Alireza Shafaei - April 2016









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 $\mathcal{P}(\mathcal{D}) = \int_{\Theta} \mathcal{P}(\mathcal{D}|\theta)\pi(\theta) \,\mathrm{d}\theta$



• Previously we looked at the general problem of handling **high-dimensional integrals** and **unnormalized probability** functions.

$$\mathcal{P}(x) = \frac{1}{Z} p^*(x)$$



- Rejection Sampling Given $p^*(x), q(x), M \text{ s.t.} \frac{p^*(x)}{q(x)} \le M \forall x.$ $x \sim q(x)$
 - Accept x with probability

 $\frac{p^*(x)}{M \cdot q(x)}$



$$\int_x f(x) p(x) \,\mathrm{d}x$$



$$\int_{x} f(x)p(x) \, \mathrm{d}x = \frac{\int_{x} f(x)p^{*}(x) \, \mathrm{d}x}{\int_{x} p^{*}(x) \, \mathrm{d}x}$$



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$$\int_{x} f(x)p(x) \, \mathrm{d}x \approx \frac{\sum_{l=1}^{L} f(x_{l}) \frac{p^{*}(x_{l})}{q(x_{l})}}{\sum_{l=1}^{L} \frac{p^{*}(x_{l})}{q(x_{l})}}$$



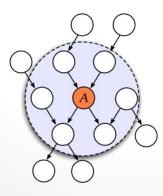
• Markov chain Monte Carlo



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 - Reversible Jump MCMC (non-parametric)





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$$\mathcal{P}(\theta|\mathcal{D}) = \frac{\mathcal{P}(\mathcal{D}|\theta)\pi(\theta)}{\mathcal{P}(\mathcal{D})} \propto \mathcal{P}(\mathcal{D}|\theta)\pi(\theta)$$

• What if we can't calculate $\mathcal{P}(\mathcal{D}|\theta)$?





• Apparently applies to a lot of problems in biology.



The Problem

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- Given a parameter θ you can simulate the execution.



The Problem

- Apparently applies to a lot of problems in biology.
- Given a parameter θ you can simulate the execution.
- $\mathcal{P}(\mathcal{D}|\theta)$ Could be intractable or simply no mathematical derivation of it exists.



1. Draw $\theta \sim \pi(\theta)$



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 - We want to explore the space to accept more often.
 - Sampling from the prior does not incorporate current observations.
- How do we choose $\rho(\cdot, \cdot), \mathcal{S}(\cdot), \epsilon$?



Approximate MCMC

1. Propose $\theta' \sim Q(\theta'|\theta)$



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Approximate MCMC

- 1. Propose $\theta' \sim Q(\theta'|\theta)$
- 2. Simulate $\tilde{\mathcal{D}} \sim \mathcal{P}(.|\theta')$
- 3. If $\rho(\mathcal{S}(\mathcal{D}), \mathcal{S}(\tilde{\mathcal{D}})) < \epsilon$
 - a. Accept θ' with probability

$$\min(1, \frac{\pi(\theta')Q(\theta'|\theta)}{\pi(\theta)Q(\theta|\theta')})$$



• Let's assume $\theta = (\theta_1, \theta_2)$ $\circ \mathcal{P}(\theta_1 | \mathcal{D}, \theta_2)$ is known. $\circ \mathcal{P}(\theta_2 | \mathcal{D}, \theta_1)$ is unknown.



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- 1. $\theta_1^{t+1} \sim \mathcal{P}(\theta_1 | \mathcal{D}, \theta_2)$ 2. $\theta_2^* \sim \pi(\theta_2)$ $\circ \quad \tilde{\mathcal{D}} \sim \mathcal{P}(.|\theta_1^{t+1}, \theta_2^*)$



- Let's assume $\theta = (\theta_1, \theta_2)$ • $\mathcal{P}(\theta_1 | \mathcal{D}, \theta_2)$ is known. • $\mathcal{P}(\theta_2 | \mathcal{D}, \theta_1)$ is unknown.
- $\begin{array}{ll} 1. & \theta_1^{t+1} \sim \mathcal{P}(\theta_1 | \mathcal{D}, \theta_2) \\ 2. & \theta_2^* \sim \pi(\theta_2) \\ & \circ & \tilde{\mathcal{D}} \sim \mathcal{P}(. | \theta_1^{t+1}, \theta_2^*) \\ & \circ & \rho(\mathcal{S}(\mathcal{D}), \mathcal{S}(\tilde{\mathcal{D}})) < \epsilon \Rightarrow \theta_2^{t+1} = \theta_2^* \\ & \circ & \text{else go to } 2. \end{array}$



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- Likelihood is not needed.
- Easy to implement and parallelize.

Cons

- Lot's of tuning.
- For complex problems, sampling from the prior is frustrating because it does not incorporate the evidence.
- How good is our approximation?



Thank you!



References

- 1. Wilkinson, Richard, and Simon Tavaré. "Approximate Bayesian Computation: a simulation based approach to inference."
- 2. https://en.wikipedia.org/wiki/Gibbs sampling
- 3. Barber, David. Bayesian reasoning and machine learning. Cambridge University Press, 2012.

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