# Causal Inference: <br> Classical Approaches 

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UBC MLRG

## Outline

- Potential Outcomes
- Confounding and Causal DAGs
- Granger Causality
- ICA for Causal Discovery


## Associational Inference

- Universe $U$
- For each unit $u \in U$ :
- Attribute variable $X(u)$
- Observed variable $Y(u)$
- Inference:
- $P(Y=y \mid X=x)$


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## Causal Inference

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- For each unit $u \in U$ :
- Treatment variable $T(u) \in\{1,0\}$
- Potential outcome $Y_{1}(u), Y_{0}(u)$
- Inference:
- $Y_{1}(u)-Y_{0}(u)$


## Rubin's Framework

- For each unit $u \in U$ :
- Treatment variable $T(u) \in\{1,0\}$
- Potential outcomes $Y_{1}(u), Y_{0}(u)$
- the outcome that would be observed if treatment was set to $T=0$ or 1 , on the same unit.
- (before)
- If $T(u)$ is set to 1
- $Y_{1}(u)$ is the observed outcome
- $Y_{0}(u)$ is the counterfactual outcome
- (after)


## Causal Effects

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- Abbreviated as $Y_{1}$ and $Y_{0}$
- Fundamental Problem of Causal Inference
- It is impossible to observe both $Y_{1}$ and $Y_{0}$ on the same unit, and therefore it is impossible to observe the causal effect.


## THE END

## Scientific solution to the Fundamental Problem

- Assume temporal stability and causal transience
- The value of $Y_{0}$ does not depend on when $T=0$ is applied and measured.
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- Causal effect can then be computed using $Y_{1}\left(u_{1}\right)-Y_{0}\left(u_{2}\right)$.
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- Implies the constant effect assumption: $Y_{1}(u)-Y_{0}(u)$ is the same for all $u \in U$.
- It's very difficult to argue that these are valid...


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- Average causal effect of $T=1$ (relative to $T=0$ ) over $U$ :
- $\mathbb{E}\left(Y_{1}-Y_{0}\right)=\mathbb{E}\left(Y_{1}\right)-\mathbb{E}\left(Y_{0}\right)$
- Imagine parallel universes with the same population...
- Can't observe this.
- Observed data can only give us information about the average of the outcome over $u \in U$ exposed to $T=t$.
- $\mathbb{E}\left(Y_{1} \mid T=1\right)-\mathbb{E}\left(Y_{0} \mid T=0\right)$


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- $\mathbb{E}\left(Y_{1} \mid T=1\right)-\mathbb{E}\left(Y_{0} \mid T=0\right)$
- In general, $\mathbb{E}\left(Y_{t}\right) \neq \mathbb{E}\left(Y_{t} \mid T=\mathrm{t}\right)$


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- $\mathbb{E}\left(Y_{1} \mid T=1\right)-\mathbb{E}\left(Y_{0} \mid T=0\right)$
- In general, $\mathbb{E}\left(Y_{t}\right) \neq \mathbb{E}\left(Y_{t} \mid T=\mathrm{t}\right)$
- Independence assumption hold via randomized treatment assignment allows equality to hold, which lets us compute the ACE above.


## Other assumptions

- Stable Unit Treatment Value Assumption (SUVTA)
- No interference: units do not interact with each other.
- One version of treatment.
- Consistency
- The potential outcome $Y_{t}$ is equal to the observed outcome if the actual treatment received is $T=t$.
- Positivity
- $\mathbb{P}(T(u)=t)>0$ for all $t$ and $u$.


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- Positivity
- $\mathbb{P}(T(u)=t)>0$ for all $t$ and $u$.
- Ignorability (aka no unmeasured confounders assumption)
- $Y_{0}, Y_{1} \perp \mathrm{~T} \mid \mathrm{X}$
- Among people with the same features $X$, we can think of treatment $T$ as being randomly assigned.


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## DAGs

- Useful for identifying dependencies and ways to factor and simplify the joint distribution.
- $p\left(x_{1}, \ldots, x_{n}\right)=\prod_{\{i=1\}}^{n} p\left(x_{i} \mid x_{\{p a(i)\}}\right)$


Firing squad example [Pearl, 2018]

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- Two variables $A$ and $B$ are d-separated by a set of variables $Z$ if $A$ and $B$ are conditionally independent given $Z$.
- $p(A, B \mid Z)=p(A \mid Z) p(B \mid Z)$
- Chain



## Causal DAGs

- DAGs where directions of the edges represent causal relationships.
- In contrast to Rubin's potential outcome framework, this is a structural approach to causal inference which Pearl advocates.
- They are shown to be mathematically equivalent.


Firing squad example [Pearl, 2018]

## Intervention and Pearl's do-calculus

- $d o()$ operator signals an intervention on a variable.
- Replace that variable with the actual value that we assign.
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- $d o()$ operator signals an intervention on a variable.
- Replace that variable with the actual value that we assign.
- Removes all incoming edges to that node.
- Instead of $p(D \mid A=0)$
- We want $p(D \mid d o(A=0))$
- The causal effect of $A=0$ on $D$.



## Confounding

- Confounders: variables that influences both treatment and outcome.
- Want: identify a set of variables so that ignorability holds.
- We don't need to identity specific confounders
- We just need to be able to control for confounding.
- Need to block backdoor paths from $T$ to $Y$.



## Frontdoor paths



- We are not concerned about frontdoor paths.
- We don't want to control anything along the frontdoor paths.
- Unless we care about the magnitude of the causal effect...


## Backdoor paths

- Begins with a parent of $T$ and ends at $Y$.
- Need to control these paths as they confound our causal effect.
- How?
- Identify the set of variables that blocks all backdoor paths from $T$ to $Y$.



## Backdoor criterion

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1. it blocks all backdoor paths from $T$ to $Y$, and
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- Backdoor Adjustment:
- If a set of variables $C$ satisfies the backdoor criterion relative to $T$ ane $Y$, then the causal effect of $T$ on $Y$ is given by

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- $\mathbb{P}(Y \mid d o(T=t))=\sum_{c \in C} \mathbb{P}(Y \mid T=t, c) \mathbb{P}(c)$.
- In Rubin's framework, this is equivalent to the ignorability assumption:
- Treatment assignment is effectively randomized given $C$.


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- Hypothesis test:
- $\mathbb{P}(Y(t+1) \mid I(t)) \neq \mathbb{P}\left(Y(t+1) \mid I_{\{-x\}}(t)\right)$
- $\quad I(t)$ all information up to time $t$
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- Steps:
- $y_{t}=a_{0}+a_{1} y_{\mathrm{t}-1}+a_{2} y_{t-2}+\epsilon_{t}$



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- $H_{0}$ : All $b_{1}=\cdots=b_{p}=0$

- $H_{1}$ : At least one is non-zero.
- Null hypothesis: $X$ does not Granger cause $Y$ iff no lagged values of $x$ are retained.


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## Holland

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## Pearl



## Judea Pearl @yudapearl • Jan 24

Replying to @nntaleb and @HarryDCrane
1/ This one is easy. In 1991, I had a quiet dinner with Clive Granger in Uppsala, Sweden. Between the 2nd and 3rd glass of wine, he confessed to me that he feels embarrassed by the name: "Granger causality", since it has nothing to do with causality, but he can't stop people from
$\bigcirc 1$
〔】 3
O 54


Judea Pearl @yudapearl • Jan 24
$2 / 3$ using it; they need some way to express what they wish to estimate. I think we should honor him by echoing his understanding. An easy way to see that GC has nothing to do with causality is to look at the defining equations and note that they comprise only conditional
Q 1
$\uparrow \downarrow$
O 21
$\uparrow$

Judea Pearl @yudapearl • Jan 24
$3 / 3$ probabilities, no do $(x)$ expressions, nor counterfactual terms $Y \_x$. Bingo! We are done! Whenever a concept is defined in terms of a distribution of observable variables it can't be "causal". No causes in - no causes out ( N . Cartwright) \#Bookofwhy
Q 3
$\uparrow\urcorner$
023
↔

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## Independent Component Analysis (ICA)

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- Factor analysis, data compression, etc.
- Invariant to rotation.


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- Factor analysis, data compression, etc.
- Invariant to rotation.
-ICA: $X=Z W$ usually with $k=d$
- Require the components of each $z_{i}$ to be independent, and at most one can be normally distributed.
- Independence is measured by non-normality.
- $W=\operatorname{argmax}_{\mathrm{W}} \sum_{i=1}^{n} \sum_{j=1}^{k} \operatorname{kurt}\left(w_{j}^{T} x_{i}\right)^{2}$ where $\operatorname{kurt}(u)=\mathbb{E}\left(u^{4}\right)-3\left(\mathbb{E}\left(u^{2}\right)\right)^{2}$
- Up to permutation and scaling, we can identify the factors W .


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- Structural equation modeling (SEM): path analysis integrating factor analysis and latent variables for non-experimental data.
- But requires background knowledge BEFORE collecting and analyzing data.
- Normality constraints.
- Causal directions are often unknown.
- Model 1: $x_{1}=b_{12} x_{2}+\epsilon_{1}$ and Model 2: $x_{2}=b_{21} x_{1}+\epsilon_{1}$, both are saturated with the same covariance matrix.


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- Model 1: $x_{1}=b_{12} x_{2}+\epsilon_{1}$ and Model 2: $x_{2}=b_{21} x_{1}+\epsilon_{1}$, both are saturated with the same covariance matrix.
- With non-normality, ICA can be used to find the causal ordering of any number of observed variables based on non-experimental data. (Shimizu et al. 2006)


## Causal Order

- Causality (a causal order) from a random variable $x_{1}$ to another random variable $x_{2}$, denoted by $x_{1} \rightarrow x_{2}$ is confirmed if the following holds:
- $x_{2}=f\left(x_{1}, \epsilon_{2}\right)$ where $\epsilon_{2}$ is some perturbation independently distributed from $x_{1}$.
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- Note that we can't just have uncorrelatedness from $x_{1}$ and $\epsilon_{2}$ :
- Independence: $\mathbb{E}(g(x) h(y))=\mathbb{E}(g(x)) \mathbb{E}(h(y))$ for any deterministic functions $g$ and $h$.
- Uncorrelatedness (covariance is 0 ): $\mathbb{E}(x y)=\mathbb{E}(x) \mathbb{E}(y)$, weaker assumption.


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- Uncorrelatedness (covariance is 0 ): $\mathbb{E}(x y)=\mathbb{E}(x) \mathbb{E}(y)$, weaker assumption.
- If $x_{1}$ and $\epsilon_{2}$ are only uncorrelated, then there could exist unobserved confounder $z$ that affects both $x_{1}$ and $x_{2}$.
- Suppose $x_{1}=\alpha z+\epsilon_{1}$ and $x_{2}=\beta z+b_{21} x_{1}+\epsilon_{2}$, then $\operatorname{Cov}\left(x_{1}, x_{2}\right)=b_{21} \operatorname{Var}\left(x_{1}\right)+\alpha \beta \operatorname{Var}(z)$ can be non-zero even if $b_{21}$ is 0 .


## Finding a Causal Order

- Suppose we have $N$ measurements of $x_{1}$ and $x_{2}$.
- $\overline{x_{1}^{2}}=\frac{1}{N} \sum_{i=1}^{N} x_{1 i}^{2}$, similarly for $\overline{x_{2}^{2}}$ and $\overline{x_{1} x_{2}}$
- Model 1: $x_{1}=b_{12} x_{2}+\epsilon_{1}$
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- Model 1: $x_{1}=b_{12} x_{2}+\epsilon_{1}$
- Model 2: $x_{2}=b_{21} x_{1}+\epsilon_{2}$
- Second order moments of Model 1:
$E\left[\begin{array}{c}\overline{x_{1}^{2}} \\ \overline{x_{1} x_{2}} \\ \overline{x_{2}^{2}}\end{array}\right]=\left[\begin{array}{c}b_{12}^{2} E\left(x_{2}^{2}\right)+E\left(\xi_{1}^{2}\right) \\ b_{12} E\left(x_{2}^{2}\right) \\ E\left(x_{2}^{2}\right)\end{array}\right] \begin{aligned} & \text { which we denote by } \\ & \left.\boldsymbol{\tau}_{2}=\left[E\left(x_{2}^{2}\right), E\left(\xi_{1}^{2}\right), b_{12}\right]^{T}\right]=\boldsymbol{\sigma}_{2}\left(\boldsymbol{\tau}_{2}\right)\end{aligned}$
- Symmetric for Model 2.
- Undistinguishable.


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- Fourth order moments for Model 1 and assume the residuals are NOT normally distributed:

which we denote by $E\left[\boldsymbol{m}_{4}\right]=\boldsymbol{\sigma}_{4}\left(\boldsymbol{\tau}_{4}\right)$,

$$
\boldsymbol{\tau}_{4}=\left[\boldsymbol{\tau}_{2}^{T}, E\left(x_{2}^{4}\right), E\left(\xi_{1}^{4}\right)\right]^{T} .
$$

$E\left[\begin{array}{c}\overline{x_{1}^{2}} \\ \overline{x_{1} x_{2}} \\ \overline{x_{2}^{2}}\end{array}\right]=\left[\begin{array}{c}b_{12}^{2} E\left(x_{2}^{2}\right)+E\left(\xi_{1}^{2}\right) \\ b_{12} E\left(x_{2}^{2}\right) \\ E\left(x_{2}^{2}\right)\end{array}\right] \begin{aligned} & \text { which we denote by } E\left[\boldsymbol{m}_{2}\right]=\boldsymbol{\sigma}_{2}\left(\boldsymbol{\tau}_{2}\right) \\ & \boldsymbol{\tau}_{2}=\left[E\left(x_{2}^{2}\right), E\left(\xi_{1}^{2}\right), b_{12}\right]^{T}\end{aligned}$

- Symmetric for Model 2.
- Undistinguishable.


## Finding a Causal Order

- Suppose we have $N$ measurements of $x_{1}$ and $x_{2}$.
- $\overline{x_{1}^{2}}=\frac{1}{N} \sum_{i=1}^{N} x_{1 i}^{2}$, similarly for $\overline{x_{2}^{2}}$ and $\overline{x_{1} x_{2}}$
- Model 1: $x_{1}=b_{12} x_{2}+\epsilon_{1}$
- Model 2: $x_{2}=b_{21} x_{1}+\epsilon_{2}$
- Second order moments of Model 1:

$$
E\left[\begin{array}{c}
\overline{x_{1}^{2}} \\
\overline{x_{1} x_{2}} \\
\overline{x_{2}^{2}}
\end{array}\right]=\left[\begin{array}{c}
b_{12}^{2} E\left(x_{2}^{2}\right)+E\left(\xi_{1}^{2}\right) \\
b_{12} E\left(x_{2}^{2}\right) \\
E\left(x_{2}^{2}\right)
\end{array}\right] \begin{aligned}
& \\
& \text { which we denote by } \\
& \boldsymbol{\tau}_{2}=\left[E\left(x_{2}^{2}\right), E\left(\xi_{1}^{2}\right), b_{12}\right]^{T}
\end{aligned}
$$

- Symmetric for Model 2.
- Undistinguishable.
- Fourth order moments for Model 1 and assume the residuals are NOT normally distributed:

$$
E\left[\begin{array}{c}
\overline{x_{1}^{4}} \\
\overline{x_{1}^{3} x_{2}} \\
\overline{x_{1}^{2} x_{2}^{2}} \\
\overline{x_{1} x_{2}^{3}} \\
\overline{x_{2}^{4}}
\end{array}\right]=\left[\begin{array}{c}
b_{12}^{4} E\left(x_{2}^{4}\right)+6 b_{12}^{2} E\left(x_{2}^{2}\right) E\left(\xi_{1}^{2}\right)+E\left(\xi_{1}^{4}\right) \\
b_{12}^{3} E\left(x_{2}^{4}\right)+3 b_{12} E\left(x_{2}^{2}\right) E\left(\xi_{1}^{2}\right) \\
b_{12}^{2} E\left(x_{2}^{4}\right)+E\left(x_{2}^{2}\right) E\left(\xi_{1}^{2}\right) \\
b_{12} E\left(x_{2}^{4}\right) \\
E\left(x_{2}^{4}\right)
\end{array}\right]
$$

$$
\text { which we denote by } E\left[\boldsymbol{m}_{4}\right]=\boldsymbol{\sigma}_{4}\left(\boldsymbol{\tau}_{4}\right) \text {, }
$$

$$
\boldsymbol{\tau}_{4}=\left[\boldsymbol{\tau}_{2}^{T}, E\left(x_{2}^{4}\right), E\left(\xi_{1}^{4}\right)\right]^{T}
$$

$$
T=N\left(\left[\begin{array}{l}
\boldsymbol{m}_{2} \\
\boldsymbol{m}_{4}
\end{array}\right]-\left[\begin{array}{l}
\boldsymbol{\sigma}_{2}\left(\hat{\boldsymbol{\tau}}_{2}\right) \\
\boldsymbol{\sigma}_{4}\left(\widehat{\boldsymbol{\tau}}_{4}\right)
\end{array}\right]\right)^{T} \hat{M}\left(\left[\begin{array}{l}
\boldsymbol{m}_{2} \\
\boldsymbol{m}_{4}
\end{array}\right]-\left[\begin{array}{l}
\boldsymbol{\sigma}_{2}\left(\hat{\boldsymbol{\tau}}_{2}\right) \\
\boldsymbol{\sigma}_{4}\left(\hat{\boldsymbol{\tau}}_{4}\right)
\end{array}\right]\right),
$$

- Measure of model fit $\approx$ distance between data and the model used
- Compare T for the two models which will imply causal direction.


## Causal ordering - more than 2 variables

- Suppose we have $n$ variables, and we want to find an ordering $i(1), \ldots, i(n)$ such that
- $x_{i(j)}=\sum_{k=1}^{j-1} b_{i(j), i(k)} x_{i(k)}+\epsilon_{i(j)}$ for all $j=1, \ldots, n$, with nonzero coefficients and $\epsilon_{i(j)}$ non-normal, independent from $x_{i(k)}$ for $k<j$.
- " $x_{i(j)}$ can be written as a linear combination of its preceding variables in that order plus an (mutually) independent error."
- Causal ordering $x_{i(1)} \rightarrow x_{i(2)} \rightarrow \cdots \rightarrow x_{i(n)}$


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- Assuming such model exists, we need find the correct mapping $i(j)$ for $j=1, \ldots, n$.
- $\tilde{x}=B \tilde{x}+\tilde{\epsilon}$ where $B$ is lower triangular and $\tilde{x}$ is a vector of the observed variables with the desired ordering.


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- $\tilde{x}=B \tilde{x}+\tilde{\epsilon}$ where $B$ is lower triangular and $\tilde{x}$ is a vector of the observed variables with the desired ordering.
- Normalize the above so that the errors have unit variance using
- $w_{i(j), i(j)}=1 / \sqrt{ } \operatorname{Var}\left(\epsilon_{i(j)}\right)$ and $w_{i(j), i(k)}=-b_{i(j), i(k)} / \sqrt{ } \operatorname{Var}\left(\epsilon_{i(j)}\right)$ for $k \neq j$
- $\operatorname{diag}(W) \tilde{x}=-\operatorname{offdiag}(W) \tilde{x}+\tilde{\epsilon}^{*}$, or $W \tilde{x}=\tilde{\epsilon}^{*}$


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- We can use ICA to estimate $W$ !
- And there exists a unique permutation to make $W$ lower triangular if the coefficients in $B$ are non-zero.
- We can also estimate B, depending whether we care about the coefficients or just the causal ordering.


## Summary (and References)

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Thank you

