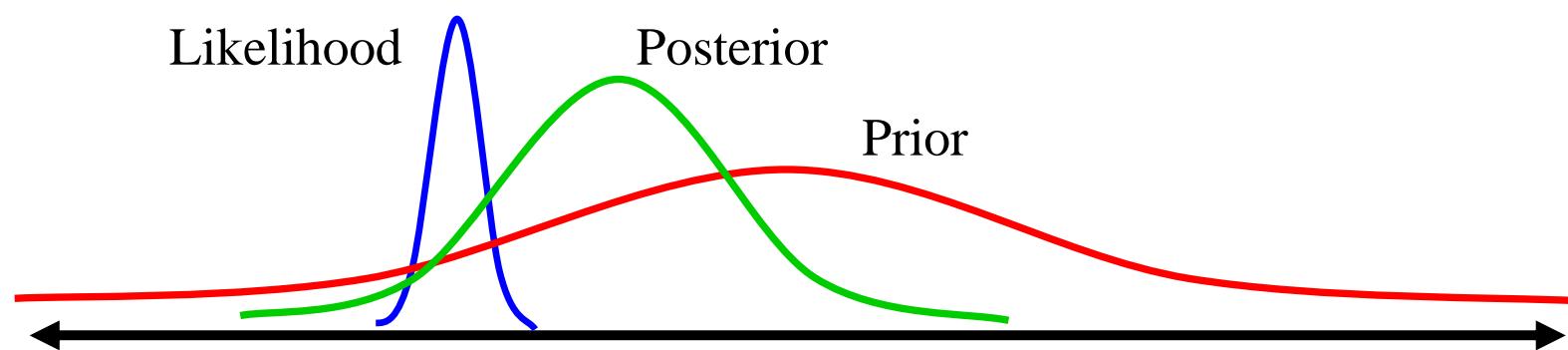
kNN not on exam

Sampling distributions (confidence intervals etc) not on exam

# Bayesian belief updating

$$p(h | d) = \frac{p(d | h)p(h)}{\sum_{h' \in H} p(d | h')p(h')}$$

Posterior probability → Likelihood ↓ Prior probability



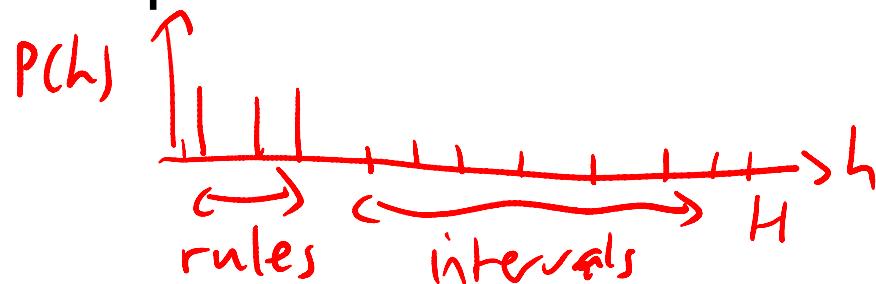
Bayesian inference = Inverse probability theory

# Number game

- Data:  $x_i \in \{1, \dots, 100\}$ .
- Hypothesis space:  $h \in \{1, \dots, H\}$      $H \sim 3000$   
(rules + intervals)
- Likelihood: strong sampling model

$$p(D|h) = \left[ \frac{1}{|h|} \right]^n I(x_1, \dots, x_n \in h)$$

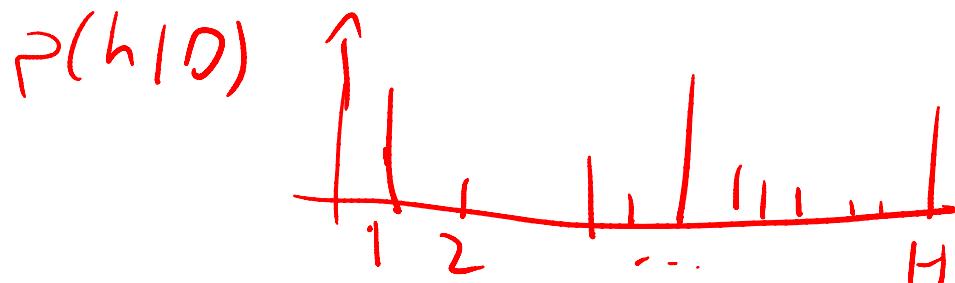
- Prior: piecewise uniform histogram



$$p(h) = \frac{\lambda I(h \in \text{rules}) + (1 - \lambda)I(h \in \text{intervals})}{H}$$

# Number game

- Posterior: histogram



- Posterior predictive: histogram



# Coin tossing

- Data:  $x_i \in \{0,1\}$
- Hypothesis space:  $\theta$  in  $[0,1]$
- Likelihood: Bernoulli

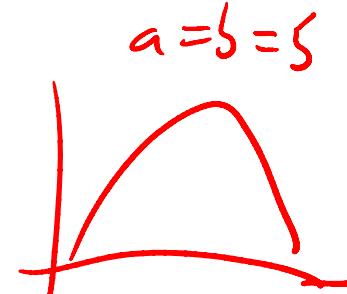
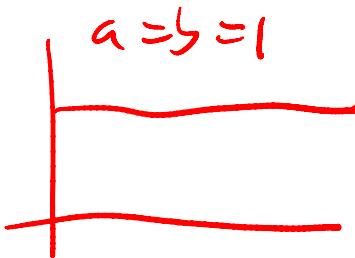
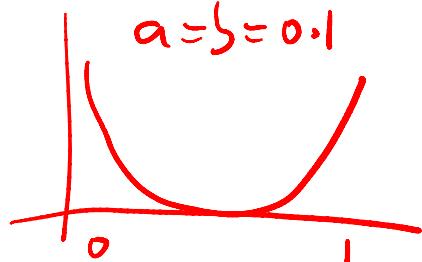
$$p(D|\theta) = \prod_{i=1}^n \theta^{I(x_i=1)} (1-\theta)^{I(x_i=0)} = \theta^{N_1} (1-\theta)^{N_0}$$

$$N_j = \sum_{i=1}^n I(x_i = j)$$

$$E[\theta] = \frac{a}{a+b}$$

- Prior: Beta

$$p(\theta) = Beta(\theta|a, b)$$



# Coin tossing

- Posterior: beta

$$p(\theta|D) = Beta(\theta|a + N_1, b + N_0)$$

- Posterior predictive: two-element histogram

$$\begin{aligned} p(X = 1|D) &= \int p(X = 1|\theta)p(\theta|D)d\theta \\ &= \int \theta Beta(\theta|a', b')d\theta = \frac{a'}{a' + b'} \end{aligned}$$

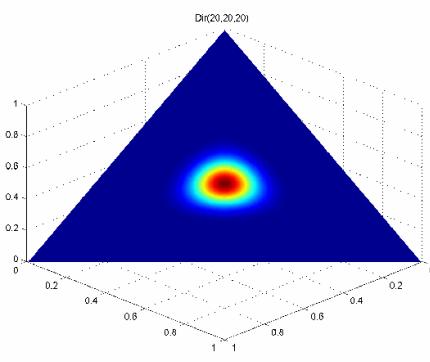
# Dice rolling

- Data:  $x_i \in \{1, \dots, K\}$
- Hypothesis space:  $(\theta_1, \dots, \theta_K) \in [0,1]^K$  st  $\sum_k \theta_k = 1$   
(probability simplex)
- Likelihood: Multinomial

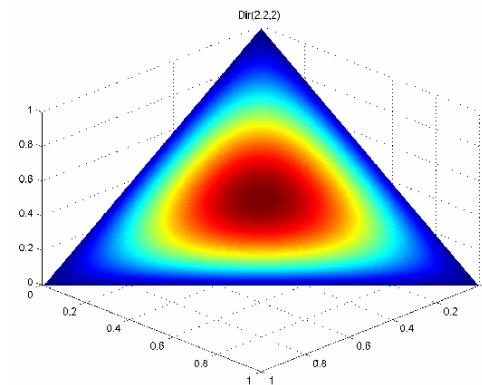
$$p(D|\boldsymbol{\theta}) = \prod_k \theta_k^{N_k}$$

- Prior: Dirichlet

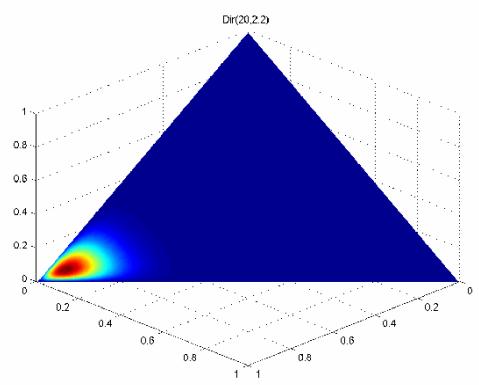
$$p(\boldsymbol{\theta}) = Dir(\boldsymbol{\theta}|\boldsymbol{\alpha})$$



(20,20,20)



(2,2,2)



(20,2,2)

# Dice rolling

- Posterior: Dirichlet

$$p(\boldsymbol{\theta}|D) = Dir(\boldsymbol{\theta}|\boldsymbol{\alpha} + \mathbf{N})$$

- Posterior predictive: K-element histogram

$$p(X = k|D) = E[\theta_k|D] = \frac{\alpha_k + N_k}{\sum_{k'} \alpha_{k'} + N_{k'}}$$

# Real values

- Data:  $x_i \in \mathbb{R}$
- Hypothesis space:  $\mu \in \mathbb{R}$  ( $\lambda$  known)
- Likelihood: Gaussian

$$\begin{aligned} p(D|\mu) &= \frac{1}{(2\pi)^{n/2}} \lambda^{n/2} \exp \left( -\frac{\lambda}{2} \sum_{i=1}^n (x_i - \mu)^2 \right) \\ &= \frac{1}{(2\pi)^{n/2}} \lambda^{n/2} \exp \left( -\frac{\lambda}{2} \left[ n(\mu - \bar{x})^2 + \sum_{i=1}^n (x_i - \bar{x})^2 \right] \right) \end{aligned}$$

- Prior: Gaussian

$$p(\mu) = \mathcal{N}(\mu|\mu_0, \lambda_0^{-1})$$

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

# Real values

- Posterior: Gaussian

$$p(\mu|D) = \mathcal{N}(\mu|\mu_n, \lambda_n^{-1})$$

$$\lambda_n = \lambda_0 + n\lambda \quad \text{Precisions add}$$

$$\mu_n = \frac{\bar{x}n\lambda + \mu_0\lambda_0}{\lambda_n} \quad \text{Convex combination}$$

- Posterior predictive: Gaussian

$$\begin{aligned} p(x|D) &= \int p(x|\mu)p(\mu|D)d\mu \\ &= \int \mathcal{N}(x|\mu, \sigma^2)\mathcal{N}(\mu|\mu_n, \sigma_n^2)d\mu \\ &= \mathcal{N}(x|\mu_n, \sigma_n^2 + \sigma^2) \end{aligned}$$

↑                      ↑  
Uncertainty about  $\mu$       noise

# Real values

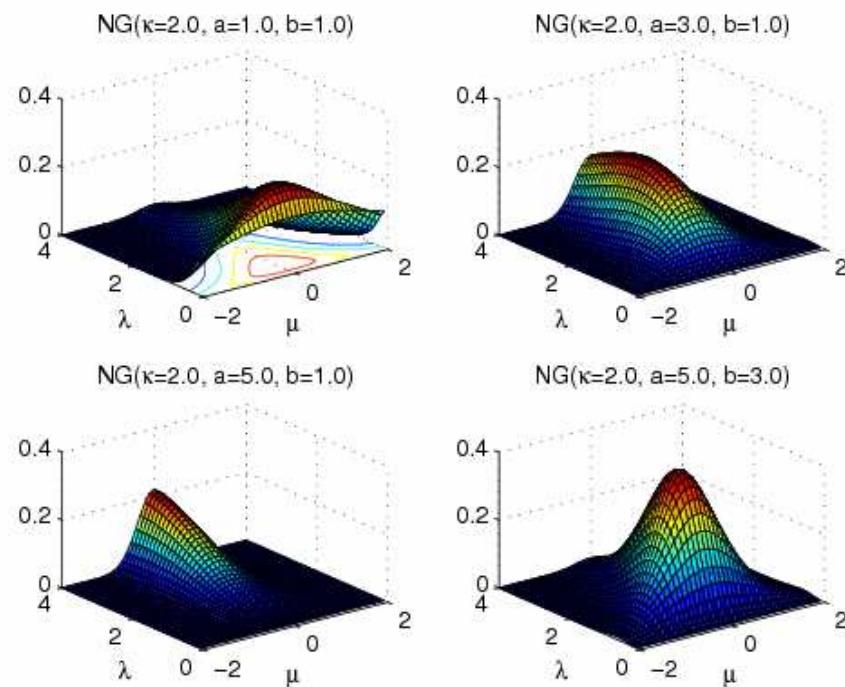
- Data:  $x_i \in \mathbb{R}$
- Hypothesis space:  $\mu \in \mathbb{R}, \lambda \in \mathbb{R}^+$
- Likelihood: Gaussian

$$\begin{aligned} p(D|\mu) &= \frac{1}{(2\pi)^{n/2}} \lambda^{n/2} \exp \left( -\frac{\lambda}{2} \sum_{i=1}^n (x_i - \mu)^2 \right) \\ &= \frac{1}{(2\pi)^{n/2}} \lambda^{n/2} \exp \left( -\frac{\lambda}{2} \left[ n(\mu - \bar{x})^2 + \sum_{i=1}^n (x_i - \bar{x})^2 \right] \right) \end{aligned}$$

# Real values

- Prior: Normal Gamma

$$\begin{aligned} NG(\mu, \lambda | \mu_0, \kappa_0, \alpha_0, \beta_0) &\stackrel{\text{def}}{=} \mathcal{N}(\mu | \mu_0, (\kappa_0 \lambda)^{-1}) Ga(\lambda | \alpha_0, \text{rate} = \beta_0) \\ &\propto \lambda^{\frac{1}{2}} \exp\left(-\frac{\kappa_0 \lambda}{2}(\mu - \mu_0)^2\right) \lambda^{\alpha_0 - 1} e^{-\lambda \beta_0} \end{aligned}$$



# Real values

- Posterior: Normal Gamma

$$p(\mu, \lambda | D) = NG(\mu, \lambda | \mu_n, \kappa_n, \alpha_n, \beta_n)$$

$$\mu_n = \frac{\kappa_0 \mu_0 + n \bar{x}}{\kappa_0 + n}$$

$$\kappa_n = \kappa_0 + n$$

$$\alpha_n = \alpha_0 + n/2$$

$$\beta_n = \beta_0 + \frac{1}{2} \sum_{i=1}^n (x_i - \bar{x})^2 + \frac{\kappa_0 n (\bar{x} - \mu_0)^2}{2(\kappa_0 + n)}$$

- Posterior predictive: student T (long-tailed Gaussian)

$$p(x | D) = t_{2\alpha_n}(x | \mu_n, \frac{\beta_n(\kappa_n + 1)}{\alpha_n \kappa_n})$$

# Posterior summaries

- Common to quote the posterior mean  $E[\theta|D]$  as a point estimate
- 95% Credible interval  $(\ell(D), u(D))$   
$$p(\ell \leq \theta \leq u|D) \geq 0.95$$
- Can also summarize using samples from posterior  
 $\theta^s \sim p(\theta|D)$

# MAP estimation

- We will often use the posterior mode as an approximation to the full posterior.

$$\begin{aligned}\hat{\theta}_{MAP} &= \arg \max_{\theta} p(\theta|D) \\ &= \arg \max_{\theta} \log p(D|\theta) + \log p(\theta) \\ p(\theta|D) &\approx \delta(\theta - \hat{\theta}_{MAP})\end{aligned}$$

- This ignores uncertainty in our estimate, and will result in overconfident predictions.
- However, it is often computationally much cheaper than a fully Bayesian solution.
- If  $p(\theta) \propto 1$  (uninformative prior), then MAP = MLE.

# Topics

- Bayesian statistics
- Information theory
- Decision theory

# Information theory

- Entropy = min num bits to encode samples from  $p(x)$  using optimal code (and knowledge of  $p(x)$ )

$$H(X) = - \sum_x p(x) \log_2 p(x)$$

- KL divergence = extra num bits to encode samples coming from  $p(x)$  using code based on  $q(x)$

$$\begin{aligned} KL(p||q) &= \sum_x p(x) \log \frac{p(x)}{q(x)} \\ &= \underbrace{- \sum_x p(x) \log q(x)}_{\text{Cross entropy from } p \text{ to } q} - H(p) \end{aligned}$$

Cross entropy from  $p$  to  $q$

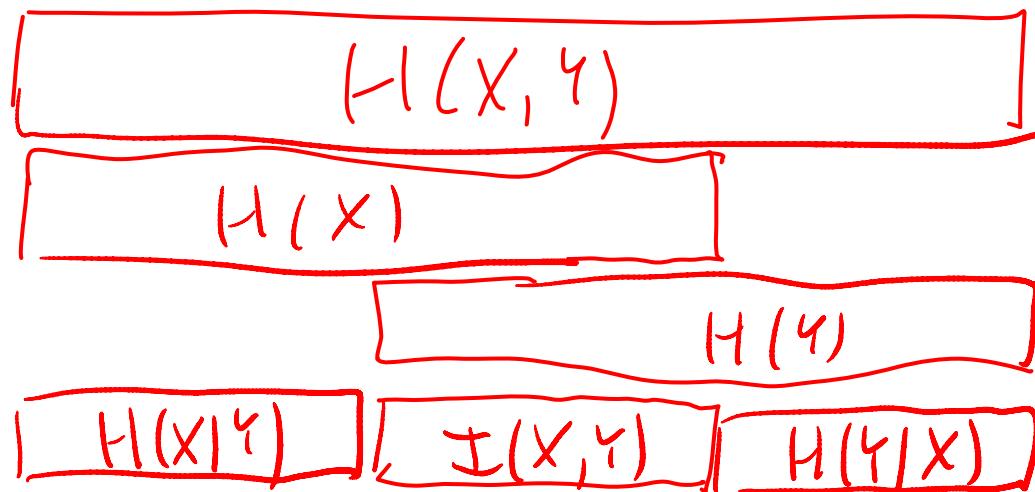
# Mutual information

- $I(X, Y)$  is how much our uncertainty about  $Y$  decreases when we observe  $X$

$$\begin{aligned} I(X, Y) &\stackrel{\text{def}}{=} \sum_y \sum_x p(x, y) \log \frac{p(x, y)}{p(x)p(y)} = KL(p(x, y) || p(x)p(y)) \\ &= -H(X, Y) + H(X) + H(Y) \\ &= H(X) - H(X|Y) = H(Y) - H(Y|X) \end{aligned}$$

- Hence

$$H(X, Y) = H(X|Y) + H(Y|X) + I(X, Y)$$



Mackay 9.1

# Topics

- Bayesian statistics
- Information theory
- Decision theory

# Bayesian decision theory

- Pick action  $\hat{\theta}(D)$  to minimize expected loss wrt current belief state  $p(\theta|D)$

$$\hat{\theta}(D) = \arg \min_{\hat{\theta}} EL(\theta, \hat{\theta}) = \arg \min_{\hat{\theta}} \int L(\theta, \hat{\theta}) p(\theta|D) d\theta$$

- Squared error (L2) loss

$$L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2 \quad \text{Posterior mean}$$
$$\frac{d}{d\hat{\theta}} EL(\theta, \hat{\theta}) = 0 \Rightarrow \hat{\theta}(D) = E[\theta|D]$$

- Zero-one loss

$$L(\theta, \hat{\theta}) = \delta(\theta - \hat{\theta}) \quad \text{Posterior mode}$$
$$\frac{d}{d\hat{\theta}} EL(\theta, \hat{\theta}) = 0 \Rightarrow \hat{\theta}(D) = \arg \max_{\theta} p(\theta|D)$$

# Decision theory: classifiers

- Given belief state  $p(y|x)$ , pick action  $\hat{y}(x)$  to minimize expected loss

$$\hat{y}(x) = \arg \min_{\hat{y}} EL(y, \hat{y}) = \arg \min_{\hat{y}} \sum_y L(y, \hat{y}) p(y|x)$$

		state	$y$
		1	2
action $\hat{y}$	1	True positive $\lambda_{11} = 0$	False positive $\lambda_{12}$
	2	False negative $\lambda_{21}$	True negative $\lambda_{22} = 0$

$$\hat{y}(x) = 1 \text{ iff } \frac{p(Y=1|x)}{p(Y=2|x)} > \frac{\lambda_{12}}{\lambda_{21}}$$

# Decision theory: model selection

- Given belief state  $p(m|D)$ , pick model  $\hat{m}(D)$  to minimize expected loss
- For 0-1 loss, pick most probable model

$$m^*(D) = \arg \max_m p(m|D)$$

$$p(m|D) = \frac{p(m)p(D|m)}{p(D)}$$

$$p(D) = \sum_{m \in \mathcal{M}} p(m)p(D|m)$$

# Bayes factors

- To compare 2 models with equal priors, use the Bayes factor (c.f. likelihood ratio)

$$BF(i, j) = \frac{p(D|m_i)}{p(D|m_j)}$$

- The marginal likelihood  $p(D|m)$  is the probability that model m can generate D using parameters sampled from its prior

$$p(D|m) = \int p(D|\theta, m)p(\theta|m)d\theta$$

- This automatically penalizes complex models (Occam's razor)

# Predicting the future

- Consider predicting  $y=x_{n+1}$  given  $x_{1:n}$ .
- Use a loss function that measures your surprise

$$L(m, y) = -\log p(y|m)$$

- Pick  $m$  to minimize expected loss (risk)

$$\begin{aligned} R(m) &= \int p(y|x_{1:n}) L(m, y) dy \\ &= \int -p(y|x_{1:n}) \log p(y|x_{1:n}, m) = E_y f_m(y, x_{1:n}) \end{aligned}$$

- To minimize this cross entropy, we should pick the model whose predictions  $p(y|x_{1:n}, m)$  come closest to our predictions of the future, given by this Bayes model average

$$p(y|x_{1:n}) = \sum_{m \in \mathcal{M}} p(y|m, x_{1:n}) p(m|x_{1:n})$$

# Cross validation

- If we don't think the true model is in our model class  $\mathcal{M}$ , we can approximate  $p(y|x_{1:n})$  empirically.
- Leave one out cross validation (LOOCV) uses  $x_i$  as test and  $x_{-i}$  as training, and averages over  $i$

$$E_y f_m(y, x_{1:n}) \approx \frac{1}{n} \sum_{i=1}^n f_m(x_i, x_{-i})$$

- We then pick the model  $m$  with the minimal empirical cross entropy loss.
- We can reduce the variance of this estimate using larger splits, eg 10-fold cross validation.