

Graphical Models

Markov Chain Monte Carlo Inference

Siamak Ravanbakhsh

Winter 2018

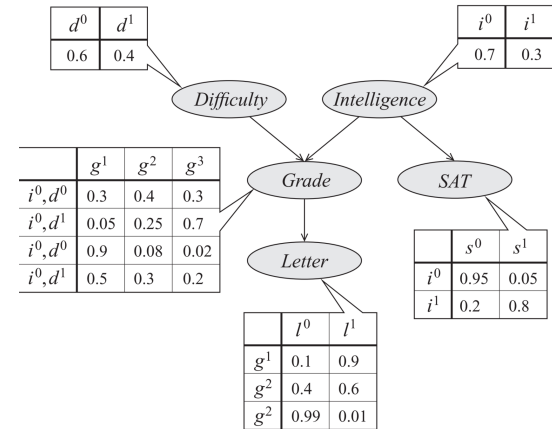
Learning objectives

- Markov chains
- the idea behind Markov Chain Monte Carlo (MCMC)
- two important examples:
 - Gibbs sampling
 - Metropolis-Hastings algorithm

Problem with **likelihood weighting**

Recap

- use a topological ordering
- sample conditioned on the parents
- if observed:
 - keep the observed value
 - update the weight



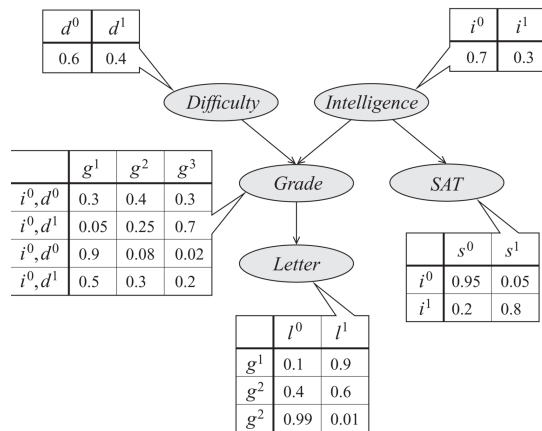
Problem with **likelihood weighting**

Recap

- use a topological ordering
- sample conditioned on the parents
- if observed:
 - keep the observed value
 - update the weight

Issues

- observing the child does not affect the parent's assignment
- only applies to Bayes-nets



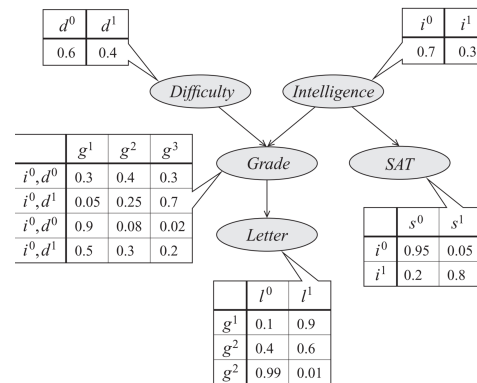
Gibbs sampling

Idea

- iteratively sample each var. condition on its Markov blanket

$$X_i \sim p(x_i \mid X_{MB(i)})$$

- if X_i is observed: keep the observed value



- after many Gibbs sampling iterations $X \sim P$

Gibbs sampling

Idea

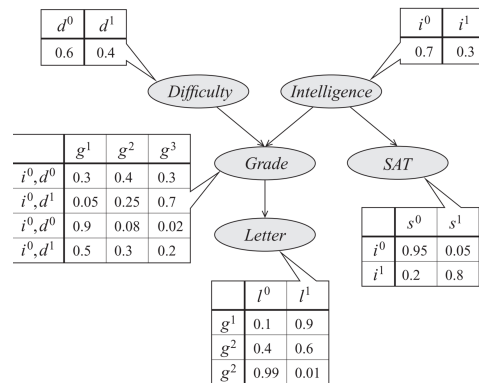
- iteratively sample each var. condition on its Markov blanket

$$X_i \sim p(x_i \mid X_{MB(i)})$$

- if X_i is observed: keep the observed value

equivalent to

- first simplifying the model by removing observed vars
 - sampling from the simplified Gibbs dist.
- after many Gibbs sampling iterations $X \sim P$

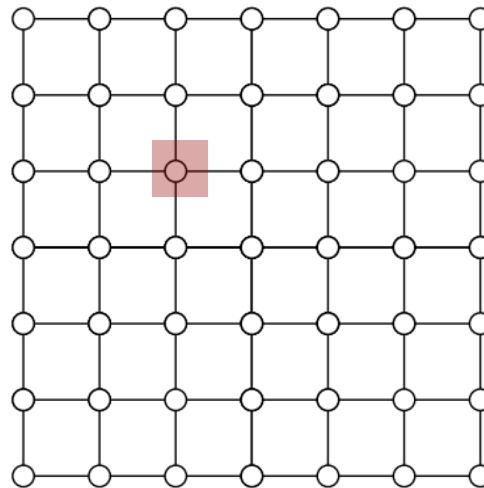


Example: Ising model

recall the Ising model:

$$p(x) \propto \exp(\sum_i x_i h_i + \sum_{i,j \in \mathcal{E}} x_i x_j J_{i,j})$$

$$x_i \in \{-1, +1\}$$



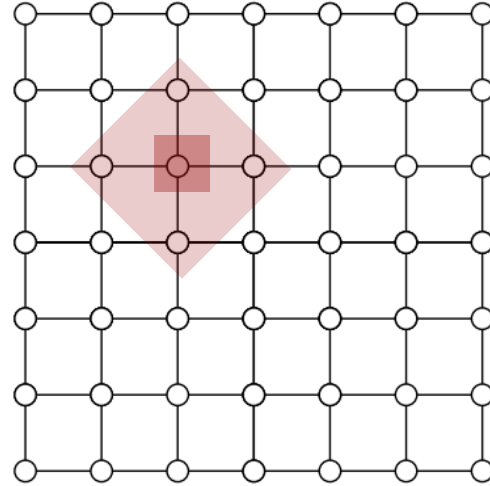
Example: Ising model

recall the Ising model:

$$p(x) \propto \exp(\sum_i x_i h_i + \sum_{i,j \in \mathcal{E}} x_i x_j J_{i,j})$$
$$x_i \in \{-1, +1\}$$

sample each node i :

$$p(x_i = +1 \mid X_{MB(i)}) =$$
$$\frac{\exp(h_i + \sum_{j \in Mb(i)} J_{i,j} X_j)}{\exp(h_i + \sum_{j \in Mb(i)} J_{i,j} X_j) + \exp(-h_i - \sum_{j \in Mb(i)} J_{i,j} X_j)} =$$



Example: Ising model

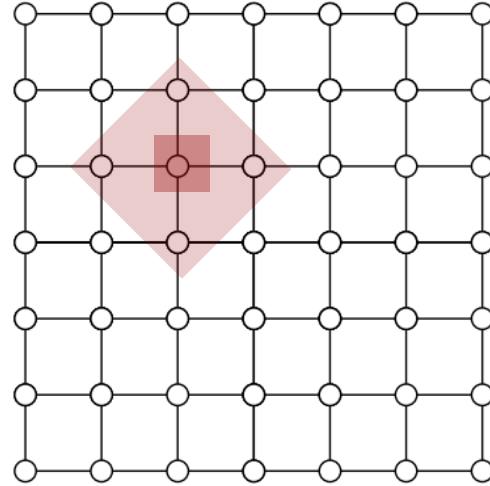
recall the Ising model:

$$p(x) \propto \exp(\sum_i x_i h_i + \sum_{i,j \in \mathcal{E}} x_i x_j J_{i,j})$$
$$x_i \in \{-1, +1\}$$

sample each node i :

$$p(x_i = +1 \mid X_{MB(i)}) =$$
$$\frac{\exp(h_i + \sum_{j \in Mb(i)} J_{i,j} X_j)}{\exp(h_i + \sum_{j \in Mb(i)} J_{i,j} X_j) + \exp(-h_i - \sum_{j \in Mb(i)} J_{i,j} X_j)} =$$

$$\sigma(2h_i + 2 \sum_{j \in Mb(i)} J_{i,j} X_j)$$



Example: Ising model

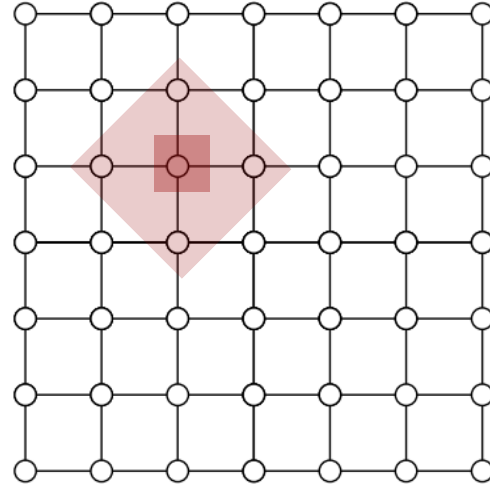
recall the Ising model:

$$p(x) \propto \exp(\sum_i x_i h_i + \sum_{i,j \in \mathcal{E}} x_i x_j J_{i,j})$$
$$x_i \in \{-1, +1\}$$

sample each node i :

$$p(x_i = +1 \mid X_{MB(i)}) = \frac{\exp(h_i + \sum_{j \in Mb(i)} J_{i,j} X_j)}{\exp(h_i + \sum_{j \in Mb(i)} J_{i,j} X_j) + \exp(-h_i - \sum_{j \in Mb(i)} J_{i,j} X_j)} =$$

$$\sigma(2h_i + 2 \sum_{j \in Mb(i)} J_{i,j} X_j) \quad \text{compare with mean-field} \quad \sigma(2h_i + 2 \sum_{j \in Mb(i)} J_{i,j} \mu_j)$$

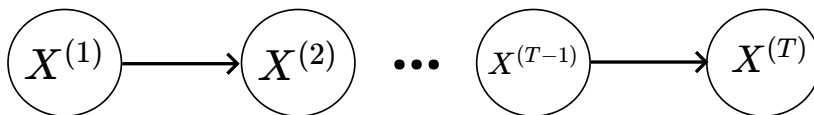


Markov Chain

a sequence of random variables with **Markov property**

$$P(X^{(t)} | X^{(1)}, \dots, X^{(t-1)}) = P(X^{(t)} | X^{(t-1)})$$

its graphical model



many applications:

- **language modeling:** X is a word or a character
- **physics:** with correct choice of X , the world is Markov

Transition model

we assume a **homogeneous** chain: $P(X^{(t)}|X^{(t-1)}) = P(X^{(t+1)}|X^{(t)}) \quad \forall t$

cond. probabilities remain the same across time-steps

notation: conditional probability $P(X^{(t)} = x | X^{(t-1)} = x') = T(x, x')$

is called the **transition model**

think of this as a matrix T

Transition model

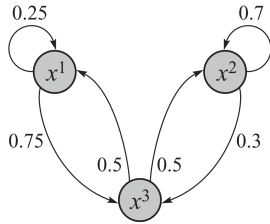
we assume a **homogeneous** chain: $P(X^{(t)}|X^{(t-1)}) = P(X^{(t+1)}|X^{(t)}) \quad \forall t$
cond. probabilities remain the same across time-steps

notation: conditional probability $P(X^{(t)} = x | X^{(t-1)} = x') = T(x, x')$

is called the **transition model**

think of this as a matrix T

state-transition diagram



its transition matrix

$$T = \begin{bmatrix} .25 & 0 & .75 \\ 0 & .7 & .3 \\ .5 & .5 & 0 \end{bmatrix}$$

Transition model

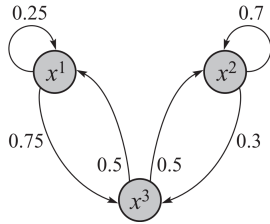
we assume a **homogeneous** chain: $P(X^{(t)}|X^{(t-1)}) = P(X^{(t+1)}|X^{(t)}) \quad \forall t$
cond. probabilities remain the same across time-steps

notation: conditional probability $P(X^{(t)} = x|X^{(t-1)} = x') = T(x, x')$

is called the **transition model**

think of this as a matrix T

state-transition diagram



its transition matrix

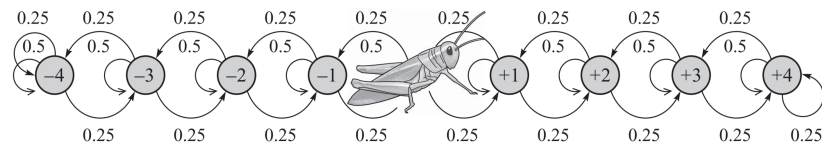
$$T = \begin{bmatrix} .25 & 0 & .75 \\ 0 & .7 & .3 \\ .5 & .5 & 0 \end{bmatrix}$$

evolving the distribution $P(X^{(t+1)} = x) = \sum_{x' \in \text{Val}(X)} P(X^{(t)} = x')T(x', x)$

Markov Chain Monte Carlo (MCMC)

Example

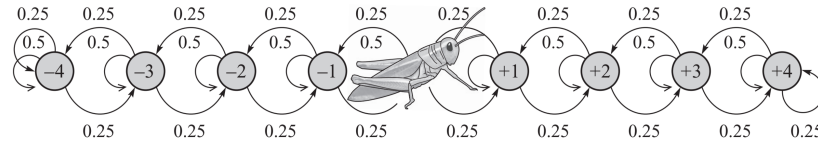
state-transition diagram for grasshopper random walk



initial distribution $P^{(0)}(X = 0) = 1$

Markov Chain Monte Carlo (MCMC)

Example state-transition diagram for grasshopper random walk



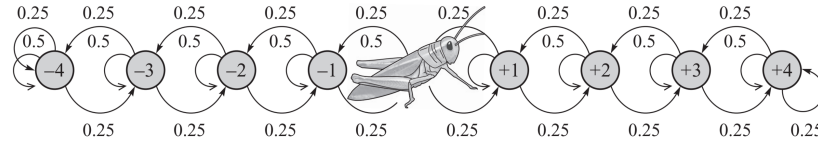
initial distribution $P^{(0)}(X = 0) = 1$

after $t=50$ steps, the distribution is almost uniform $P^t(x) \approx \frac{1}{9} \quad \forall x$

Markov Chain Monte Carlo (MCMC)

Example

state-transition diagram for grasshopper random walk



initial distribution $P^{(0)}(X = 0) = 1$

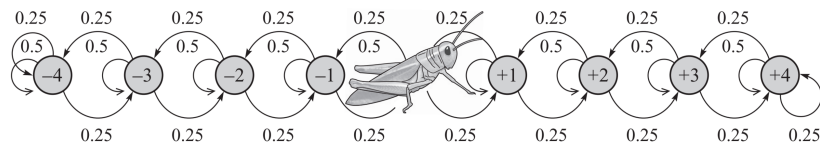
after $t=50$ steps, the distribution is almost uniform $P^t(x) \approx \frac{1}{9} \quad \forall x$

use the chain to sample from the uniform distribution $P^t(X) \approx \frac{1}{9}$

Markov Chain Monte Carlo (MCMC)

Example

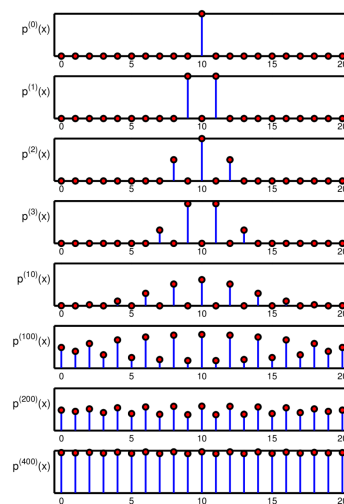
state-transition diagram for grasshopper random walk



initial distribution $P^{(0)}(X = 0) = 1$

after $t=50$ steps, the distribution is almost uniform $P^t(x) \approx \frac{1}{9} \quad \forall x$

use the chain to sample from the uniform distribution $P^t(X) \approx \frac{1}{9}$



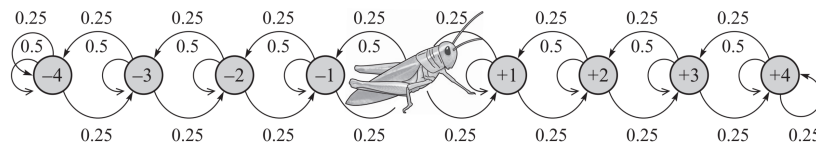
why is it uniform?

(mixing image: Murphy's book)

Markov Chain Monte Carlo (MCMC)

Example

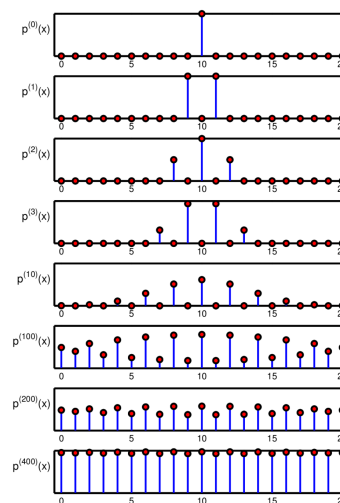
state-transition diagram for grasshopper random walk



initial distribution $P^{(0)}(X = 0) = 1$

after $t=50$ steps, the distribution is almost uniform $P^t(x) \approx \frac{1}{9} \quad \forall x$

use the chain to sample from the uniform distribution $P^t(X) \approx \frac{1}{9}$



why is it uniform?

MCMC

generalize this idea beyond uniform dist.

- we **want to sample** from P^*
- pick the **transition model** such that $P^\infty(X) = P^*(X)$

Stationary distribution

given a transition model $T(x, x')$ if the chain **converges**:

global balance equation $P^{(t)}(x) \approx P^{(t+1)}(x) = \sum_{x'} P^{(t)}(x')T(x', x)$

Stationary distribution

given a transition model $T(x, x')$ if the chain **converges**:

global balance equation $P^{(t)}(x) \approx P^{(t+1)}(x) = \sum_{x'} P^{(t)}(x')T(x', x)$

this condition defines the **stationary distribution**: π

$$\pi(X = x) = \sum_{x' \in \text{Val}(X)} \pi(X = x')T(x, x')$$

Stationary distribution

given a transition model $T(x, x')$ if the chain **converges**:

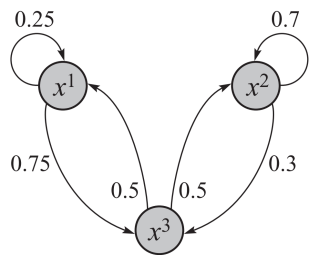
global balance equation $P^{(t)}(x) \approx P^{(t+1)}(x) = \sum_{x'} P^{(t)}(x')T(x', x)$

this condition defines the **stationary distribution**: π

$$\pi(X = x) = \sum_{x' \in \text{Val}(X)} \pi(X = x')T(x, x')$$

Example

finding the stationary dist.



$$\pi(x^1) = .25\pi(x^1) + .5\pi(x^3)$$

$$\pi(x^2) = .7\pi(x^2) + .5\pi(x^3)$$

$$\pi(x^3) = .75\pi(x^1) + .3\pi(x^2)$$

$$\pi(x^1) + \pi(x^2) + \pi(x^3) = 1$$



$$\pi(x^1) = .2$$

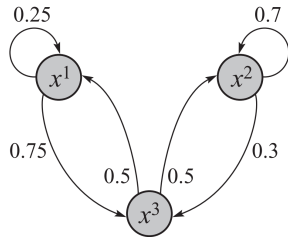
$$\pi(x^2) = .5$$

$$\pi(x^3) = .3$$

Stationary distribution as an eigenvector

Example

finding the stationary dist.



$$\pi(x^1) = .25\pi(x^1) + .5\pi(x^3)$$

$$\pi(x^2) = .7\pi(x^2) + .5\pi(x^3)$$

$$\pi(x^3) = .75\pi(x^1) + .3\pi(x^2)$$

$$\pi(x^1) + \pi(x^2) + \pi(x^3) = 1$$



$$\pi(x^1) = .2$$

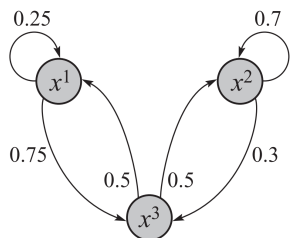
$$\pi(x^2) = .5$$

$$\pi(x^3) = .3$$

Stationary distribution as an eigenvector

Example

finding the stationary dist.



$$\pi(x^1) = .25\pi(x^1) + .5\pi(x^3)$$

$$\pi(x^2) = .7\pi(x^2) + .5\pi(x^3)$$

$$\pi(x^3) = .75\pi(x^1) + .3\pi(x^2)$$

$$\pi(x^1) + \pi(x^2) + \pi(x^3) = 1$$



$$\pi(x^1) = .2$$

$$\pi(x^2) = .5$$

$$\pi(x^3) = .3$$

viewing $T(\cdot, \cdot)$ as a matrix and $P^t(x)$ as a vector

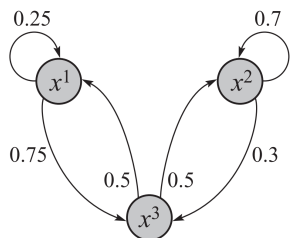
- evolution of dist $P^t(x) : P^{(t+1)} = T^\top P^{(t)}$
- multiple steps: $P^{(t+m)} = (T^\top)^m P^{(t)}$

$$\begin{matrix} \begin{bmatrix} .25 & 0 & .5 \\ 0 & .7 & .5 \\ .75 & .3 & 0 \end{bmatrix} & \begin{bmatrix} .2 \\ .5 \\ .3 \end{bmatrix} & = & \begin{bmatrix} .2 \\ .5 \\ .3 \end{bmatrix} \\ T^\top & \pi & & \pi \end{matrix}$$

Stationary distribution as an eigenvector

Example

finding the stationary dist.



$$\pi(x^1) = .25\pi(x^1) + .5\pi(x^3)$$

$$\pi(x^2) = .7\pi(x^2) + .5\pi(x^3)$$

$$\pi(x^3) = .75\pi(x^1) + .3\pi(x^2)$$

$$\pi(x^1) + \pi(x^2) + \pi(x^3) = 1$$



$$\pi(x^1) = .2$$

$$\pi(x^2) = .5$$

$$\pi(x^3) = .3$$

viewing $T(\cdot, \cdot)$ as a matrix and $P^t(x)$ as a vector

- evolution of dist $P^t(x) : P^{(t+1)} = T^\top P^{(t)}$
- multiple steps: $P^{(t+m)} = (T^\top)^m P^{(t)}$
- for stationary dist: $\pi = T^\top \pi$

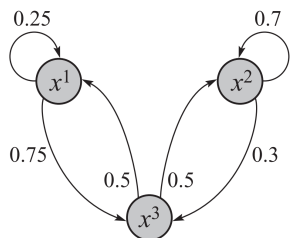
$$\begin{bmatrix} .25 & 0 & .5 \\ 0 & .7 & .5 \\ .75 & .3 & 0 \end{bmatrix} \begin{bmatrix} .2 \\ .5 \\ .3 \end{bmatrix} = \begin{bmatrix} .2 \\ .5 \\ .3 \end{bmatrix}$$

$T^\top \quad \pi \quad \pi$

Stationary distribution as an eigenvector

Example

finding the stationary dist.



$$\pi(x^1) = .25\pi(x^1) + .5\pi(x^3)$$

$$\pi(x^2) = .7\pi(x^2) + .5\pi(x^3)$$

$$\pi(x^3) = .75\pi(x^1) + .3\pi(x^2)$$

$$\pi(x^1) + \pi(x^2) + \pi(x^3) = 1$$



$$\pi(x^1) = .2$$

$$\pi(x^2) = .5$$

$$\pi(x^3) = .3$$

viewing $T(\cdot, \cdot)$ as a matrix and $P^t(x)$ as a vector

- evolution of dist $P^t(x) : P^{(t+1)} = T^\top P^{(t)}$

- multiple steps: $P^{(t+m)} = (T^\top)^m P^{(t)}$

- for stationary dist: $\pi = T^\top \pi$

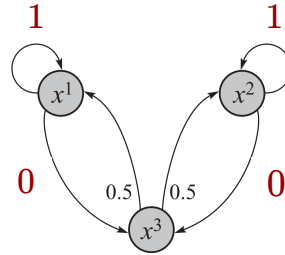
- π is an eigenvector of T^\top with eigenvalue 1 (produce it using the power method)

$$\begin{matrix} \begin{bmatrix} .25 & 0 & .5 \\ 0 & .7 & .5 \\ .75 & .3 & 0 \end{bmatrix} & \begin{bmatrix} .2 \\ .5 \\ .3 \end{bmatrix} & = & \begin{bmatrix} .2 \\ .5 \\ .3 \end{bmatrix} \\ T^\top & \pi & & \pi \end{matrix}$$

Stationary distribution: existence & uniqueness

irreducible

- we should be able to reach any x' from any x
- otherwise, π is not unique



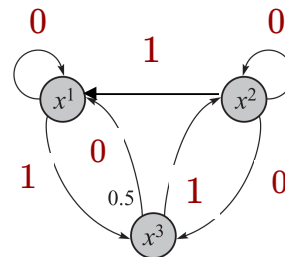
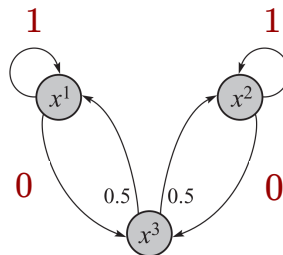
Stationary distribution: existence & uniqueness

irreducible

- we should be able to reach any x' from any x
- otherwise, π is not unique

aperiodic

- the chain should not have a fixed cyclic behavior
- otherwise, the chain does not converge (it oscillates)



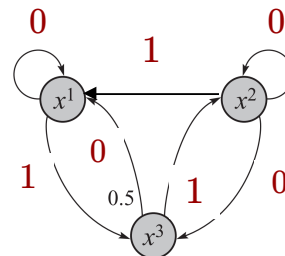
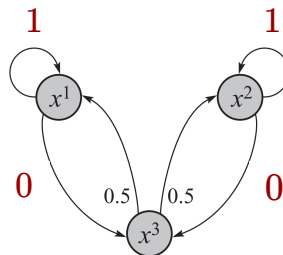
Stationary distribution: existence & uniqueness

irreducible

- we should be able to reach any x' from any x
- otherwise, π is not unique

aperiodic

- the chain should not have a fixed cyclic behavior
- otherwise, the chain does not converge (it oscillates)



every **aperiodic** and **irreducible** chain (with a finite domain) has a unique limiting distribution π

such that
$$\pi(X = x) = \sum_{x' \in \text{Val}(X)} \pi(X = x') T(x, x')$$

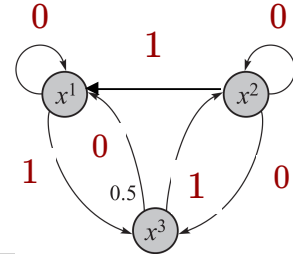
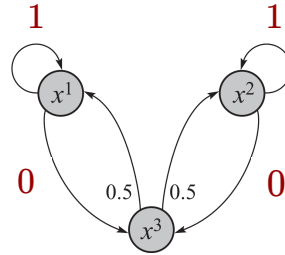
Stationary distribution: existence & uniqueness

irreducible

- we should be able to reach any x' from any x
- otherwise, π is not unique

aperiodic

- the chain should not have a fixed cyclic behavior
- otherwise, the chain does not converge (it oscillates)



every **aperiodic** and **irreducible** chain (with a finite domain) has a unique limiting distribution π

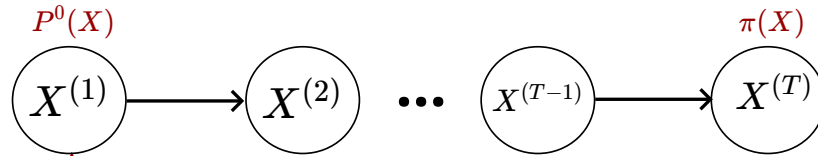
such that
$$\pi(X = x) = \sum_{x' \in Val(X)} \pi(X = x') T(x, x')$$

regular chain

a **sufficient condition**: there exists a K , such that the probability of reaching any destination from any source in K steps is positive (applies to discrete & continuous domains)

MCMC in graphical models

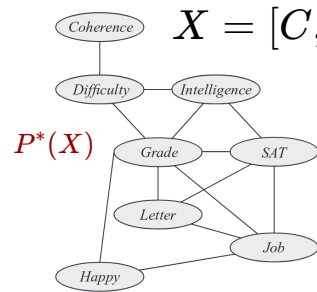
distinguishing the "graphical models" involved



1: the Markov chain

2: state-transition diagram (not shown)
that has exponentially many nodes

$$\#nodes = |Val(X)|$$

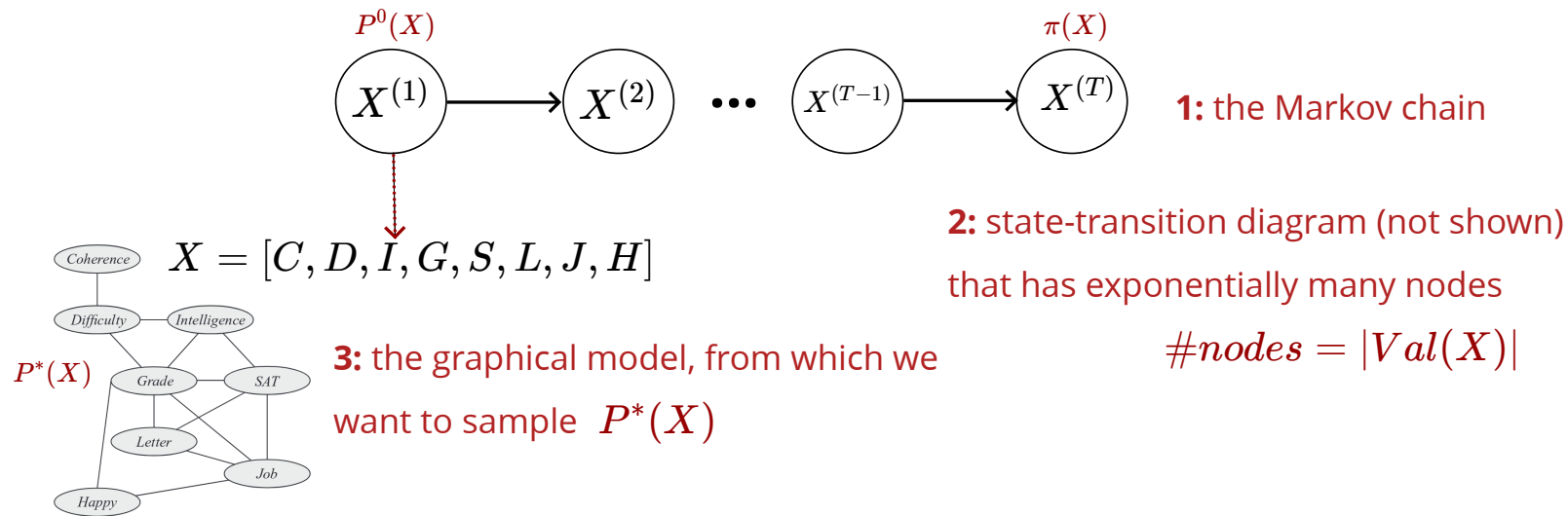


$$X = [C, D, I, G, S, L, J, H]$$

3: the graphical model, from which we
want to sample $P^*(X)$

MCMC in graphical models

distinguishing the "graphical models" involved



objective: design the Markov chain transition so that $\pi(X) = P^*(X)$

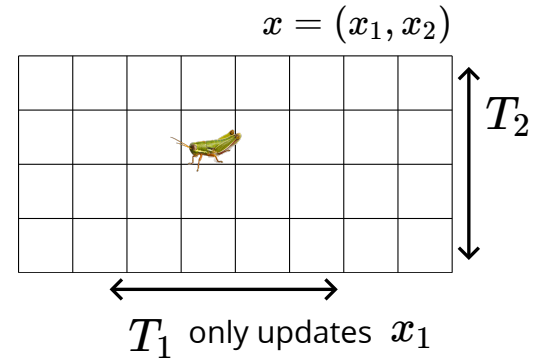
Multiple transition models

idea

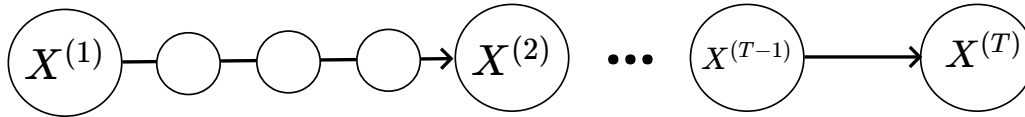
aka, **kernels**

have multiple transition models $T_1(x, x'), T_2(x, x'), \dots, T_n(x, x')$

each making local changes to x



using a single kernel we may not be able to visit all the states while their combination is "ergodic"



Multiple transition models

idea

aka, **kernels**

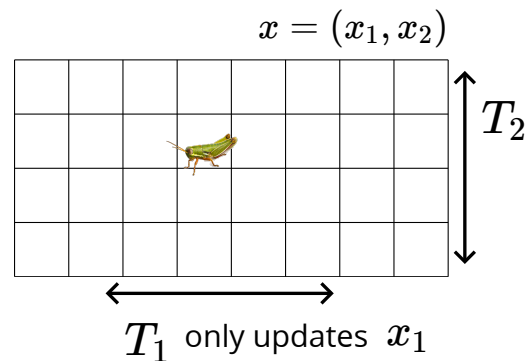
have multiple transition models $T_1(x, x'), T_2(x, x'), \dots, T_n(x, x')$

each making local changes to x

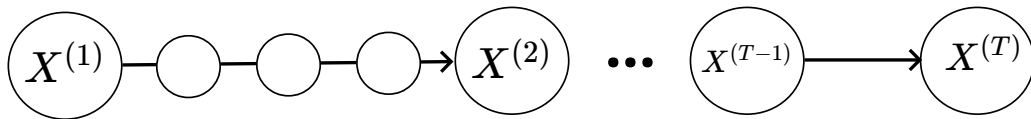
if $\pi(X = x) = \sum_{x' \in \text{Val}(X)} \pi(X = x') T_k(x, x') \quad \forall k$

then we can combine the kernels:

- mixing them $T(x, x') = \sum_k p(k) T_k(x, x')$
- cycling them $T(x, x') = \int_{x^{[1]}, x^{[2]}, \dots, x^{[n]}} T_1(x, x^{[1]}) T_2(x^{[1]}, x^{[2]}), \dots, T_n(x^{[n-1]}, x') dx^{[1]} dx^{[2]} \dots dx^{[n]}$



using a single kernel we may not be able to visit all the states while their combination is "ergodic"

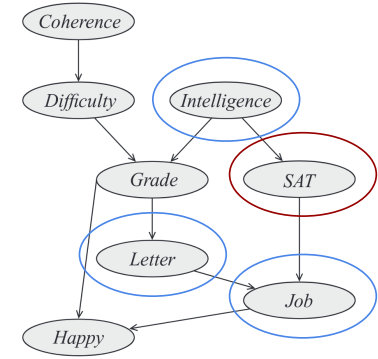
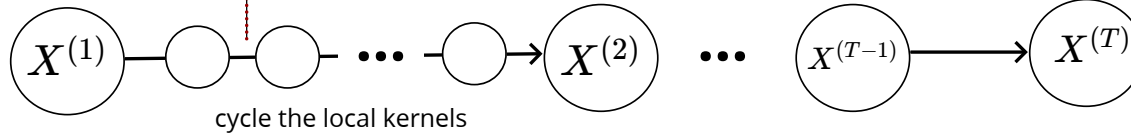


Revisiting Gibbs sampling

one kernel for each variable
perform local, conditional updates

$$T_i(x, x') = P^*(x_i | x'_{-i}) \mathbb{I}(x_{-i} = x'_{-i})$$

$$P^*(x_i | x'_{-i}) = P^*(x_i | x'_{MB(i)})$$



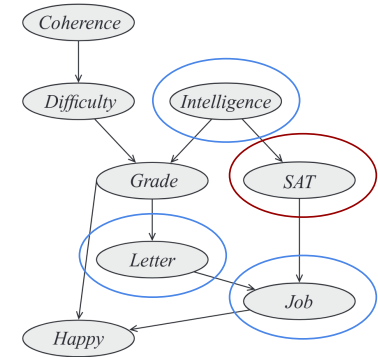
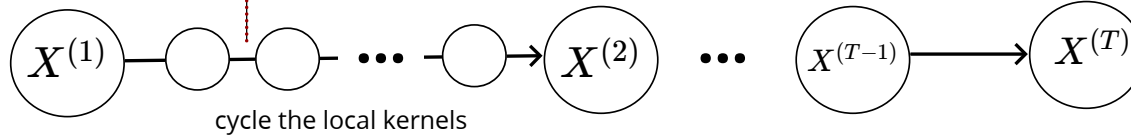
$\pi(X) = P^*(X)$ is the stationary dist. for this Markov chain

Revisiting Gibbs sampling

one kernel for each variable
perform local, conditional updates

$$T_i(x, x') = P^*(x_i | x'_{-i}) \mathbb{I}(x_{-i} = x'_{-i})$$

$$P^*(x_i | x'_{-i}) = P^*(x_i | x'_{MB(i)})$$



$\pi(X) = P^*(X)$ is the stationary dist. for this Markov chain

if $P^*(x) > 0 \quad \forall x$ then this chain is **regular**

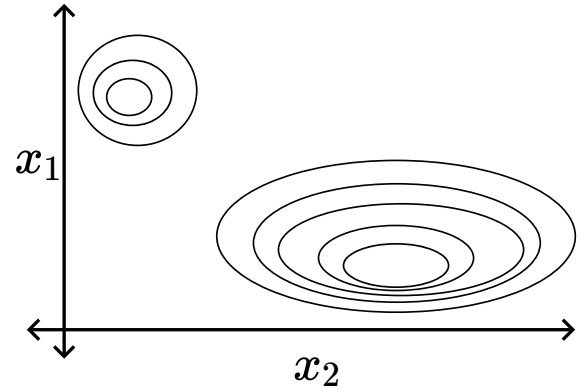
i.e., converges to its unique stationary dist.

Block Gibbs sampling

local moves can get stuck in modes of $P^*(X)$

updates using $P(x_1 | x_2), P(x_2 | x_1)$ will have problem

exploring these modes



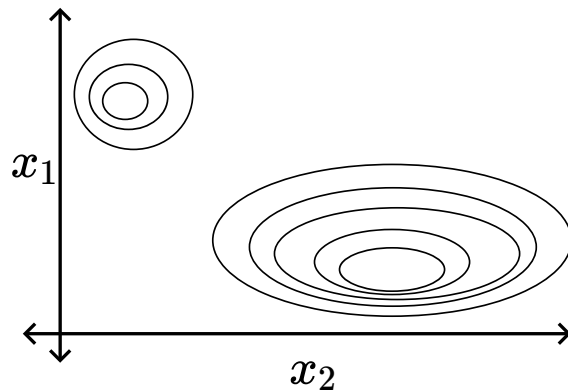
Block Gibbs sampling

local moves can get stuck in modes of $P^*(X)$

updates using $P(x_1 | x_2), P(x_2 | x_1)$ will have problem

exploring these modes

idea: each kernel updates a block of variables

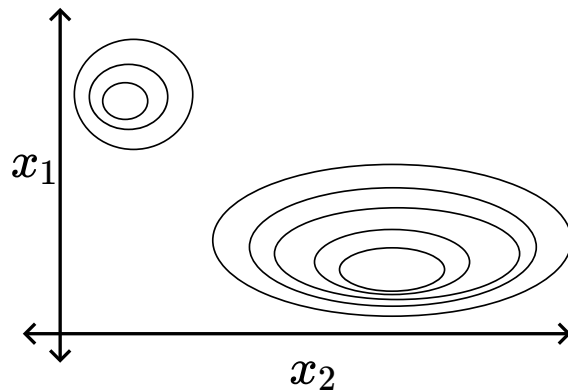


Block Gibbs sampling

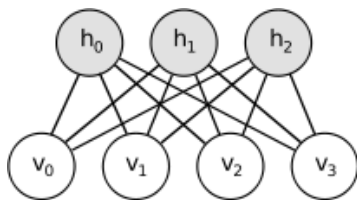
local moves can get stuck in modes of $P^*(X)$

updates using $P(x_1 | x_2), P(x_2 | x_1)$ will have problem

exploring these modes



idea: each kernel updates a block of variables



Restricted Boltzmann Machine (RBM)

- update h given v $P(h_0, h_1, h_2 | v_0, \dots, v_3) = P(h_0 | v) \dots P(h_2 | v)$
- update v given h $P(v_0, \dots, v_3 | h_0, h_1, h_2) = P(v_0 | h) \dots P(v_3 | h)$

Detailed balance

A Markov chain is **reversible** if for a unique π

detailed balance $\pi(x)T(x, x') = \pi(x')T(x', x) \quad \forall x, x'$
same frequency in both directions

Detailed balance

A Markov chain is **reversible** if for a unique π

$$\text{detailed balance } \pi(x)T(x, x') = \pi(x')T(x', x) \quad \forall x, x'$$

same frequency in both directions

$$\int_{x'} \pi(x)T(x, x')dx' = \pi(x) \int_{x'} T(x, x')dx' = \pi(x) \quad = \quad \int_{x'} \pi(x')T(x', x)dx'$$

left-hand side **global balance** *right-hand side*

Detailed balance

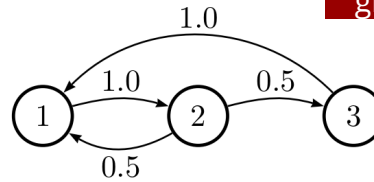
A Markov chain is **reversible** if for a unique π

detailed balance $\pi(x)T(x, x') = \pi(x')T(x', x) \quad \forall x, x'$
same frequency in both directions

$$\int_{x'} \pi(x)T(x, x')dx' = \pi(x) \int_{x'} T(x, x')dx' = \pi(x) \quad \equiv \quad \int_{x'} \pi(x')T(x', x)dx'$$

left-hand side global balance *right-hand side*

detailed balance is a stronger condition



$$\pi = [.4, .4, .2]$$

global balance ✓

detailed balance ✗

Detailed balance

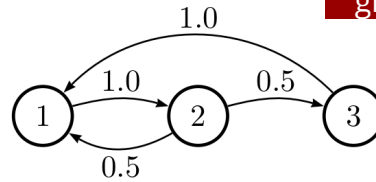
A Markov chain is **reversible** if for a unique π

detailed balance $\pi(x)T(x, x') = \pi(x')T(x', x) \quad \forall x, x'$
same frequency in both directions

$$\int_{x'} \pi(x)T(x, x')dx' = \pi(x) \int_{x'} T(x, x')dx' = \pi(x) \quad \equiv \quad \int_{x'} \pi(x')T(x', x)dx'$$

left-hand side **global balance** *right-hand side*

detailed balance is a stronger condition



$$\pi = [.4, .4, .2]$$

global balance ✓

detailed balance ✗

if Markov chain is **regular** and π satisfies **detailed balance**,
then π is the unique stationary distribution

Detailed balance

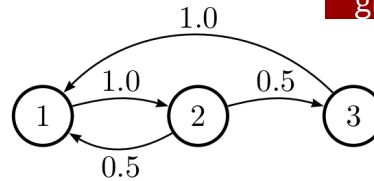
A Markov chain is **reversible** if for a unique π

detailed balance $\pi(x)T(x, x') = \pi(x')T(x', x) \quad \forall x, x'$
same frequency in both directions

$$\int_{x'} \pi(x)T(x, x')dx' = \pi(x) \int_{x'} T(x, x')dx' = \pi(x) \quad \equiv \quad \int_{x'} \pi(x')T(x', x)dx'$$

left-hand side **global balance** *right-hand side*

detailed balance is a stronger condition



$$\pi = [.4, .4, .2]$$

global balance ✓

detailed balance ✗

if Markov chain is **regular** and π satisfies **detailed balance**, then π is the unique stationary distribution

- *analogous to the theorem for global balance*
- *checking for detailed balance is easier*

Detailed balance

A Markov chain is **reversible** if for a unique π

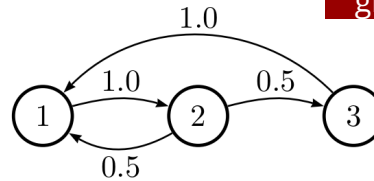
detailed balance $\pi(x)T(x, x') = \pi(x')T(x', x) \quad \forall x, x'$
same frequency in both directions

$$\int_{x'} \pi(x)T(x, x')dx' = \pi(x) \int_{x'} T(x, x')dx' = \pi(x) \quad = \quad \int_{x'} \pi(x')T(x', x)dx'$$

left-hand side

global balance *right-hand side*

detailed balance is a stronger condition



$$\pi = [.4, .4, .2]$$

global balance ✓

detailed balance ✗

if Markov chain is **regular** and π satisfies **detailed balance**,
then π is the unique stationary distribution

- *analogous to the theorem for global balance*
- *checking for detailed balance is easier*

what happens if T is symmetric?

(example: Murphy's book)

Using a proposal for the chain

Given P^* design a chain to sample from P^*

idea

Using a proposal for the chain

Given P^* design a chain to sample from P^*

idea

- use a proposal transition $T^q(x, x')$
- we can sample from $T^q(x, \cdot)$
- $T^q(x, x')$ is a regular chain (reaching every state in K steps has a non-zero probability)

Using a proposal for the chain

Given P^* design a chain to sample from P^*

idea

- use a proposal transition $T^q(x, x')$
- we can sample from $T^q(x, \cdot)$
- $T^q(x, x')$ is a regular chain (reaching every state in K steps has a non-zero probability)
- accept the proposed move with probability $A(x, x')$
 - to achieve detailed balance

Metropolis algorithm

- use a proposal transition $T^q(x, x')$
- we can sample from $T^q(x, \cdot)$
- $T^q(x, x')$ is a regular chain (reaching every state in K steps has a non-zero probability)
- accept the proposed move with probability $A(x, x')$

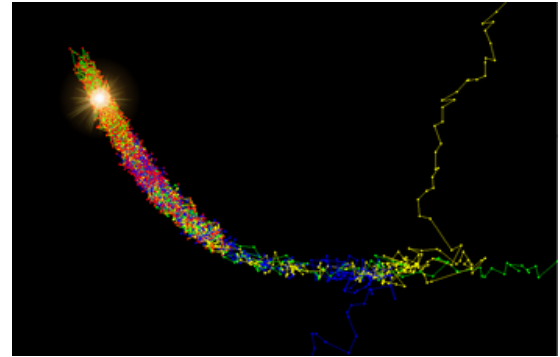
- to achieve detailed balance

- **proposal is symmetric** $T(x, x') = T(x', x)$

$$A(x, x') \triangleq \min\left(1, \frac{p(x')}{p(x)}\right)$$

accepts the move if it increases P^*

may accept it otherwise



(image: Wikipedia)

Metropolis-Hastings algorithm

if the proposal is **NOT symmetric**, then $A(x, x') \triangleq \min\left(1, \frac{p(x')T^q(x', x)}{p(x)T^q(x, x')}\right)$

Metropolis-Hastings algorithm

if the proposal is **NOT symmetric**, then $A(x, x') \triangleq \min\left(1, \frac{p(x')T^q(x', x)}{p(x)T^q(x, x')}\right)$

why does it sample from P^* ?

Metropolis-Hastings algorithm

if the proposal is **NOT symmetric**, then $A(x, x') \triangleq \min(1, \frac{p(x')T^q(x', x)}{p(x)T^q(x, x')})$

why does it sample from P^* ?

$T(x, x') = T^q(x, x')A(x, x') \quad \forall x \neq x' \quad \Big| \quad \text{move to a different state is accepted}$

Metropolis-Hastings algorithm

if the proposal is **NOT symmetric**, then $A(x, x') \triangleq \min(1, \frac{p(x')T^q(x', x)}{p(x)T^q(x, x')})$

why does it sample from P^* ?

$T(x, x') = T^q(x, x')A(x, x') \quad \forall x \neq x' \quad | \quad \text{move to a different state is accepted}$

$T(x, x) = T^q(x, x) + \sum_{x \neq x'} (1 - A(x, x'))T(x, x') \quad | \quad \begin{array}{l} \text{proposal to stay is always accepted} \\ \text{move to a new state is rejected} \end{array}$

Metropolis-Hastings algorithm

if the proposal is **NOT symmetric**, then $A(x, x') \triangleq \min(1, \frac{p(x')T^q(x',x)}{p(x)T^q(x,x')})$

why does it sample from P^* ?

$T(x, x') = T^q(x, x')A(x, x') \quad \forall x \neq x' \quad \Big| \quad \text{move to a different state is accepted}$

$T(x, x) = T^q(x, x) + \sum_{x' \neq x} (1 - A(x, x'))T(x, x') \quad \Big| \quad \begin{array}{l} \text{proposal to stay is always accepted} \\ \text{move to a new state is rejected} \end{array}$

substitute this into **detailed balance** (does it hold?)

$$\pi(x)T^q(x, x')A(x, x') \stackrel{?}{=} \pi(x')T^q(x', x)A(x', x)$$
$$\min(1, \frac{\pi(x')T^q(x',x)}{\pi(x)T^q(x,x')}) \qquad \min(1, \frac{\pi(x)T^q(x,x')}{\pi(x')T^q(x',x)})$$



Metropolis-Hastings algorithm

if the proposal is **NOT symmetric**, then $A(x, x') \triangleq \min(1, \frac{p(x')T^q(x', x)}{p(x)T^q(x, x')})$

why does it sample from P^* ?

$T(x, x') = T^q(x, x')A(x, x') \quad \forall x \neq x' \quad \Big| \quad \text{move to a different state is accepted}$

$T(x, x) = T^q(x, x) + \sum_{x' \neq x} (1 - A(x, x'))T(x, x') \quad \Big| \quad \begin{array}{l} \text{proposal to stay is always accepted} \\ \text{move to a new state is rejected} \end{array}$

substitute this into **detailed balance** (does it hold?)

$$\pi(x)T^q(x, x')A(x, x') \stackrel{?}{=} \pi(x')T^q(x', x)A(x', x)$$
$$\min(1, \frac{\pi(x')T^q(x', x)}{\pi(x)T^q(x, x')}) \qquad \min(1, \frac{\pi(x)T^q(x, x')}{\pi(x')T^q(x', x)})$$

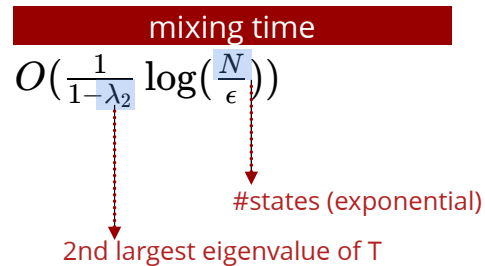


Gibbs sampling is a special case, with $A(x, x') = 1$ all the time!

Sampling from the chain

at the limit $T \rightarrow \infty$, $P^\infty = \pi = P^*$

how long should we wait for $D(P^T, \pi) < \epsilon$?



Sampling from the chain

at the limit $T \rightarrow \infty$, $P^\infty = \pi = P^*$

how long should we wait for $D(P^T, \pi) < \epsilon$?

- run the chain for a **burn-in** period (T steps)
- collect samples (few more steps)
- multiple restarts can ensure a better coverage

mixing time

$$O\left(\frac{1}{1-\lambda_2} \log\left(\frac{N}{\epsilon}\right)\right)$$

↓ #states (exponential)

↓ 2nd largest eigenvalue of T

Sampling from the chain

at the limit $T \rightarrow \infty$, $P^\infty = \pi = P^*$

how long should we wait for $D(P^T, \pi) < \epsilon$?

- run the chain for a **burn-in** period (T steps)
- collect samples (few more steps)
- multiple restarts can ensure a better coverage

Example Potts model

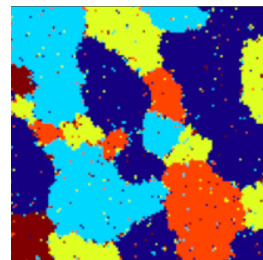
- model $p(x) \propto \exp(\sum_i h(x_i) + \sum_{i,j \in \mathcal{E}} .66\mathbb{I}(x_j = x_i))$
- $|Val(X)| = 5$ different colors
- 128x128 grid
- Gibbs sampling

mixing time

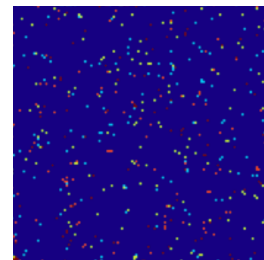
$$O\left(\frac{1}{1-\lambda_2} \log\left(\frac{N}{\epsilon}\right)\right)$$

2nd largest eigenvalue of T

#states (exponential)



200 iterations



10,000 iterations

image : Murphy's book

Diagnosing convergence

- heuristics for diagnosing non-convergence
- difficult problem
- run multiple chains (compare sample statistics)
- auto-correlation within each chain

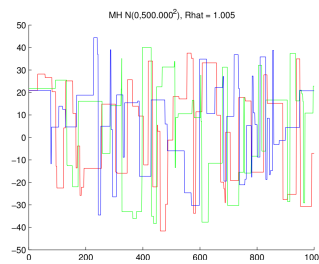
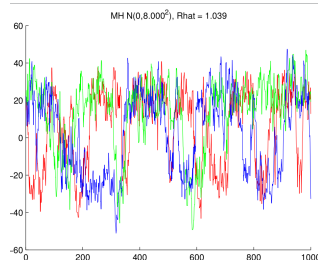
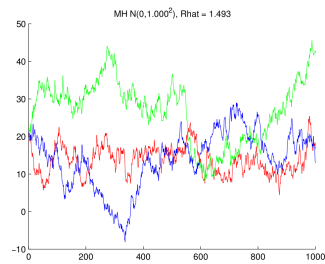
Diagnosing convergence

- heuristics for diagnosing **non**-convergence
- difficult problem
- run multiple chains (compare sample statistics)
- auto-correlation within each chain

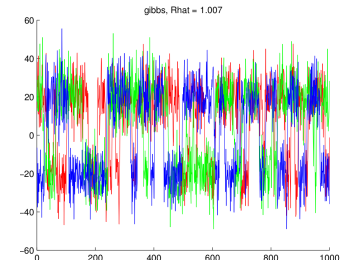
example

sampling from a mixture of two 1D Gaussians (3 chains: colors)

metropolis-hastings (MH) with increasing step sizes for the proposal



Gibbs sampling



Summary

Markov Chain:

- can model the "evolution" of an initial distribution
- converges to a **stationary distribution**

Summary

Markov Chain:

- can model the "evolution" of an initial distribution
- converges to a **stationary distribution**

Markov Chain Monte Carlo:

- **design** a Markov chain: stationary dist. is what we want to sample
- run the chain to produce samples

Summary

Markov Chain:

- can model the "evolution" of an initial distribution
- converges to a **stationary distribution**

Markov Chain Monte Carlo:

- **design** a Markov chain: stationary dist. is what we want to sample
- run the chain to produce samples

Two MCMC methods:

- Gibbs sampling
- Metropolis-Hastings