

Lecture Stat 302

Introduction to Probability - Slides 20

AD

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Conditional Distributions: Discrete Case

- Given a joint p.m.f. for two r.v. X, Y it is possible to compute the conditional p.m.f. X given $Y = y$.
- Assume X, Y are discrete-valued r.v. with a joint p.m.f. $p(x, y)$ then the conditional p.m.f. of X given $Y = y$ is

$$p_{X|Y}(x|y) = \frac{p(x, y)}{p_Y(y)} = \frac{p_{Y|X}(y|x) p_X(x)}{p_Y(y)}$$

- The conditional expectation of $g(X)$ is given by

$$E(g(X) | Y = y) = \sum_x g(x) \cdot p_{X|Y}(x|y).$$

- $E(g(X) | Y = y)$ is a function of y and $E(g(X) | Y)$ is a r.v..

Example: Warm-up

- Let X, Y be two discrete r.v. of joint p.m.f.

$$p(x, y) = \frac{(x + y)}{21}$$

for $x = 1, 2, 3$ and $y = 1, 2$.

- (a) Show that X and Y are not independent.
- (b) Compute $p_{X|Y}(x|y)$.
- (c) Compute $E(X|Y = y)$.

Example: Warm-up

- (a) X and Y are independent if and only if $p(x, y) = p_X(x) p_Y(y)$. We have

$$\begin{aligned} p_X(x) &= \sum_y p(x, y) = p(x, 1) + p(x, 2) \\ &= \frac{(x+1)}{21} + \frac{(x+2)}{21} = \frac{2x}{21} + \frac{1}{7} \end{aligned}$$

and

$$\begin{aligned} p_Y(y) &= \sum_x p(x, y) = p(1, y) + p(2, y) + p(3, y) \\ &= \frac{(1+y)}{21} + \frac{(2+y)}{21} + \frac{(3+y)}{21} = \frac{2}{7} + \frac{3y}{21}. \end{aligned}$$

Clearly $p(x, y) \neq p_X(x) p_Y(y)$ so X, Y are not independent.

- Remark: A safety check consists of checking that $\sum_x p_X(x) = \sum_y p_Y(y) = 1$.

Example: Warm-up

- (b) The conditional p.m.f. is given by

$$\begin{aligned} p_{X|Y}(x|y) &= \frac{p(x,y)}{p_Y(y)} = \frac{(x+y)/21}{(6+3y)/21} \\ &= \frac{x+y}{6+3y} \end{aligned}$$

- (c) The conditional mean is given by

$$\begin{aligned} E(X|Y=y) &= \sum_x x \cdot p_{X|Y}(x|y) \\ &= \frac{1+y}{6+3y} + 2 \times \frac{2+y}{6+3y} + 3 \times \frac{3+y}{6+3y} \\ &= \frac{14+6y}{6+3y} \end{aligned}$$

Example: Insurance Company

- An insurance company provides insurance to three groups of staff with the following characteristics

Age	# ind.	Proba claim (in a year)	Expect amount (in a year)
<30	20	0.02	\$500
30 to 50	50	0.04	\$1000
>50	30	0.06	\$1500

- What is the proba that a randomly selected individual is below the age of 30 and will make a claim in a year?
- What is the proba that a randomly selected individual will not make any claim in a year?
- Given that the randomly selected individual has made a claim in a year, what are the proba that he/she is below 30, aged 30 to 50 or older than 50?
- Given that the randomly selected individual has made a claim in a year and is older than 30, what is the expected claim?

Example: Insurance Company

① $\Pr(" < 30" \cap \text{"Claim"}) = \frac{20}{100} \times 0.02 = 0.004.$

② We have

$$\begin{aligned}\Pr(\text{No Claim}) &= \Pr(\text{No Claim} | " < 30") \Pr(" < 30") \\ &+ \Pr(\text{No Claim} | " 30 - 50") \Pr(" 30 - 50") \\ &+ \Pr(\text{No Claim} | " > 50") \Pr(" > 50") \\ &= (1 - 0.02) \frac{20}{100} + (1 - 0.04) \frac{50}{100} + (1 - 0.06) \frac{30}{100} \\ &= 0.958\end{aligned}$$

③ We want

$$\begin{aligned}\Pr(" < 30" | \text{Claim}) &= \frac{\Pr(\text{Claim} | " < 30") \Pr(" < 30")}{\Pr(\text{Claim})} \\ &= \frac{0.02 \times (20/100)}{1 - 0.958} = 0.0952.\end{aligned}$$

Similarly, we obtain $\Pr(" 30 - 50" | \text{Claim}) = 0.4762$ and $\Pr(" > 50" | \text{Claim}) = 0.4286.$

Example: Insurance Company

4a. We have

$$\begin{aligned} & \Pr ("30 - 50" | \text{Claim} \cap \{ "30 - 50" \cup " > 50" \}) \\ &= \frac{\Pr(\text{Claim} | \{ "30 - 50" \cup " > 50" \}) \Pr("30 - 50" | "30 - 50" \cup " > 50")}{\Pr(\text{Claim} \cap \{ "30 - 50" \cup " > 50" \})} \\ &= \frac{\Pr(\text{Claim} | "30 - 50") \Pr("30 - 50" | "30 - 50" \cup " > 50")}{\Pr(\text{Claim} \cap \{ "30 - 50" \cup " > 50" \})} \end{aligned}$$

$$\begin{aligned} & \text{where } \Pr(\text{Claim} \cap \{ "30 - 50" \cup " > 50" \}) = \\ & \Pr(\text{Claim} | "30 - 50") \Pr("30 - 50" | "30 - 50" \cup " > 50") + \\ & \Pr(\text{Claim} | " > 50") \Pr(" > 50" | "30 - 50" \cup " > 50"). \end{aligned}$$

4b. Hence it follows that the expected amount claimed is

$$\begin{aligned} & E[\text{Amount Claim} | \text{Claim} \cap \{ "30 - 50" \cup " > 50" \}] \\ &= 1000 \times \Pr("30 - 50" | \text{Claim} \cap \{ "30 - 50" \cup " > 50" \}) \\ & \quad + 1500 \times \Pr(" > 50" | \text{Claim} \cap \{ "30 - 50" \cup " > 50" \}) \\ &= \frac{0.4762}{0.4762 + 0.4286} \times 1000 + \frac{0.4286}{0.4762 + 0.4286} \times 1500 \\ &= 1236, 8\$ \end{aligned}$$

Conditional Densities: Continuous Case

- Given a joint p.d.f. for two r.v. X, Y it is possible to compute the conditional p.d.f. of X having observed $Y = y$.
- Assume X, Y are continuous-valued r.v. with a joint p.d.f. $f(x, y)$ then the conditional pdf of X given $Y = y$ is

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

- This can be heuristically established by noting that

$$\begin{aligned} f_{X|Y}(x|y) dx &= \frac{f(x, y) dx dy}{f_Y(y) dy} \\ &\approx \frac{P(x \leq X \leq x + dx \cap y \leq Y \leq y + dy)}{P(y \leq Y \leq y + dy)} \\ &= P(x \leq X \leq x + dx | y \leq Y \leq y + dy). \end{aligned}$$

- In the case where X and Y are independent, we have $f_{X|Y}(x|y) = f_X(x)$ as $f(x, y) = f_X(x) f_Y(y)$.

Conditional Densities: Continuous Case

- We have

$$f(x, y) = f_{X|Y}(x|y) f_Y(y)$$

and similarly

$$f(x, y) = f_{Y|X}(y|x) f_X(x)$$

- Hence we obtain

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

which holds if $f_Y(y) > 0$.

Conditional Expectation and Variance

- We can define the mean, variance of the conditional p.m.f.
- The conditional mean is given by

$$E(X|Y = y) = \int x f_{X|Y}(x|y) dx$$

- The conditional variance is given by

$$\begin{aligned} \text{Var}(X|Y = y) &= E\left((X - E(X|Y = y))^2 | Y = y\right) \\ &= E(X^2|Y = y) - \{E(X|Y = y)\}^2 \end{aligned}$$

where

$$E(X^2|Y = y) = \int x^2 f_{X|Y}(x|y) dx$$

- $E(X|Y = y)$ and $\text{Var}(X|Y = y)$ are functions but $E(X|Y)$ and $\text{Var}(X|Y)$ are random variables.

Example: Toy example

- Find the conditional pdf of Y given X and $E(Y|X = x)$ when their joint density is

$$f(x, y) = \lambda^2 e^{-\lambda y} \text{ for } 0 \leq x \leq y \leq \infty$$

- We have for $0 \leq x \leq \infty$

$$f_X(x) = \int_x^\infty \lambda^2 e^{-\lambda y} dy = \lambda e^{-\lambda x}$$

so for $0 \leq x \leq y \leq \infty$

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \lambda e^{\lambda(x-y)}$$

- Hence, we obtain

$$\begin{aligned} E(Y|X = x) &= \int_x^\infty y f_{Y|X}(y|x) dy \text{ (change of var. } u = y - x) \\ &= \int_0^\infty (u + x) \lambda e^{-\lambda u} du = \frac{1}{\lambda} + x. \end{aligned}$$

Example: Another Toy example

- Find the conditional pdf of Y given X and $E(Y|X = x)$ when their joint density is

$$f(x, y) = xe^{-x(y+1)} \text{ for } x, y \geq 0$$

- We have for $0 \leq x \leq \infty$

$$f_X(x) = \int_0^{\infty} xe^{-x(y+1)} dy = e^{-x}$$

so that for $0 \leq y \leq \infty$

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = xe^{-xy}$$

- It follows that

$$E(Y|X = x) = \int_0^{\infty} yf_{Y|X}(y|x) dy = \frac{1}{x}.$$

Optimality of Conditional Expectation

- Let us consider two r.v. X and Y . Assume we observe Y and want to find a way to estimate X based on Y . Then in some sense, $E(X|Y)$ is the best possible estimate of X .
- **Proposition.** Consider an arbitrary function $g(Y)$ then, we have

$$E \left[(X - E(X|Y))^2 \right] \leq E \left[(X - g(Y))^2 \right],$$

that is the expected square “distance” between $g(Y)$ and X is minimized for $g(Y) = E(X|Y)$.

- This is valid for both discrete and continuous r.v.