

Graphical Models

Undirected Models

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Winter 2018

Learning objectives

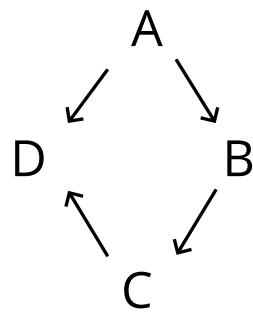
Markov networks:

- independence assumptions
 - factorization
 - representations:
 - factor-graph
 - log-linear models
- 
- Hammersley-Clifford theorem

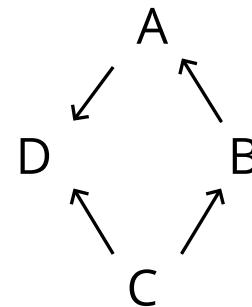
Challenge

Given the following set of CIs draw their DAG

$$\mathcal{I}(P) = \{(A \perp C \mid B, D), (D \perp B \mid A, C)\}$$



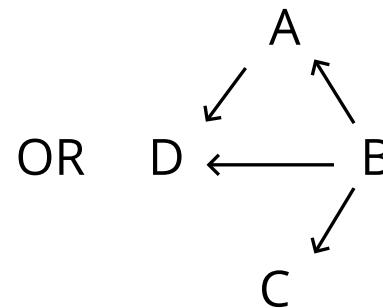
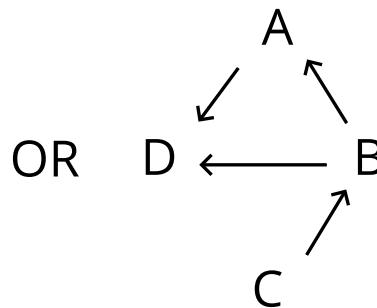
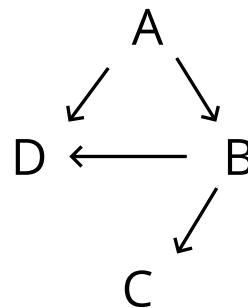
OR



Challenge

Given the following set of CIs draw their DAG

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Challenge

Given the following set of CIs draw their DAG

$$\mathcal{I}(P) = \{(A \perp C \mid B, D), (D \perp B \mid A, C)\}$$

a DAG cannot be a P-map for P

an undirected model can!

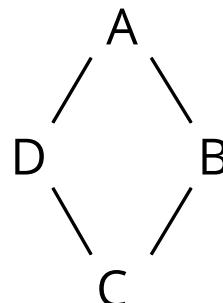
Challenge

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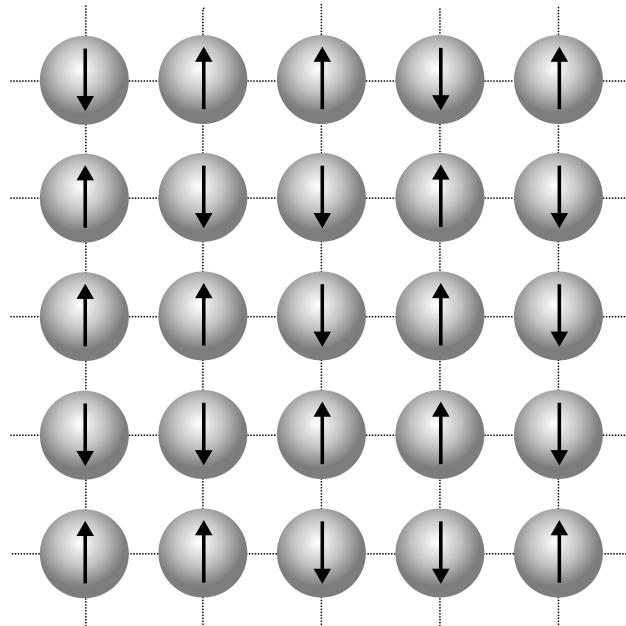
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an undirected model can!



Motivation

Statistical physics: **Ising model** of ferromagnetism

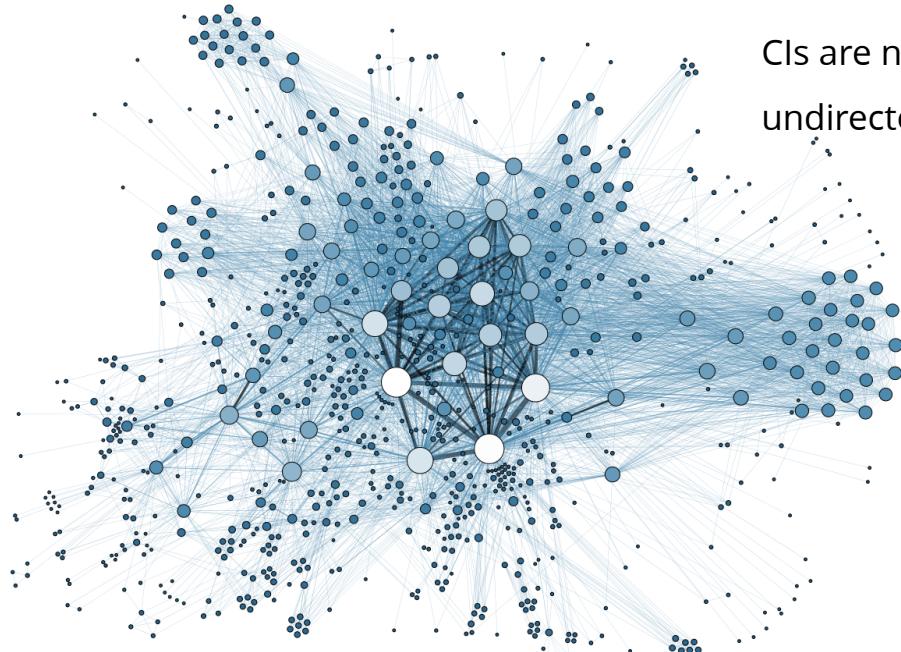


CLs are naturally expressed using an undirected model

Image: <https://web.stanford.edu/~peastman/statmech/phasetransitions.html>

Motivation

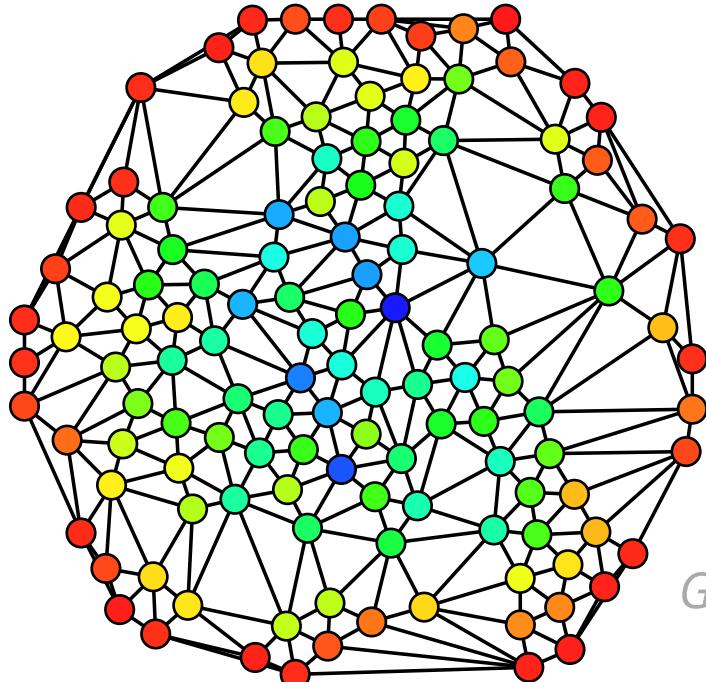
Social sciences



CIs are naturally expressed using an undirected model

Motivation

Combinatorial problems



CI_s are naturally expressed using an undirected model

Graph coloring

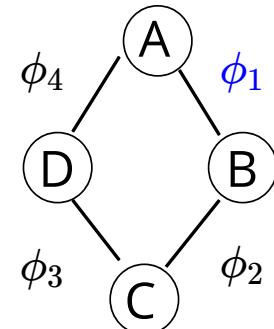
Factorization in Markov networks

$$P(A, B, C, D) = \frac{1}{Z} \phi_1(A, B) \phi_2(B, C) \phi_3(C, D) \phi_4(A, D)$$

$$Z = \sum_{a,b,c,d} \phi_1(a, b) \phi_2(b, c) \phi_3(c, d) \phi_4(a, d)$$

is a normalization constant (*partition function*)

$\phi_1 : Val(A, B) \rightarrow [0, +\infty)$ is called a factor (*potential*)



MRF; Conditional Independencies

$$P(A, B, C, D) = \left(\frac{1}{Z} \phi_1(A, B) \phi_2(B, C) \right) f(B, A, C) \phi_3(C, D) \phi_4(A, D)$$

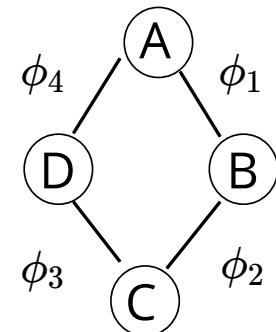


$$P \models (B \perp D \mid A, C)$$

$$P(A, B, C, D) = \left(\frac{1}{Z} \phi_1(A, B) \phi_2(A, D) \right) \phi_3(C, D) \phi_4(B, C)$$



$$P \models (A \perp C \mid B, D)$$



Product of factors

$$P(A, B, C, D) = \frac{1}{Z} \phi_1(A, B) \phi_2(B, C) \phi_3(C, D) \phi_4(A, D)$$

$$\psi(A, B, C) : Val(A, B, C) \rightarrow \Re^+$$

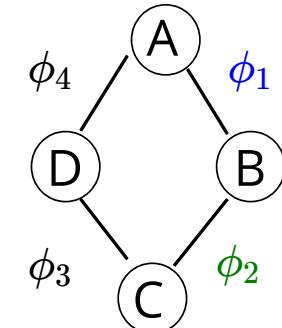
$$\phi_1 : Val(A, B) \rightarrow \Re^+$$

a^1	b^1	0.5
a^1	b^2	0.8
a^2	b^1	0.1
a^2	b^2	0
a^3	b^1	0.3
a^3	b^2	0.9

$$\phi_2 : Val(B, C) \rightarrow \Re^+$$

b^1	c^1	0.5
b^1	c^2	0.7
b^2	c^1	0.1
b^2	c^2	0.2

a^1	b^1	c^1	$0.5 \cdot 0.5 = 0.25$
a^1	b^1	c^2	$0.5 \cdot 0.7 = 0.35$
a^1	b^2	c^1	$0.8 \cdot 0.1 = 0.08$
a^1	b^2	c^2	$0.8 \cdot 0.2 = 0.16$
a^2	b^1	c^1	$0.1 \cdot 0.5 = 0.05$
a^2	b^1	c^2	$0.1 \cdot 0.7 = 0.07$
a^2	b^2	c^1	$0 \cdot 0.1 = 0$
a^2	b^2	c^2	$0 \cdot 0.2 = 0$
a^3	b^1	c^1	$0.3 \cdot 0.5 = 0.15$
a^3	b^1	c^2	$0.3 \cdot 0.7 = 0.21$
a^3	b^2	c^1	$0.9 \cdot 0.1 = 0.09$
a^3	b^2	c^2	$0.9 \cdot 0.2 = 0.18$



$Val(A) \times Val(B) \times Val(C)$ similar to a 3D tensor

Q: Do factors represent marginals?

Simplified example: $P(A, B, C) = \frac{1}{Z} \phi_1(A, B) \phi_2(B, C)$

ϕ_1			ϕ_2		
a^1	b^1	0.5	b^1	c^1	0.5
a^1	b^2	0.8	b^1	c^2	0.7
a^2	b^1	0.1	b^2	c^1	0.1
a^2	b^2	0	b^2	c^2	0.2
a^3	b^1	0.3			
a^3	b^2	0.9			



a^1	b^1	c^1	$0.5 \cdot 0.5 = 0.25$
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$$Z = .25 + .35 + \dots = 1.55$$

Marginal probabilities:

$$P(a^1, b^1) = (.25 + .35)/Z \approx .38$$

$$P(a^1, b^2) = (.08 + .16)/Z \approx .15$$

Compare to ϕ_1

$$\phi_1(a^1, b^1) = .5$$

$$\phi_1(a^1, b^2) = .8$$

Factorization: general form

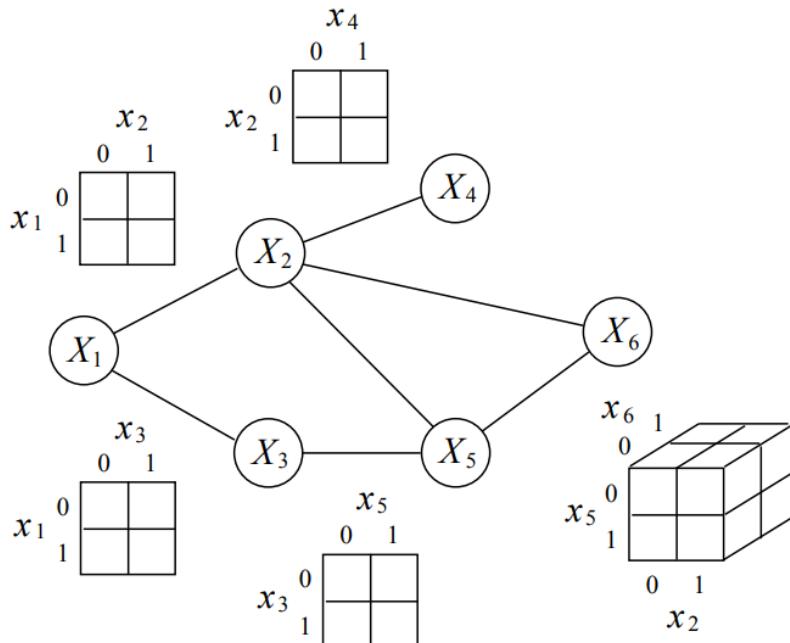
\mathbf{P} factorizes over the *cliques*



$$P(\mathbf{X}) = \frac{1}{Z} \prod_k \phi_k(\mathbf{D}_k)$$

Gibbs distribution

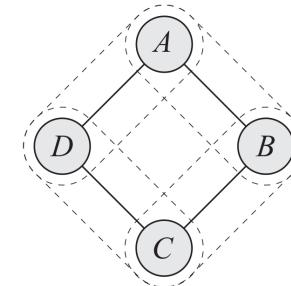
Can always convert to
factorization over *maximal cliques*



Factorization: general form

\mathbf{P} factorizes over *cliques*

$$P(\mathbf{X}) = \frac{1}{Z} \prod_k \phi_k(\mathbf{D}_k)$$



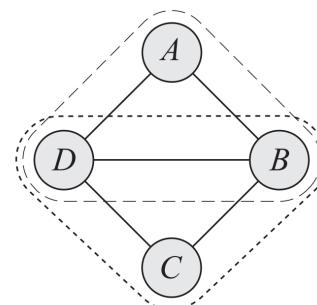
Rewrite as factorization over **maximal cliques**

- original form of \mathbf{P}

$$P(A, B, C, D) = \phi_1(A, B)\phi_2(A, D)\phi_3(B, D)\phi_4(C, D)\phi_5(B, C)$$

- factorized over cliques

$$P(A, B, C, D) = \psi_1(A, B, C)\psi_2(B, C, D)$$



Factorized form: directed vs undirected

Markov Networks:

$$P(\mathbf{X}) = \frac{1}{Z} \prod_k \phi_k(\mathbf{D}_k)$$

Bayesian Networks:

$$P(\mathbf{X}) = \prod_k P(X_i \mid Pa_{X_i})$$

- No *partition function*
- Each factor is a *cond. distribution*
- One factor per variable

Conditioning on the evidence

given $P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{D}_k)$, how to obtain $P(\mathbf{X} | U = u)$?

fix the evidence in the relevant factors $P(\mathbf{X} | U = u) \propto \prod_k \phi_k[U = u]$

$\phi_k(A, B, C)$

a^1	b^1	c^1	$0.5 \cdot 0.5 = 0.25$
a^1	b^1	c^2	$0.5 \cdot 0.7 = 0.35$
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conditioned on $C = c^1$



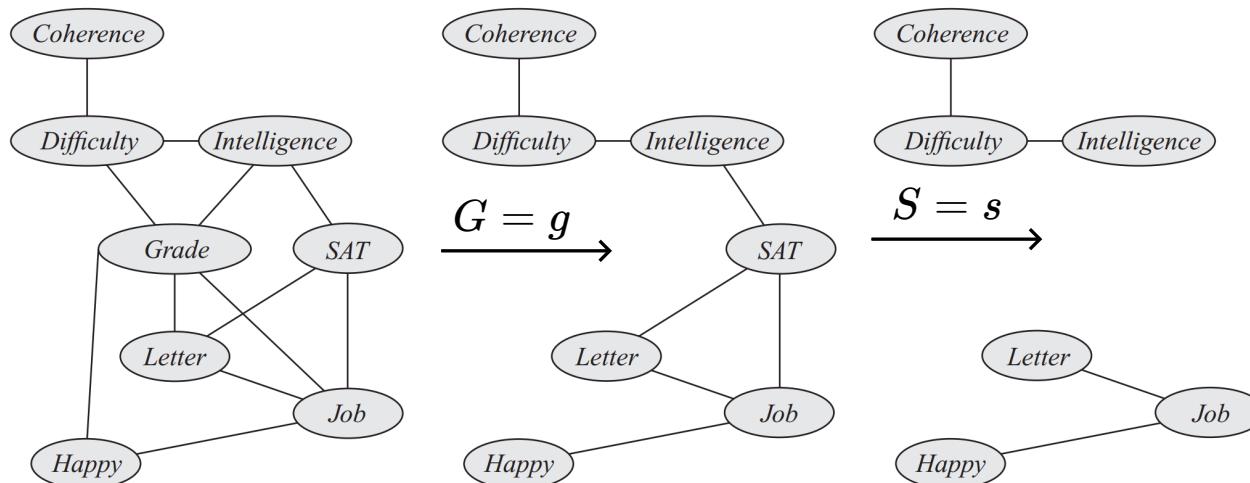
reduced factor

a^1	b^1	c^1	0.25
a^1	b^2	c^1	0.08
a^2	b^1	c^1	0.05
a^2	b^2	c^1	0
a^3	b^1	c^1	0.15
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Conditioning on the evidence

effect on the graphical model

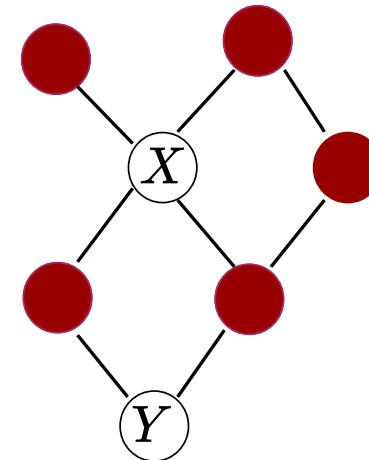
- cannot create new dependencies
- compare this to colliders in **Bayes-nets**



Pairwise conditional independencies

Non-adjacent nodes are independent given everything else

$$X \perp Y \mid \mathcal{X} - \{X, Y\}$$

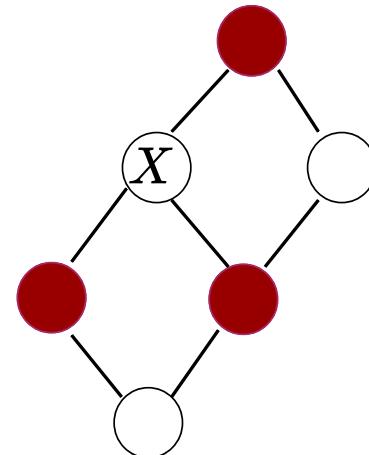


Local conditional Independencies

$MB^{\mathcal{H}}(X)$: **Markov blanket** of node X in graph H

$$X \perp \mathcal{X} - X - MB^{\mathcal{H}}(X) \mid MB^{\mathcal{H}}$$

Given its Markov blanket X is independent of every other variable



Local conditional Independencies

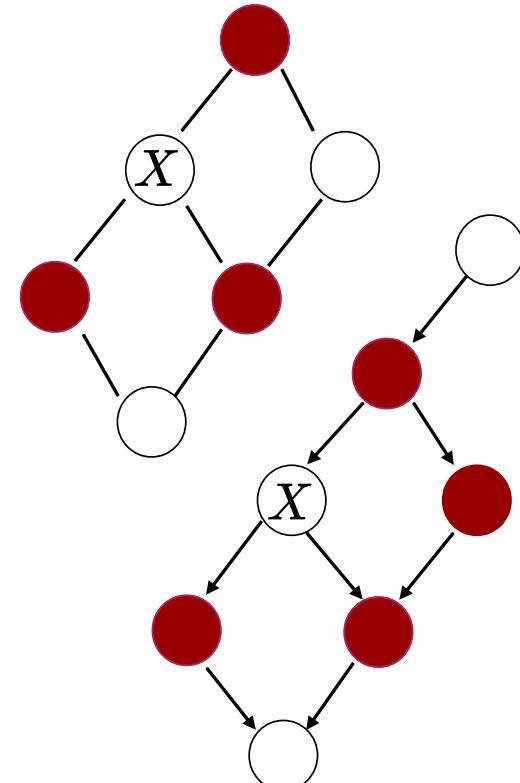
$MB^{\mathcal{H}}(X)$: **Markov blanket** of X in graph H

$$X \perp \mathcal{X} - X - MB^{\mathcal{H}}(X) \mid MB^{\mathcal{H}}$$

$MB^{\mathcal{G}}(X)$: **Markov blanket** of X in DAG G

- Parents
- *Children*
- *Parents of children*

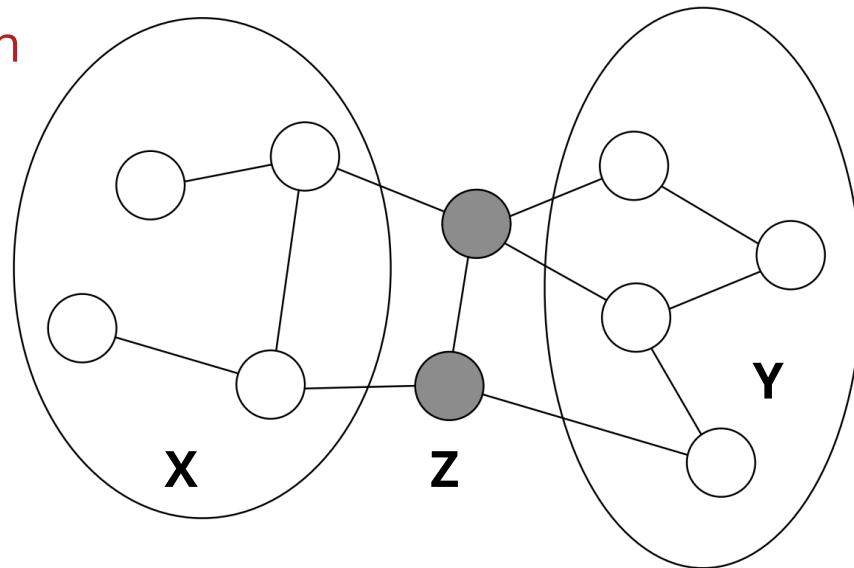
$$X \perp \mathcal{X} - X - MB^{\mathcal{G}}(X) \mid MB^{\mathcal{G}}$$



Global conditional Independencies

$X \perp Y \mid Z$ iff every path between **X** and **Y** is blocked by **Z**

much simpler than D-separation



Relationship between the three

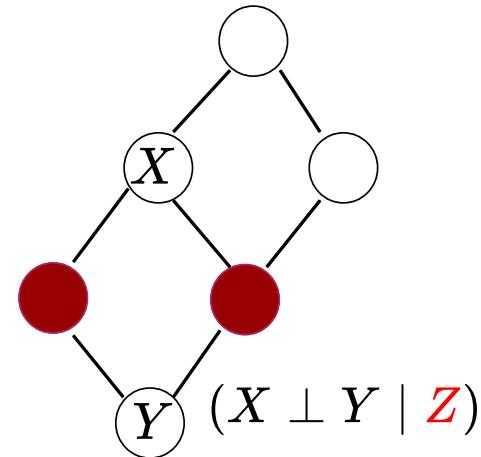
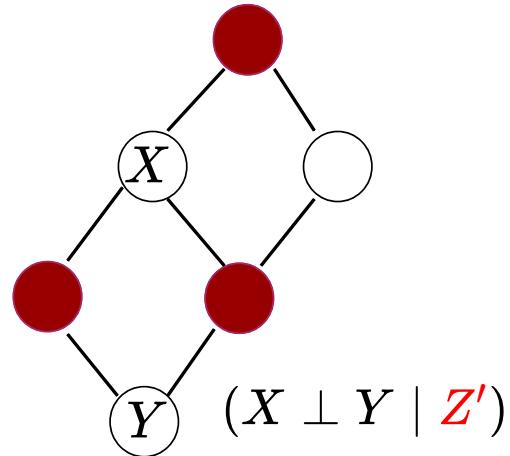
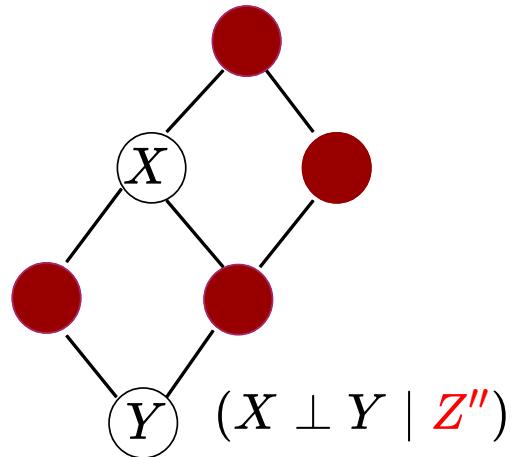
pairwise \mathcal{I}_p

\Leftarrow

local \mathcal{I}_ℓ

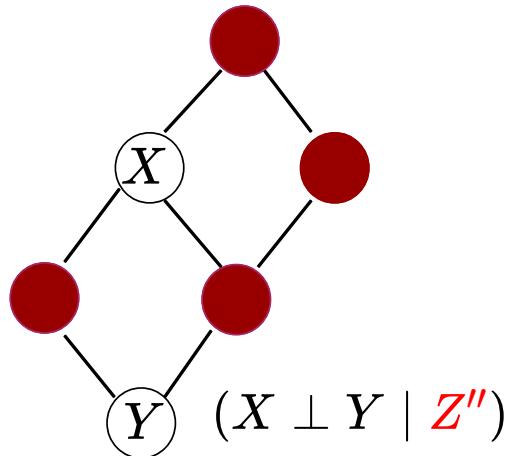
\Leftarrow

global \mathcal{I}



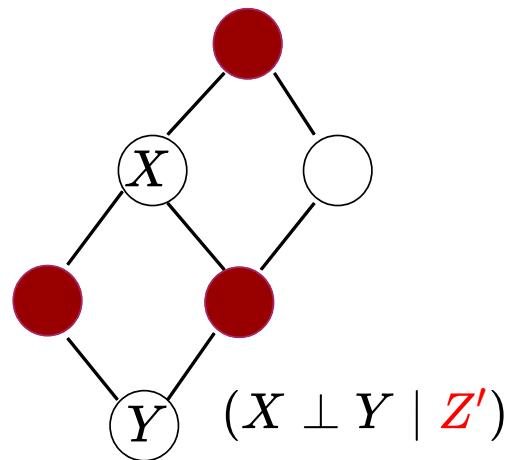
Relationship between the three

pairwise \mathcal{I}_p



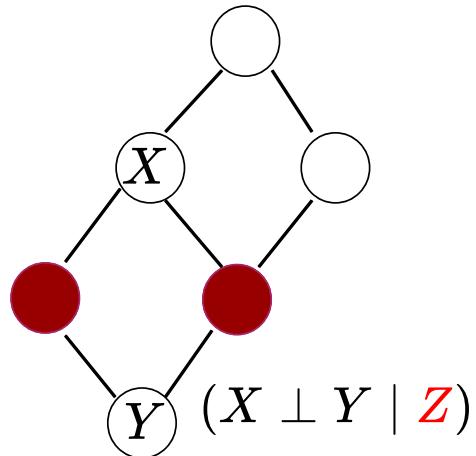
\Leftarrow

local \mathcal{I}_ℓ



\Leftarrow

global \mathcal{I}



P>0: pairwise $\mathcal{I}_p \Rightarrow$

local $\mathcal{I}_\ell \Rightarrow$

global \mathcal{I}

Factorization & independence

Recall this relationship in **Bayesian Networks**:

- **Factorization** according to a DAG
- **Local & global** CIs

Equivalent

(same family of distributions)

Is it similar for **Markov Networks**?

- **Factorization** according to an *undirected graph*
- **Pairwise, local & global** CIs

Equivalent?

Factorization & Independence

Is it similar for **Markov Networks**?

- **Factorization** according to an *undirected graph*
- **Pairwise, local & global** CIs

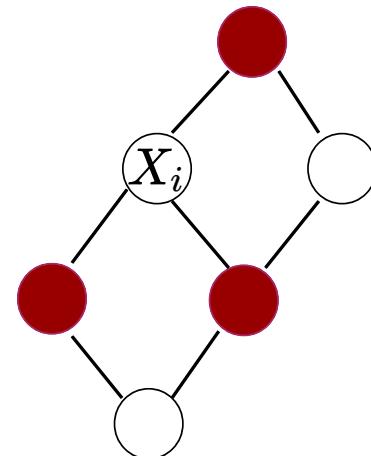


Short answer:

- for positive distributions they are equivalent

Factorization \Rightarrow CI

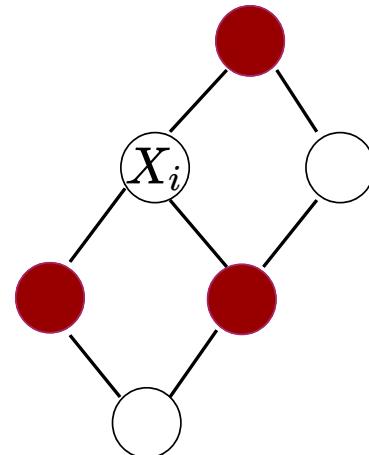
given $P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{C}_k)$ does **local** CI hold?



Factorization \Rightarrow CI

given $P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{C}_k)$ does **local** CI hold?

proof



Factorization \Rightarrow CI

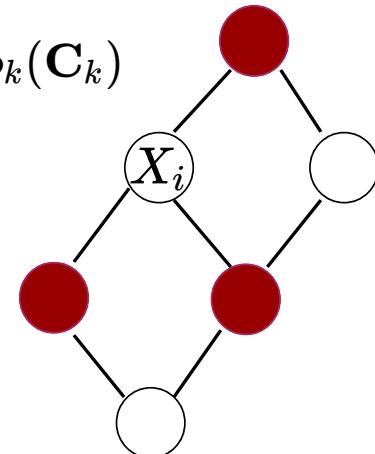
given $P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{C}_k)$ does **local** CI hold?

proof

$$P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{C}_k) = \prod_{\mathbf{C}_k \in MB(X_i)} \phi_k(\mathbf{C}_k) \prod_{\mathbf{C}_k \notin MB(X_i)} \phi_k(\mathbf{C}_k)$$

$$= f(X_i, MB(X_i)) g(\mathcal{X} - X_i) \quad \Rightarrow$$

$$X_i \perp \mathcal{X} - MB^{\mathcal{H}}(X_i) - X_i \mid MB^{\mathcal{H}}(X_i)$$



CI \Rightarrow factorization

Hammersely-Clifford theorem:

If P is *strictly positive* satisfying CI $\mathcal{I}(\mathcal{H})$
then P factorizes over \mathcal{H}

proof

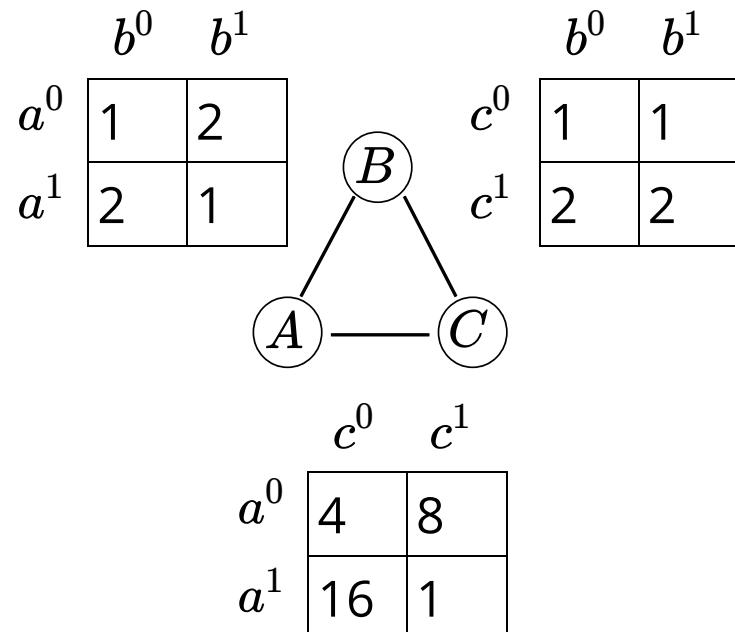
needs canonical parametrization



Parametrization: redundancy

is this representation of P unique?

$$P(A, B, C) \propto \phi_1(A, B)\phi_2(B, C)\phi_3(C, A)$$



Parametrization: redundancy

is this representation of P unique?

$$P(A, B, C) = \frac{1}{Z} \phi_1(A, B) \phi_2(B, C) \phi_3(C, A)$$

		b^0	b^1			b^0	b^1
a^0	10	20		c^0	10	10	
a^1	20	10		c^1	20	20	

Graph structure:

```
graph TD; A((A)) --- B((B)); A --- C((C));
```

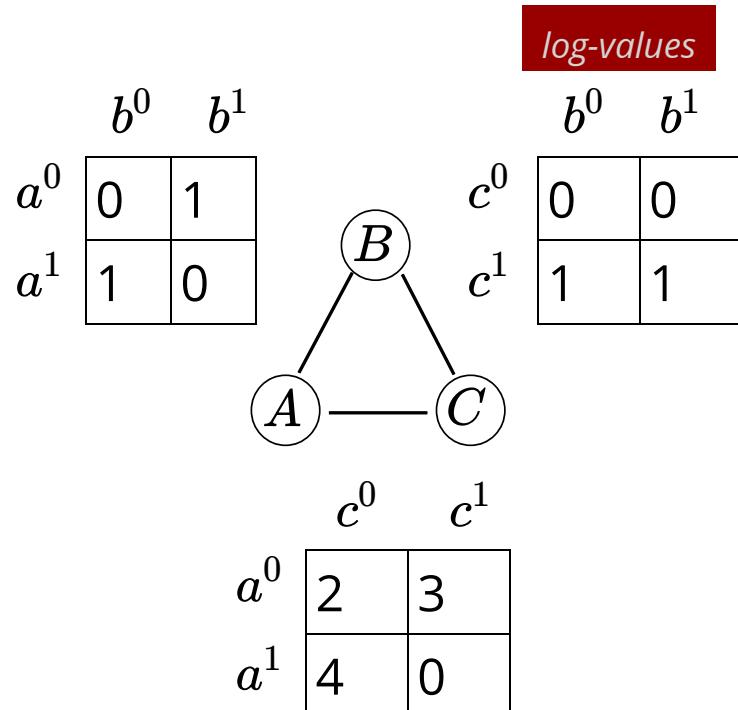
multiplying all factors by a constant
only affects Z

	c^0	c^1
a^0	.4	.8
a^1	1.6	.1

Parametrization: redundancy

is this representation of P unique?

$$P(A, B, C) = \frac{1}{Z} \phi_1(A, B) \phi_2(B, C) \phi_3(C, A)$$



use the logarithmic form

$$P(A, B, C) = \frac{1}{Z} 2^{(\psi_1(A, B) + \psi_2(B, C) + \psi_3(C, A))}$$

Parametrization: redundancy

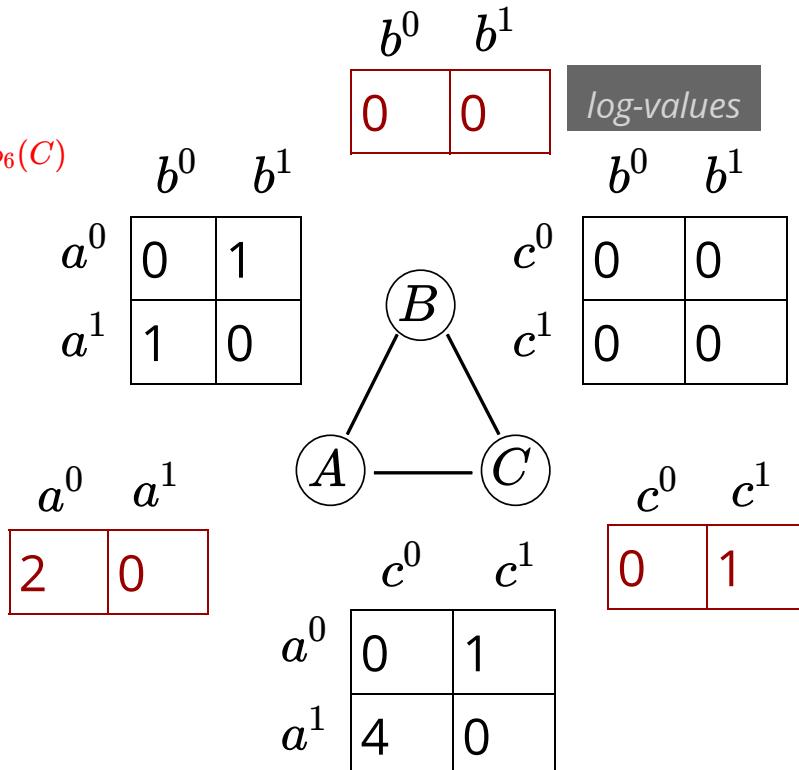
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$$P(A, B, C) = \frac{1}{Z} \phi_1(A, B) \phi_2(B, C) \phi_3(C, A) \phi_4(B) \phi_5(A) \phi_6(C)$$

use the logarithmic form

$$P(A, B, C) = \frac{1}{Z} 2^{(\psi_1(A, B) + \psi_2(B, C) + \psi_3(C, A))}$$

simplify using local potentials



Parametrization: redundancy

is this representation of P unique?

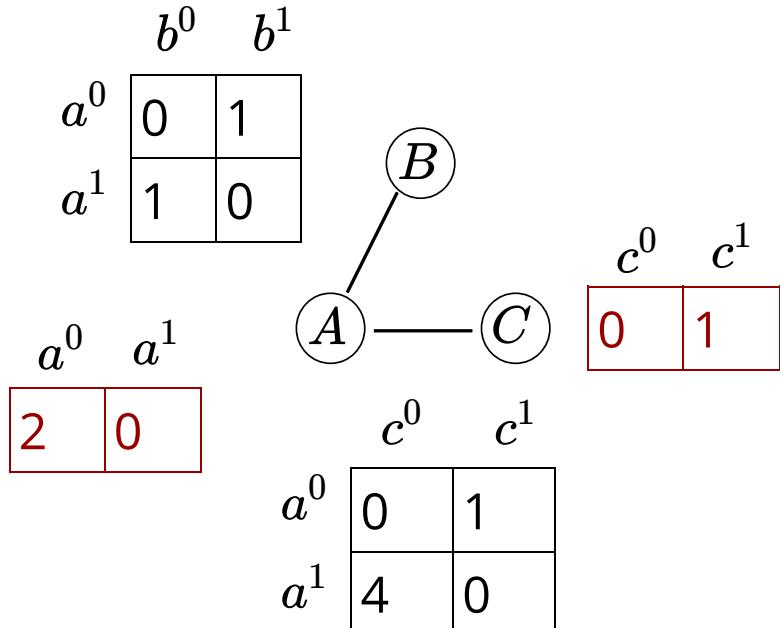
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use the logarithmic form

$$P(A, B, C) = \frac{1}{Z} 2^{(\psi_1(A, B) + \psi_2(B, C) + \psi_3(C, A))}$$

simplify using local potentials

log-values



Parametrization: example (Ising model)

Ising model: $Val(X_i) = \{-1, +1\}$ $p(\mathbf{x}) = \frac{1}{Z(t)} \exp \left(-\frac{1}{t} \left(\sum_i h_i x_i + \frac{1}{2} \sum_{i,j \in \mathcal{E}} x_i J_{ij} x_j \right) \right)$

can represent all positive, pairwise Markov networks over the binary domain

interactions		local field	
-1	J	-1	$+1$
$-J$	J	$-h$	h

log-values

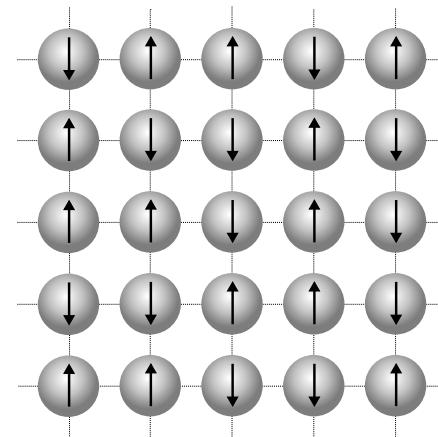


Image:
<https://web.stanford.edu/~peastman/statmech/phasetransitions.html>

Parametrization: example (Boltzmann machine)

Boltzmann machine: $Val(X_i) = \{0, 1\}$

$$p(\mathbf{x}) = \frac{1}{Z} \exp \left(- \sum_i b_i x_i - \frac{1}{2} \sum_{i,j \in \mathcal{E}} x_i W_{ij} x_j \right)$$

		interaction weights		local bias	
		-1	+1	-1	+1
-1	-1	0	0	0	h
	+1	0	J		
		log-values			

Parametrization: log-linear model

for a positive distribution:

$$P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{D}_k) = \exp\left(-\sum_k \psi_k(\mathbf{D}_k)\right)$$

energy $\frac{-\log(\phi_k(\mathbf{D}_k))}{\text{energy}}$

Parametrization: log-linear model

for a positive distribution:

$$P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{D}_k) = \exp\left(-\sum_k \psi_k(\mathbf{D}_k)\right)$$

energy $- \log(\phi_k(\mathbf{D}_k))$

linearly parameterize it:

$$P_w(\mathbf{X}) \propto \exp\left(-\sum_k w_k \underline{f_k(\mathbf{D}_k)}\right)$$

feature/sufficient statistics

Parametrization: log-linear model

for a positive distribution:

$$P(\mathbf{X}) \propto \prod_k \phi_k(\mathbf{D}_k) = \exp\left(-\sum_k \psi_k(\mathbf{D}_k)\right)$$

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linearly parameterize it:

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feature/sufficient statistics

revisit in the *exponential family*

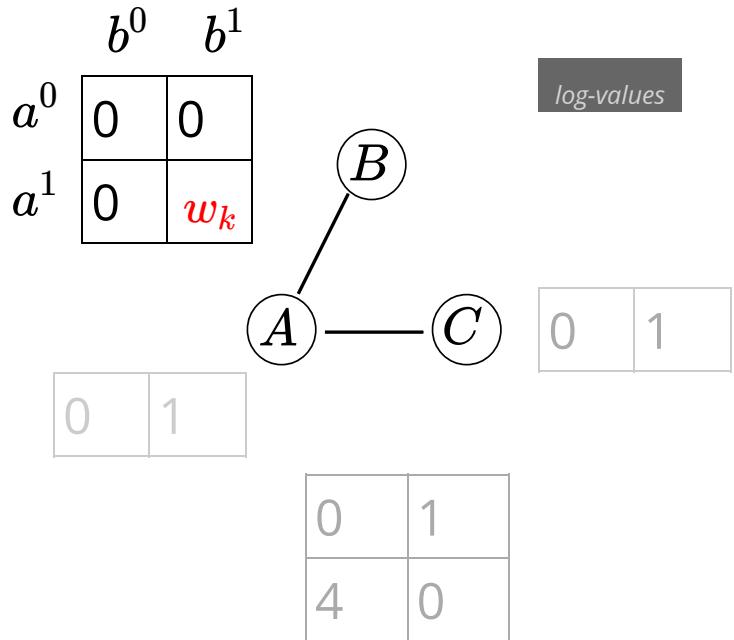
Parametrization: log-linear model

features in **discrete** distributions:

$$P_w(\mathbf{X}) \propto \exp(-\sum_k w_k f_k(\mathbf{D}_k))$$



$$f_{1,1}(A, B) = \mathbb{I}(A = a^1, B = b^1)$$



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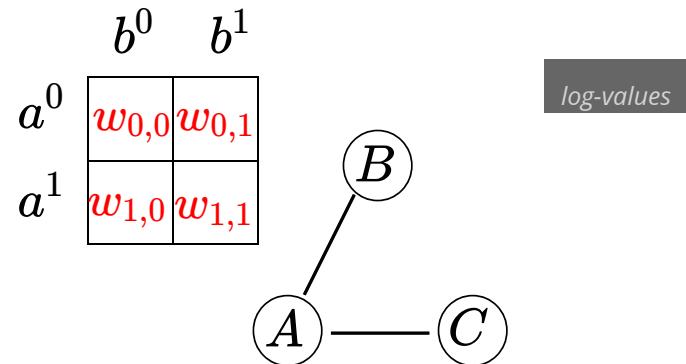


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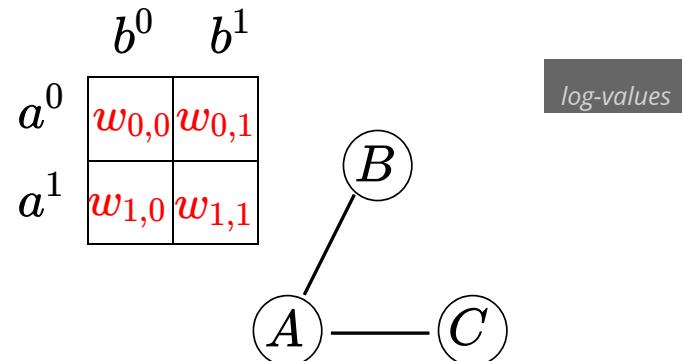


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Overparameterized model: $\{\mathbf{w}_k\} \rightarrow P_w$ is not one-to-one

Parametrization: log-linear model

$$P_w(\mathbf{X}) \propto \exp(-\sum_k w_k f_k(\mathbf{D}_k))$$

Redundant \equiv linearly dependent features



$$\sum_k \alpha_k f_k(\mathbf{D}) = \alpha \quad \forall \mathbf{D}$$

$$P_w(\mathbf{X}) \propto \exp(-\sum_k w_k f_k(\mathbf{D}_k)) \propto \exp(-\sum_k (w_k + \alpha_k) f_k(\mathbf{D}_k)) \propto P_{w+\alpha}(\mathbf{X})$$

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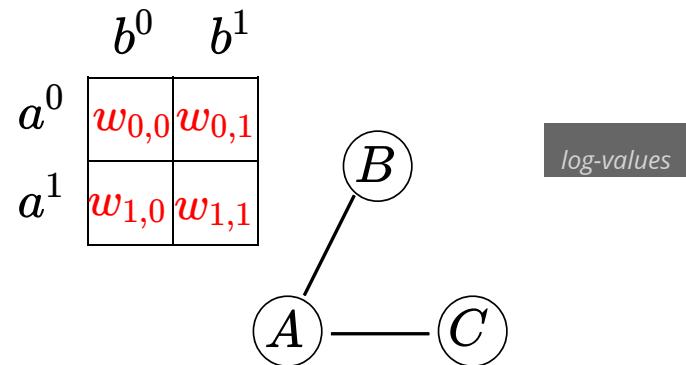


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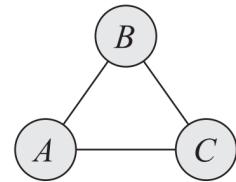
Linear dependency of features:

$$f_{0,0}(A, B) + f_{1,0}(A, B) + f_{0,1}(A, B) + f_{1,1}(A, B) = 1$$

Parametrization: factor-graph

Markov network representation:

- ✓ • identifies CI
- defines the factorized form
- ⌚ ■ is not fine-grained enough



$$P(A, B, C) = \phi_1(A, B)\phi_2(B, C)\phi_3(C, A) \quad ?$$

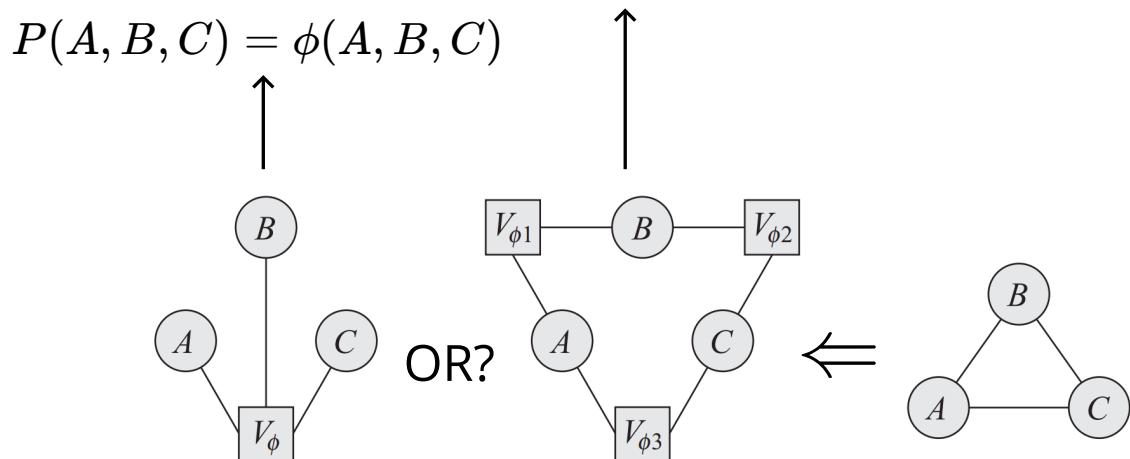
$$P(A, B, C) = \phi_1(A, B, C) \quad ?$$

Parametrization: factor-graph

use a bipartite structure:

- factors (square)
- variables (circle)

$$P(A, B, C) = \phi_1(A, B)\phi_2(B, C)\phi_3(C, A)$$



Summary

- similar to directed models:
 - factorization of the probability over cliques
 - set of conditional independencies
 - (pariwise, local, global)
- $P > 0 \Rightarrow$
same family of dists.

Summary

- similar to directed models:
 - factorization of the probability over cliques
 - set of conditional independencies
 - (pariwise, local, global)
 - parametrization
 - redundancy (same dist. different params/factors)
 - log-linear model
 - factor-graph (finer-grained specification of the factors)
- $P > 0 \Rightarrow$

same family of dists.

Bonus Slides

Parametrization: canonical form

reparameterize a given Gibbs dist.

$$P(\mathbf{X}) \propto \exp(-\sum_k \psi_k(\mathbf{D}_k))$$

such that low order interactions are automatically moved to smaller cliques

need to fixed an assignment $\xi^* = (x_1^*, \dots, x_n^*)$ e.g., $\xi^* = (0, \dots, 0)$

Mobius inversion lemma

For two functions $f, g : 2^{\mathcal{X}} \rightarrow \mathbb{R}$ defined over all subsets $\mathcal{Z} \subseteq \mathcal{X}$ the following are equivalent:

$$\forall \mathcal{Z} \subseteq \mathcal{X} \quad f(\mathcal{Z}) = \sum_{\mathcal{S} \subseteq \mathcal{Z}} g(\mathcal{S})$$

$$\forall \mathcal{Z} \subseteq \mathcal{X} \quad g(\mathcal{Z}) = \sum_{\mathcal{S} \subseteq \mathcal{Z}} (-1)^{|\mathcal{Z}-\mathcal{S}|} f(\mathcal{S}) \quad \text{Mobius inversion}$$

Parametrization: canonical form

Given a *fixed* an assignment $\xi^* = (x_1^*, \dots, x_n^*)$ e.g., $\xi^* = (0, \dots, 0)$

$f(x_{\mathcal{Z}}) \triangleq \log P(x_{\mathcal{Z}}, \xi_{-\mathcal{Z}}^*)$ is defined for all $\mathcal{Z} \subseteq \{1, \dots, N\}$

$$f(x) = \log P(x)$$

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Problem: one factor per subset of nodes

Proof of Hammersly-Clifford theorem:

When \mathcal{Z} is not a clique $\psi_{\mathcal{Z}}(x_{\mathcal{Z}})$ becomes zero.

Proof of the Hammersley-Clifford

Recap:

- fix an assignment
- define factors over each subset of nodes as:

$$\psi(x_{\mathcal{Z}}) = \sum_{\mathcal{S} \subseteq \mathcal{Z}} (-1)^{|\mathcal{Z}-\mathcal{S}|} \log P(x_{\mathcal{S}}, \xi^*_{-\mathcal{S}})$$

- if \mathcal{Z} is not a clique in \mathcal{H} then $\exists_{i,j \in \mathcal{Z}} X_i \perp X_j \mid \mathbf{X} - \{X_i, X_j\}$
 - we can show that $\psi_{\mathcal{Z}}(x_{\mathcal{Z}}) = 0 \quad \forall x_{\mathcal{Z}}$