PROBABILISTIC GRAPHICAL MODELS CPSC 532C (TOPICS IN AI) STAT 521A (TOPICS IN MULTIVARIATE ANALYSIS)

Lecture 6

Kevin Murphy

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• Discussion section on Thursday, 3.30-4, in 304 (this week only).

- There are several kinds of queries we can make.
- Suppose the joint is $P(Y, E, W) = P(Y, W) \times P(E|Y, W)$.
- Conditional probability queries (sum-product):

$$P(Y|E=e) \propto \sum_{W} P(Y,W) \times P(e|Y,W)$$

• Most probable explanation (MPE) queries (max-product, MAP):

$$(y, w)^* = \arg\max_y \max_w P(Y, W) \times P(e|Y, W)$$

 Maximum A Posteriori (MAP) queries (max-sum-product, marginal MAP)

$$y^* = \arg\max_{y} \sum_{w} P(Y, W) \times P(e|Y, W)$$

INFERENCE IN HIDDEN MARKOV MODELS (HMM)



• Conditional probability queries, e.g. estimate current state given past evidence (online filtering)

$$P(X_t|e_{1:t}) = \sum_{x_{1:t-1}} P(x_{1:t-1}, X_t|e_{1:t})$$

• Most probable explanation (MPE) queries, e.g., most probable sequence of states (Viterbi decoding)

$$x_{1:t}^* = \arg\max_{x_{1:t}} P(x_{1:t}|e_{1:t})$$

• Note: Most probable sequence of states not necessarily equal to sequence of most probable states.

• e.g.,
$$X_1 \to X_2$$

 $P(X_1) \begin{pmatrix} 0.4 \\ 0.6 \end{pmatrix} P(X_2 | X_1) \begin{pmatrix} 0.1 & 0.9 \\ 0.5 & 0.5 \end{pmatrix} P(X_1, X_2) \begin{pmatrix} 0.04 & 0.36 \\ 0.3 & 0.3 \end{pmatrix}$
 $\arg \max_{x_1} P(X_1) = 1, \ \arg \max_{x_1} \max_{x_2} P(X_1, X_2) = (0, 1)$

• Viterbi decoding minimizes word error rate

$$x_{1:t}^* = \arg\max_{x_{1:t}} P(x_{1:t}|e_{1:t})$$

• To minimize bit error rate, use most marginally likely state

$$P(X_t|y_{1:t}) = \sum_{\substack{x_{1:t-1} \\ x_t^* = \max_x}} P(x_{1:t-1}, X_t|e_{1:t})$$

MAP VS MARGINAL MAP



- Consider a Dynamic Bayes Net (DBN) for speech recognition, where W = word and Q = phoneme.
- Most likely sequence of *states* (Viterbi/ MAP, max-product):

$$\arg \max_{q_{1:t}, w_{1:t}} P(q_{1:t}, w_{1:t} | e_{1:t})$$

• Most likely sequence of *words* (Marginal MAP, max-sum-product):

$$\arg\max_{w_{1:t}} \sum_{q_{1:t}} P(w_{1:t}, q_{1:t} | e_{1:t})$$

• Max-product often used as computationally simpler approximation to max-sum-product (or can use A^* decoding).

- Determinining if $P_B(X = x) > 0$ for some (discrete) variable X and some Bayes net B is NP-complete.
- What does this mean?
- Roughly: The best algorithm for exact inference (in discrete-state models) probably takes exponential time, in the worst case.
- More formally: we need a review of basic computational complexity theory.

- Defn: a decision problem is a task of the form: does there exist a solution which satisfies these conditions?
- Example: boolean satisfiability:

$$(q_1 \lor \neg q_2 \lor q_3) \land (\neg q_q \lor q_2 \lor \neg q_3)$$

is satisfiable ($q_1 = q_2 = q_3$ =true)

• 3-SAT is boolean satisfiability where $\phi = C_1 \wedge C_2 \dots \wedge C_n$, and every clause C_i has 3 literals.

- Defn: A decision problem Π is in P if it can be solved in polynomial time.
- Defn: ∏ is in NP if it can be solved in polynomial time using a non-deterministic oracle (i.e., you can verify its guesses in polytime).
- Defn: Π is NP-hard if $\forall \Pi' \in NP$. $\exists T \in P$. $\Pi' \xrightarrow{T} \Pi$.
- Defn: Π is NP-complete if it is NP-hard and in NP.
- Conjecture: $P \neq NP$



- Thm: 3-SAT is NP-complete.
- To show Π is NP-hard, it suffices to find a transformation $T \in P$ from another NP-hard problem Π' (e.g., 3-SAT) since

$$NP \xrightarrow{T'} \Pi' \xrightarrow{T} \Pi$$

• To show Π is NP-complete, show it is NP-hard and that you can check (oracular) guesses in poly-time.



EXACT INFERENCE IN DISCRETE BAYES NETS IS NP-COMPLETE

- Thm: the decision problem "Is $P_B(X_i = x) > 0$?" is NP-complete.
- Proof. To show in NP: Given an assignment $X_{1:n}$, we can check if $X_i = x$ and then check if $P(X_{1:n}) > 0$ in poly-time. To show NP-hard: we can encode any 3SAT problem as a polynomially sized Bayes net, as shown below.
- $P(X = 1|q_{1:n}) > 0$ iff $q_{1:n}$ is a satisfying assignment.



- Defn: An estimate ρ has absolute error ϵ for P(y|e) if $|P(y|e) \rho| \le \epsilon$.
- \bullet Defn: An estimate ρ has relative error ϵ for P(y|e) if

$$\frac{\rho}{1+\epsilon} \le P(y|e) \le \rho(1+\epsilon)$$

- Thm: Computing $P(X_i = x)$ with relative error ρ is NP-hard.
- Thm: Computing $P(X_i|e)$ with absolute error for any $\epsilon \in (0, 0.5)$ is NP-hard.
- But: special cases may have error bounds.
- And: heuristics often work well.

- We will show later that exact inference can always be done in time $O(Nv^W)$, where W_G is the *tree-width* of the graph (to be defined later) and $v = \max_i |X_i|$ is the max number of values (states) each node can take.
- The NP-hardness proof shows that, in the worst case, we have $W_G \sim N$.
- But for many models used in practice, we have $W_G \sim \text{constant}$.
- Also, for *Gaussian* graphical models, exact inference is $O(N^3)$ no matter what the graph structure is!

Exact inference in Gaussian models takes ${\cal O}(N^3)$ time

- \bullet For Gaussian graphical models, exact inference is ${\cal O}(N^3)$ no matter what the graph structure is!
- c.f., linear programming easier than integer programming.
- Lecture 3: Any undirected graphical model in which potentials have the form

$$\psi_{ij} = \exp(X_i - \mu_i) \Sigma_{ij}^{-1} (X_j - \mu_j)$$

can be converted to a joint Gaussian distribution.

 Book chap 4: any directed graphical model in which CPDs have the form

$$p(X_i|X_{\pi_i}) = \mathcal{N}(X_i; WX_{\pi_i} + \mu_i, \Sigma_i)$$

can be converted to a joint Gaussian distribution.

• Exact inference in a Gaussian graphical model = matrix inversion.

VARIABLE ELIMINATION ALGORITHM



- Key idea 1: push sum inside products.
- Key idea 2: use (non-serial) dynamic programming to cache shared subexpressions.

$$P(J) = \sum_{L} \sum_{S} \sum_{G} \sum_{H} \sum_{I} \sum_{D} \sum_{C} P(C, D, I, G, S, L, J, H)$$

$$= \sum_{L} \sum_{S} \sum_{G} \sum_{H} \sum_{I} \sum_{D} \sum_{C} P(C)P(D|C)P(I)P(G|I, D)P(S|I)P(L|G)P(J|L, S)P(H|G, J)$$

$$= \sum_{L} \sum_{S} \sum_{G} \sum_{H} \sum_{I} \sum_{D} \sum_{C} \phi_{C}(C)\phi_{D}(D, C)\phi_{I}(I)\phi_{G}(G, I, D)\phi_{S}(S, I)\phi_{L}(L, G)\phi_{J}(J, L, S)\phi_{H}(H, G, J)$$

$$= \sum_{L} \sum_{S} \phi_{J}(J, L, S) \sum_{G} \phi_{L}(L, G) \sum_{H} \phi_{H}(H, G, J) \sum_{I} \phi_{S}(S, I)\phi_{I}(I) \sum_{D} \phi_{C}(C)\phi_{D}(D, C)\phi_{I}(D, C)\phi_{I}(I)\phi_{I}(I) \sum_{D} \phi_{I}(G, I, D) \sum_{C} \phi_{C}(C)\phi_{D}(D, C)\phi_{I}(D, C)\phi_{I}(I)\phi_{I}(I)\phi_{I}(I)\phi_{I}(I)\phi_{I}(I)\phi_{I}(I, D)\phi_{I}(I, D)\phi_{I}(I,$$

$$\begin{split} P(J) &= \sum_{L} \sum_{S} \phi_J(J, L, S) \sum_{G} \phi_L(L, G) \sum_{H} \phi_H(H, G, J) \sum_{I} \phi_S(S, I) \phi_I(I) \sum_{D} \phi_I(G, I, D) \underbrace{\sum_{C} \phi_C(C) \phi_D(D, C)}_{\tau_1(D)} \\ &= \sum_{L} \sum_{S} \phi_J(J, L, S) \sum_{G} \phi_L(L, G) \sum_{H} \phi_{HI}(H, G, J) \sum_{I} \phi_S(S, I) \phi_I(I) \underbrace{\sum_{D} \phi_I(G, I, D) \tau_1(D)}_{\tau_2(G, I)} \\ &= \sum_{L} \sum_{S} \phi_J(J, L, S) \sum_{G} \phi_L(L, G) \sum_{H} \phi_H(H, G, J) \underbrace{\sum_{I} \phi_S(S, I) \phi_I(I) \tau_2(G, I)}_{\tau_3(G, S)} \\ &= \sum_{L} \sum_{S} \phi_J(J, L, S) \underbrace{\sum_{G} \phi_L(L, G) \sum_{H} \phi_H(H, G, J) \tau_3(G, S)}_{\tau_5(J, L, S)} \\ &= \sum_{L} \underbrace{\sum_{S} \phi_J(J, L, S) \sum_{G} \phi_L(L, G) \tau_1(G, J) \tau_3(G, S)}_{\tau_6(J, L)} \\ &= \sum_{L} \underbrace{\sum_{T} \phi_J(J, L, S) \sum_{T} \phi_J(J, L, S)}_{\tau_6(J, L)} \\ &= \sum_{L} \underbrace{\sum_{T} \tau_6(J, L)}_{\tau_7(J)} \end{split}$$

DIFFERENT ORDERING

$$\begin{split} P(J) &= \sum_{D} \sum_{C} \phi_{D}(D,C) \sum_{H} \sum_{L} \sum_{S} \phi_{J}(J,L,S) \sum_{I} \phi_{I}(I) \phi_{S}(S,I) \underbrace{\sum_{G} \phi_{G}(G,I,D) \phi_{L}(L,) \phi_{H}(H,G,J)}_{\tau_{I}(I,D,L,J,H)} \\ &= \sum_{D} \sum_{C} \phi_{D}(D,C) \sum_{H} \sum_{L} \sum_{S} \phi_{J}(J,L,S) \underbrace{\sum_{I} \phi_{I}(I) \phi_{S}(S,I) \tau_{1}(I,D,L,J,H)}_{\tau_{2}(D,L,S,J,H)} \\ &= \sum_{D} \sum_{C} \phi_{D}(D,C) \sum_{H} \underbrace{\sum_{L} \sum_{S} \phi_{J}(J,L,S) \tau_{2}(D,L,S,J,H)}_{\tau_{3}(D,L,J,H)} \\ &= \sum_{D} \sum_{C} \phi_{D}(D,C) \sum_{H} \underbrace{\sum_{L} \tau_{3}(D,L,J,H)}_{\tau_{4}(D,J,H)} \\ &= \sum_{D} \sum_{C} \phi_{D}(D,C) \underbrace{\sum_{H} \tau_{4}(D,J,H)}_{\tau_{5}(D,J)} \\ &= \sum_{D} \sum_{C} \phi_{D}(D,C) \underbrace{\sum_{H} \tau_{4}(D,J,H)}_{\tau_{6}(D,J)} \\ &= \sum_{D} \sum_{C} \phi_{D}(D,C) \tau_{5}(D,J) \\ &= \underbrace{\sum_{D} \tau_{6}(D,J)}_{\tau_{7}(J)} \end{split}$$

DEALING WITH EVIDENCE: METHOD 1



- We can instantiate observed variables to their observed value: $P(J|I = 1, H = 0) = \frac{P(J, I = 1, H = 0)}{P(I = 1, H = 0)} \propto P(J, I = 1, H = 0)$ $= \sum_{C,D,G,L,S} P(C, D, I = 1, G, S, L, J, H = 0)$
- The denominator is P(e) = P(I = 1, H = 0).
- For Markov networks, the denominator is $P(e) \times Z$.

• We can associate a local evidence potential with every node, and set $\phi_i(X_i) = \delta(X_i, x_i^*)$ if X_i is observed to have value x_i^* , and $\phi_i(X_i) = 1$ otherwise:

$$P(X_{1:n}|ev) \propto P(X_{1:n}) \prod_{i} P(ev_i|X_i)$$

• e.g.,

$$P(J|I = 1, H = 0) \propto \sum_{C,D,I,G,S,L,J,H} P(C, D, I, G, S, L, J, H) \delta_I(I, 1) \delta_H(H, 0)$$